ALEA, Lat. Am. J. Probab. Math. Stat. 12 (1), 309-356 (2015)



# The Stein and Chen-Stein Methods for Functionals of Non-Symmetric Bernoulli Processes

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**Abstract.** Based on a new multiplication formula for discrete multiple stochastic integrals with respect to non-symmetric Bernoulli random walks, we extend the results of Nourdin et al. (2010) on the Gaussian approximation of symmetric Rademacher sequences to the setting of possibly non-identically distributed independent Bernoulli sequences. We also provide Poisson approximation results for these sequences, by following the method of Peccati (2011). Our arguments use covariance identities obtained from the Clark-Ocone representation formula in addition to those usually based on the inverse of the Ornstein-Uhlenbeck operator.

## 1. Introduction

Malliavin calculus and the Stein method were combined for the first time for Gaussian fields in the seminal paper Nourdin and Peccati (2009), whose results have later been extended to other settings, including Poisson processes Peccati (2011); Peccati et al. (2010). In particular, the Stein method has been applied in Nourdin et al. (2010) to Rademacher sequences  $(X_n)_{n\in\mathbb{N}}$  of independent and identically distributed Bernoulli random variables with  $P(X_1 = 1) = P(X_1 = -1) = 1/2$ , in order to derive bounds on distances between the probability laws of functionals of  $(X_n)_{n\in\mathbb{N}}$  and the law  $\mathcal{N}(0,1)$  of a standard  $\mathcal{N}(0,1)$  normal random variable Z. Those approaches exploit a covariance representation based on the number (or Ornstein-Uhlenbeck) operator L and its inverse  $L^{-1}$ .

Received by the editors March 31, 2014; accepted April 6, 2015.

<sup>2010</sup> Mathematics Subject Classification. 60F05, 60G57, 60H07.

*Key words and phrases.* Bernoulli processes; Chen-Stein method; Stein method; Malliavin calculus; Clark-Ocone representation .

From here onwards, we denote by  $C_b^2$  the set of all real-valued bounded functions with bounded derivatives up to the second order. In particular, for  $h \in C_b^2$ , using a chain rule proved in the symmetric case, the bound

$$\mathbb{E}[h(F)] - \mathbb{E}[h(Z)]| \le A_1 \min\{4\|h\|_{\infty}, \|h''\|_{\infty}\} + \|h''\|_{\infty}A_2,$$
(1.1)

has been derived in Nourdin et al. (2010) (see Theorem 3.1 therein) for centered functionals F of a symmetric Bernoulli random walk  $(X_n)_{n \in \mathbb{N}}$ . Here,  $(X_n)_{n \in \mathbb{N}}$  is built as the sequence of canonical projections on  $\Omega := \{-1, 1\}^{\mathbb{N}}$  and

$$A_1 = \mathbf{E}\left[\left|1 - \langle DF, -DL^{-1}F \rangle_{\ell^2(\mathbb{N})}\right|\right], \qquad A_2 = \frac{20}{3} \mathbf{E}\left[\langle |DL^{-1}F|, |DF|^3 \rangle_{\ell^2(\mathbb{N})}\right],$$

where  $\langle \cdot, \cdot \rangle_{\ell^2(\mathbb{N})}$  is the usual inner product on  $\ell^2(\mathbb{N})$ , and D is the symmetric gradient defined as

$$D_k F(\omega) = \frac{1}{2} (F(\omega_+^k) - F(\omega_-^k)), \quad k \in \mathbb{N},$$

where, given  $\omega = (\omega_0, \omega_1, \ldots) \in \Omega$ , we let

$$\omega_+^k = (\omega_0, \dots, \omega_{k-1}, +1, \omega_{k+1}, \dots)$$

and

$$\omega_{-}^{k} = (\omega_0, \ldots, \omega_{k-1}, -1, \omega_{k+1}, \ldots).$$

The above bound (1.1) can be used to control the Wasserstein distance between  $\mathcal{N}(0, 1)$  and the law of F as in Corollary 3.6 in Nourdin et al. (2010). In addition, the right-hand side of (1.1) yields explicit bounds in the case where F is a single discrete stochastic integral (see Corollary 3.3 in Nourdin et al. (2010)) or a multiple discrete stochastic integral (see Section 4 in Nourdin et al. (2010)). In this latter case the derivation of explicit bounds is based on a multiplication formula proved in the symmetric case (see Proposition 2.9 in Nourdin et al. (2010)).

In this paper we provide Gaussian and Poisson approximations for functionals of not-necessarily symmetric Bernoulli sequences via the Stein and Chen-Stein methods, respectively. See Krokowski et al. (2015) for recent related results on Gaussian approximation, without relying on a multiplication formula for discrete multiple stochastic integrals.

The normal and Poisson approximations are based on suitable chain rules in Propositions 2.1 and 2.2 and on an extension to the non-symmetric case of the multiplication formula for discrete multiple stochastic integrals (see Proposition 5.1 and Section 9 for its proof). In addition to using the Ornstein-Uhlenbeck operator L for covariance representations, we also derive error bounds for the normal and Poisson approximations using covariance representations based on the Clark-Ocone formula, following the argument implemented in Privault and Torrisi (2013). Indeed the operator L is of a more delicate use in applications to functionals whose multiple stochastic integral expansion is not explicitly known. In contrast with covariance identities based on the number operator, which rely on the divergence-gradient composition, the Clark-Ocone formula only requires the computation of a gradient and a conditional expectation.

A bound for the Wasserstein distance between a standard Gaussian random variable and a (standardized) function of a finite sequence of independent random variables has been obtained in Chatterjee (2008), via the construction of an auxiliary random variable which allows one to approximate the Stein equation. Although the results in our paper are restricted to the Bernoulli case, they may be applied to functionals of an infinite sequence of Bernoulli distributed random variables. A

comparison between a bound in our paper and that one in Chatterjee (2008) is given at the end of the first example of Section 4.

As far as the Gaussian approximation is concerned, using a covariance representation based on the Clark-Ocone formula, in Theorem 3.2 below we find sufficient conditions on centered functionals F of a not necessarily symmetric Bernoulli random walk so that

$$|\mathbf{E}[h(F)] - \mathbf{E}[h(Z)]| \le B_1 \min\{4\|h\|_{\infty}, \|h''\|_{\infty}\} + \|h'\|_{\infty}B_2 + \|h''\|_{\infty}B_3 \qquad (1.2)$$

for any  $h \in \mathbb{C}_b^2$ , and for some positive constants  $B_1, B_2, B_3 > 0$ ; similarly, using a covariance representation based on the Ornstein-Uhlenbeck operator, in Theorem 3.4 below we provide alternate sufficient conditions on centered functionals Fof a not necessarily symmetric Bernoulli random walk so that the bound (1.2) holds for different positive constants  $C_1, C_2, C_3 > 0$ , in place of  $B_1, B_2, B_3$  respectively. In Theorem 3.6 below we show that the bound (1.2) can be used to control the Fortet-Mourier distance  $d_{\rm FM}$  between F and the standard  $\mathcal{N}(0, 1)$  normal random variable Z, i.e. we prove

$$d_{\rm FM}(F,Z) \le \sqrt{2(B_1 + B_3)(5 + {\rm E}[|F|])} + B_2$$

A similar bound holds, under alternate conditions on F, with the constant  $B_i$  replaced by  $C_i$  (i = 1, 2, 3). Replacing the Stein method by the Chen-Stein method, we also show that this approach applies to the Poisson approximation in addition to the Gaussian approximation, and treat discrete multiple stochastic integrals as examples in both cases.

This paper is organized as follows. In Section 2 we recall some elements of stochastic analysis of Bernoulli processes, including chain rules for finite difference operators. In Section 3 we present the two different upper bounds for the quantity |E[h(F)] - E[h(Z)]|,  $h \in C_b^2$ , described above and the related application to the Fortet-Mourier distance. Section 4 contains explicit first chaos bounds with application to determinantal processes, while Section 5 is concerned with bounds for the *n*th chaoses. The important case of quadratic functionals (second chaoses) is treated in a separate paragraph. In Section 6 we apply our arguments to the Poisson approximation and in Sections 7 and 8 we investigate the case of single and multiple discrete stochastic integrals. Finally, Section 9 deals with the new multiplication formula for discrete multiple stochastic integrals in the non-symmetric case, whose proof is modeled on normal martingales that are solution of a deterministic structure equation.

#### 2. Stochastic analysis of Bernoulli processes

In this section we provide some preliminaries. The reader is directed to Privault (2008) and references therein for more insight into the stochastic analysis of Bernoulli processes.

From now on we assume that the canonical projections  $X_n : \Omega \to \{-1, 1\}$ ,  $\Omega = \{-1, 1\}^{\mathbb{N}}$ , are considered under the not necessarily symmetric measure P given on cylinder sets by

$$P(\{\varepsilon_0, \dots, \varepsilon_n\} \times \{-1, 1\}^{\mathbb{N}}) = \prod_{k=0}^n p_k^{(1+\varepsilon_k)/2} q_k^{(1-\varepsilon_k)/2}, \quad \varepsilon_k \in \{-1, 1\}, \ k = 0, \dots, n.$$

Given  $\omega = (\omega_0, \omega_1, \ldots) \in \Omega$  and  $\omega_+^k, \omega_-^k$  defined as above, for any  $F : \Omega \to \mathbb{R}$  we consider the finite difference operator

$$D_k F(\omega) = \sqrt{p_k q_k} (F(\omega_+^k) - F(\omega_-^k)), \quad k \in \mathbb{N}$$

and, denoting by  $\kappa$  the counting measure on  $\mathbb{N}$ , we consider the  $L^2(\Omega \times \mathbb{N}) = L^2(\Omega \times \mathbb{N}, P \otimes \kappa)$ -valued operator D defined for any  $F : \Omega \to \mathbb{R}$ , by  $DF = (D_k F)_{k \in \mathbb{N}}$ . Given  $n \geq 1$  we denote by  $\ell^2(\mathbb{N})^{\otimes n} = \ell^2(\mathbb{N}^n)$  the class of functions on  $\mathbb{N}^n$  that are square integrable with respect to  $\kappa^{\otimes n}$ , we denote by  $\ell^2(\mathbb{N})^{\circ n}$  the subspace of  $\ell^2(\mathbb{N})^{\otimes n}$  formed by functions that are symmetric in n variables. The  $L^2$  domain of D is given by

$$Dom(D) = \{ F \in L^{2}(\Omega) : DF \in L^{2}(\Omega \times \mathbb{N}) \} = \{ F \in L^{2}(\Omega) : E[\|DF\|^{2}_{\ell^{2}(\mathbb{N})}] < \infty \}.$$

We let  $(Y_n)_{n\geq 0}$  denote the sequence of centered and normalized random variables defined by

$$Y_n = \frac{q_n - p_n + X_n}{2\sqrt{p_n q_n}},$$

which satisfies the discrete structure equation

$$Y_n^2 = 1 + \frac{q_n - p_n}{2\sqrt{p_n q_n}} Y_n.$$
 (2.1)

Given  $f_1 \in \ell^2(\mathbb{N})$  we define the first order discrete stochastic integral of  $f_1$  as

$$J_1(f_1) = \sum_{k \ge 0} f_1(k) Y_k,$$

and we let

$$J_n(f_n) = \sum_{(i_1,\dots,i_n)\in\Delta_n} f_n(i_1,\dots,i_n)Y_{i_1}\dots Y_{i_n}$$

denote the discrete multiple stochastic integral of order n of  $f_n$  in the subspace  $\ell^2_{\mathfrak{s}}(\Delta_n)$  of  $\ell^2(\mathbb{N})^{\circ n}$  composed of symmetric kernels that vanish on diagonals, i.e. on the complement of

$$\Delta_n = \{ (k_1, \dots, k_n) \in \mathbb{N}^n : k_i \neq k_j, \ 1 \le i < j \le n \}, \quad n \ge 1$$

As a convention we identify  $\ell^2(\mathbb{N}^0)$  to  $\mathbb{R}$  and let  $J_0(f_0) = f_0, f_0 \in \mathbb{R}$ . Hereafter, we shall refer to the set of functionals of the form  $J_n(f)$  as the *n*-chaos. The multiple stochastic integrals satisfy the isometry formula

$$\mathbb{E}[J_n(f_n)J_m(g_m)] = \mathbb{1}_{\{n=m\}} n! \langle f_n, g_m \rangle_{\ell^2_s(\Delta_n)},$$

 $f_n \in \ell_{\mathfrak{s}}^2(\Delta_n), g_m \in \ell_{\mathfrak{s}}^2(\Delta_m)$ , cf. e.g. Proposition 1.3.2 of Privault (2009).

The finite difference operator acts on multiple stochastic integrals as follows:

$$D_k J_n(f_n) = n J_{n-1}(f_n(*,k) \mathbb{1}_{\Delta_n}(*,k)) = n J_{n-1}(f_n(*,k)),$$
(2.2)

 $k \in \mathbb{N}$ ,  $f_n \in \ell^2_{\mathfrak{s}}(\Delta_n)$ . Due to the chaos representation property any square integrable F may be represented as  $F = \sum_{n \geq 0} J_n(f_n)$ ,  $f_n \in \ell^2_{\mathfrak{s}}(\Delta_n)$ , and so the  $L^2$  domain of D may be rewritten as

Dom(D) = 
$$\left\{ F = \sum_{n \ge 0} J_n(f_n) : \sum_{n \ge 1} n \, n! \|f_n\|_{\ell^2(\mathbb{N})^{\otimes n}}^2 < \infty \right\}.$$

Next we present a chain rule for the finite difference operator that extends Proposition 2.14 in Nourdin et al. (2010) from the symmetric to the non-symmetric case.

This chain rule will be used later on for the normal approximation. In the following we write  $F_k^{\pm}$  in place of  $F(\omega_{\pm}^k)$ .

**Proposition 2.1.** Let  $F \in \text{Dom}(D)$  and  $f : \mathbb{R} \to \mathbb{R}$  be thrice differentiable with bounded third derivative. Assume moreover that  $f(F) \in \text{Dom}(D)$ . Then, for any integer  $k \ge 0$  there exists a random variable  $R_k^F$  such that

$$D_k f(F) = f'(F) D_k F - \frac{|D_k F|^2}{4\sqrt{p_k q_k}} (f''(F_k^+) + f''(F_k^-)) X_k + R_k^F, \quad a.s.$$
(2.3)

where

$$|R_k^F| \le \frac{5}{3!} \|f^{'''}\|_{\infty} \frac{|D_k F|^3}{p_k q_k}, \quad a.s.$$
(2.4)

Proof. By a standard Taylor expansion we have

$$D_{k}f(F) = \sqrt{p_{k}q_{k}}(f(F_{k}^{+}) - f(F_{k}^{-})) = \sqrt{p_{k}q_{k}}(f(F_{k}^{+}) - f(F)) -\sqrt{p_{k}q_{k}}(f(F_{k}^{-}) - f(F)) = \sqrt{p_{k}q_{k}}f'(F)(F_{k}^{+} - F) + \frac{\sqrt{p_{k}q_{k}}}{2}f''(F)(F_{k}^{+} - F)^{2} + R_{k}^{+} -\sqrt{p_{k}q_{k}}f'(F)(F_{k}^{-} - F) - \frac{\sqrt{p_{k}q_{k}}}{2}f''(F)(F_{k}^{-} - F)^{2} + R_{k}^{-} = f'(F)D_{k}F + \frac{\sqrt{p_{k}q_{k}}}{2}f''(F)[(F_{k}^{+} - F)^{2} -(F_{k}^{-} - F)^{2}] + R_{k}^{+} + R_{k}^{-},$$
(2.5)

where

$$R_k^{\pm}| \le \frac{\sqrt{p_k q_k}}{3!} \|f^{'''}\|_{\infty} |F_k^{\pm} - F|^3.$$
(2.6)

By the mean value theorem we find

$$\begin{split} f''(F) &= \frac{f''(F_k^+) + f''(F_k^-)}{2} + \frac{f''(F) - f''(F_k^+) + f''(F) - f''(F_k^-)}{2} \\ &= \frac{f''(F_k^+) + f''(F_k^-)}{2} + R_k', \end{split}$$

where

$$|R_{k}^{'}| \leq \frac{\|f^{'''}\|_{\infty}}{2}(|F_{k}^{+} - F| + |F_{k}^{-} - F|).$$

Substituting this into (2.5) we get

$$D_k f(F) = f'(F) D_k F + \frac{\sqrt{p_k q_k}}{4} (f''(F_k^+) + f''(F_k^-)) [(F_k^+ - F)^2 - (F_k^- - F)^2] + R_k^+ + R_k^- + R_k^*,$$
(2.7)

where

$$|R_k^*| \le \frac{\sqrt{p_k q_k}}{4} \|f^{'''}\|_{\infty} (|F_k^+ - F| + |F_k^- - F|) (|F_k^+ - F|^2 - |F_k^- - F|^2) \\ \le \frac{\sqrt{p_k q_k}}{4} \|f^{'''}\|_{\infty} (|F_k^+ - F| + |F_k^- - F|) |F_k^+ - F|^2.$$
(2.8)

Note that

$$F_{k}^{+} - F = (F_{k}^{+} - F) \mathbb{1}_{\{X_{k}=-1\}} + (F_{k}^{+} - F) \mathbb{1}_{\{X_{k}=1\}} = (F_{k}^{+} - F) \mathbb{1}_{\{X_{k}=-1\}}$$
  
=  $(F_{k}^{+} - F_{k}^{-}) \mathbb{1}_{\{X_{k}=-1\}},$  (2.9)

and similarly,

$$F_k^- - F = -(F_k^+ - F_k^-) \mathbb{1}_{\{X_k=1\}}.$$
(2.10)

Therefore we have  $|F_k^{\pm} - F| \leq |D_k F| / \sqrt{p_k q_k}$ , and combining this with (2.6) and (2.8) we find

$$|R_k^{\pm}| \le \frac{\|f^{'''}\|_{\infty}}{3! p_k q_k} |D_k F|^3, \quad |R_k^*| \le \frac{\|f^{'''}\|_{\infty}}{2p_k q_k} |D_k F|^3.$$
(2.11)

By (2.9) and (2.10) we also have

 $(F_k^+ - F)^2 = (F_k^+ - F_k^-)^2 \mathbb{1}_{\{X_k = -1\}} \text{ and } (F_k^- - F)^2 = (F_k^+ - F_k^-)^2 \mathbb{1}_{\{X_k = 1\}},$ therefore

$$(F_k^+ - F)^2 - (F_k^- - F)^2 = (F_k^+ - F_k^-)^2 (\mathbb{1}_{\{X_k = -1\}} - \mathbb{1}_{\{X_k = 1\}})$$
$$= -(F_k^+ - F_k^-)^2 X_k = -\frac{|D_k F|^2}{p_k q_k} X_k.$$

The claim follows substituting this expression into (2.7) and by using (2.11) to estimate the remainder.

Now we present a chain rule for the finite difference operator, which is suitable for integer-valued functionals. This chain rule will be used later on for the Poisson approximation. Given a function  $f : \mathbb{N} \to \mathbb{R}$  we define the operators

$$\Delta f(k) := f(k+1) - f(k), \quad \Delta^2 f := \Delta(\Delta f).$$

**Proposition 2.2.** Let  $F \in Dom(D)$  be an  $\mathbb{N}$ -valued random variable. Then, for any  $f : \mathbb{N} \to \mathbb{R}$  so that  $f(F) \in Dom(D)$ , we have

$$D_k f(F) = \Delta f(F) D_k F + R_k^F, \qquad (2.12)$$

where

$$|R_{k}^{F}| \leq \frac{\|\Delta^{2}f\|_{\infty}}{2} \left( \left| \frac{D_{k}F}{\sqrt{p_{k}q_{k}}} \mathbb{1}_{\{X_{k}=-1\}} \left( \frac{D_{k}F}{\sqrt{p_{k}q_{k}}} - 1 \right) \right| + \left| \frac{D_{k}F}{\sqrt{p_{k}q_{k}}} \mathbb{1}_{\{X_{k}=1\}} \left( \frac{D_{k}F}{\sqrt{p_{k}q_{k}}} + 1 \right) \right| \right).$$

$$(2.13)$$

*Proof.* As shown in the proof of Theorem 3.1 in Peccati (2011), for any  $f : \mathbb{N} \to \mathbb{R}$  and any  $k, a \in \mathbb{N}$ ,

$$|f(k) - f(a) - \Delta f(a)(k-a)| \le \frac{\|\Delta^2 f\|_{\infty}}{2} |(k-a)(k-a-1)|.$$
(2.14)

Therefore, taking first  $k = F_k^+$ , a = F and then  $k = F_k^-$ , a = F, we deduce

$$D_k f(F) = \sqrt{p_k q_k} (f(F_k^+) - f(F)) - \sqrt{p_k q_k} (f(F_k^-) - f(F))$$
  
=  $\sqrt{p_k q_k} \Delta f(F) (F_k^+ - F) + R_k^{(1)} + \sqrt{p_k q_k} \Delta f(F) (F_k^- - F) + R_k^{(2)},$ 

where by (2.14), setting  $R_k^F := R_k^{(1)} + R_k^{(2)}$ , we have

$$|R_k^F| \le \frac{\|\Delta^2 f\|_{\infty}}{2} (|(F_k^+ - F)(F_k^+ - F - 1)| + |(F_k^- - F)(F_k^- - F - 1)|).$$

The claim follows from (2.9) and (2.10).

Next we give two alternative covariance representation formulas. Let  $(\mathcal{F}_n)_{n\geq -1}$  be the filtration defined by

$$\mathfrak{F}_{-1} = \{\emptyset, \Omega\}, \qquad \mathfrak{F}_n = \sigma\{X_0, \dots, X_n\}, \quad n \ge 0.$$

By Proposition 1.10.1 of Privault (2009), for any  $F, G \in \text{Dom}(D)$  with F centered we have the Clark-Ocone covariance representation formula

$$\operatorname{Cov}(F,G) = \operatorname{E}[FG] = \operatorname{E}\left[\sum_{k\geq 0} \operatorname{E}[D_k F \mid \mathcal{F}_{k-1}]D_k G\right].$$
(2.15)

The second covariance representation formula involves the inverse of the Ornstein-Uhlenbeck operator. The domain Dom(L) of the Ornstein-Uhlenbeck operator  $L : L^2(\Omega) \to L^2_0(\Omega)$ , where  $L^2_0(\Omega)$  denotes the subspace of  $L^2(\Omega)$  composed of centered random variables, is given by

$$Dom(L) = \left\{ F = \sum_{n \ge 0} J_n(f_n) : \sum_{n \ge 1} n^2 n! \|f_n\|_{\ell^2(\mathbb{N})^{\otimes n}}^2 < \infty \right\}$$

and, for any  $F \in \text{Dom}(L)$ ,

$$LF = -\sum_{n=1}^{\infty} nJ_n(f_n).$$

The inverse of L, denoted by  $L^{-1}$ , is defined on  $L^2_0(\Omega)$  by

1

$$L^{-1}F = -\sum_{n=1}^{\infty} \frac{1}{n} J_n(f_n),$$

with the convention  $L^{-1}F = L^{-1}(F - \mathbb{E}[F])$  in case F is not centered, as in e.g. Peccati (2011). Using this convention, for any  $F, G \in \text{Dom}(D)$  we have

$$\operatorname{Cov}(F,G) = \operatorname{E}[G(F - \operatorname{E}[F])] = \operatorname{E}\left[\langle DG, -DL^{-1}F \rangle_{\ell^{2}(\mathbb{N})}\right], \qquad (2.16)$$

cf. Lemma 2.12 of Nourdin et al. (2010) in the symmetric case. For the sake of completeness, we provide an alternative expression for the covariance representation formula (2.16). Let  $(P_t)_{t\geq 0}$  be the semigroup associated to the Ornstein-Uhlenbeck operator L (we refer the reader to Section 10 of Privault (2008) for the details). Then

$$P_t J_n(f_n) = e^{-nt} J_n(f_n), \quad n \ge 1,$$

and so for any  $F = \sum_{n=0}^{\infty} J_n(f_n) \in \text{Dom}(D)$  one has

$$\int_{0}^{\infty} e^{-t} P_{t} D_{k} F dt = \sum_{n=1}^{\infty} n \int_{0}^{\infty} e^{-t} P_{t} J_{n-1}(f_{n}(*,k)) dt$$
$$= \sum_{n\geq 1} n \int_{0}^{\infty} e^{-t} e^{-(n-1)t} J_{n-1}(f_{n}(*,k)) dt$$
$$= \sum_{n=1}^{\infty} J_{n-1}(f_{n}(*,k))$$
$$= -D_{k} L^{-1} F.$$

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Consequently, the covariance representation (2.16) may be rewritten as

$$\operatorname{Cov}(F,G) = \operatorname{E}[G(F - \operatorname{E}[F])] = \operatorname{E}\left[\sum_{k=0}^{\infty} \int_{0}^{\infty} \operatorname{e}^{-t} D_{k} G P_{t} D_{k} F \, \mathrm{d}t\right],$$

for any  $F, G \in \text{Dom}(D)$ , cf. Proposition 1.10.2 of Privault (2009).

#### 3. Normal approximation of Bernoulli functionals

In this section we present two different upper bounds for the quantity |E[h(F)] - E[h(Z)]|,  $h \in C_b^2$ . The first one is obtained by using the covariance representation formula (2.15), while the second one, obtained by using the covariance representation formula (2.16), is a strict extension of the bound given in Theorem 3.1 of Nourdin et al. (2010).

Before proceeding further we recall some necessary background on the Stein method for the normal approximation and refer to Barbour (1990); Götze (1991); Stein (1972, 1986) and to Nourdin et al. (2010) for more insight into this technique.

Stein's method for normal approximation. Let Z be a standard  $\mathcal{N}(0, 1)$  normal random variable and consider the so-called Stein's equation associated with  $h : \mathbb{R} \to \mathbb{R}$ :

$$h(x) - \mathbb{E}[h(Z)] = f'(x) - xf(x), \quad x \in \mathbb{R}.$$

We refer to part (ii) of Lemma 2.15 in Nourdin et al. (2010) for the following lemma. More precisely, the estimates on the first and second derivative are proved in Lemma II.3 of Stein (1986), the estimate of the third derivative is proved in Theorem 1.1 of Daly (2008) and the alternative estimate on the first derivative may be found in Barbour (1990) and Götze (1991).

**Lemma 3.1.** If  $h \in C_b^2$ , then the Stein equation has a solution  $f_h$  which is thrice differentiable and such that  $\|f'_h\|_{\infty} \leq 4\|h\|_{\infty}$ ,  $\|f''_h\|_{\infty} \leq 2\|h'\|_{\infty}$  and  $\|f'''_h\|_{\infty} \leq 2\|h''\|_{\infty}$ . We also have  $\|f'_h\|_{\infty} \leq \|h''\|_{\infty}$ .

Combining the Stein equation with this lemma, for a generic square integrable and centered random variable F we have

$$|\mathbf{E}[h(F)] - \mathbf{E}[h(Z)]| = |\mathbf{E}[f'_{h}(F) - Ff_{h}(F)]|.$$
(3.1)

Let  $(F_n)_{n\geq 1}$  be a sequence of square integrable and centered random variables. If

$$|\mathbf{E}[h(F_n)] - \mathbf{E}[h(Z)]| \to 0, \qquad h \in \mathcal{C}_b^2$$

then  $(F_n)_{n\geq 1}$  converges to Z in distribution as n tends to infinity, and so an upper bound for the right-hand side of (3.1) may provide informations about this normal approximation.

The results of Sections 3.1 and 3.2 below are given in terms of bounds for |E[h(F)] - E[h(Z)]|, for test functions in  $\mathcal{C}_b^2$ , and they are applied in Section 3.3 to derive bounds for the Fortet-Mourier distance between the laws of two random variables X and Y, which is defined by

$$d_{\rm FM}(X,Y) = \sup_{h \in \mathcal{FM}} |\mathbf{E}[h(X)] - \mathbf{E}[h(Y)]|, \qquad (3.2)$$

where  $\mathcal{FM}$  is the class of functions h such that  $\|h\|_{BL} = \|h\|_L + \|h\|_{\infty} \leq 1$ , where  $\|\cdot\|_L$  denotes the standard Lipschitz semi-norm. Clearly, any  $h \in \mathcal{FM}$  is Lipschitz with Lipschitz constant less than or equal to 1 and so it is Lebesgue a.e. differentiable and

 $||h'||_{\infty} \leq 1$ . One can also shows that  $d_{\text{FM}}$  metrizes the convergence in distribution, see e.g. Chapter 11 in Dudley (2002).

# 3.1. Clark-Ocone bound.

**Theorem 3.2.** Let  $F \in Dom(D)$  be a centered random variable and assume that

$$B_{1}: = E\left[\left|1 - \sum_{k \geq 0} E[D_{k}F \mid \mathcal{F}_{k-1}]D_{k}F\right|\right],$$
  

$$B_{2}: = \sum_{k \geq 0} \frac{|1 - 2p_{k}|}{\sqrt{p_{k}q_{k}}} E[|E[D_{k}F \mid \mathcal{F}_{k-1}]||D_{k}F|^{2}],$$
(3.3)

$$B_3: = \frac{5}{3} \sum_{k \ge 0} \frac{1}{p_k q_k} \mathbb{E}[|\mathbb{E}[D_k F \mid \mathcal{F}_{k-1}]| |D_k F|^3]$$
(3.4)

are finite. Then we have

$$|\mathbf{E}[h(F)] - \mathbf{E}[h(Z)]| \le B_1 \min\{4\|h\|_{\infty}, \|h''\|_{\infty}\} + \|h'\|_{\infty}B_2 + \|h''\|_{\infty}B_3 \qquad (3.5)$$
  
for all  $h \in \mathcal{C}^2_b$ .

*Proof.* Since the first derivative of  $f_h$  is bounded we have that  $f_h$  is Lipschitz. So  $f_h(F) \in L^2(\Omega)$  and

$$|D_k f_h(F)| = \sqrt{p_k q_k} |f_h(F_k^+) - f_h(F_k^-)| \le ||f_h'||_{\infty} |D_k F|.$$

Consequently we have

$$\mathbf{E}[\|Df_{h}(F)\|_{\ell^{2}(\mathbb{N})}^{2}] \leq \|f_{h}^{'}\|_{\infty}\mathbf{E}[\|DF\|_{\ell^{2}(\mathbb{N})}^{2}]$$

and  $f_h(F) \in \text{Dom}(D)$ . Since F is centered, by the covariance representation (2.15) and the chain rule of Proposition 2.1 we have

$$E[Ff_{h}(F)] = E\left[\sum_{k\geq 0} E[D_{k}F \mid \mathcal{F}_{k-1}]D_{k}f_{h}(F)\right]$$
  
$$= E\left[\sum_{k\geq 0} E[D_{k}F \mid \mathcal{F}_{k-1}]D_{k}Ff_{h}'(F)\right]$$
  
$$-E\left[\sum_{k\geq 0} E[D_{k}F \mid \mathcal{F}_{k-1}]\frac{|D_{k}F|^{2}}{4\sqrt{p_{k}q_{k}}}(f_{h}''(F_{k}^{+}) + f_{h}''(F_{k}^{-}))X_{k}\right]$$
  
$$+E\left[\sum_{k\geq 0} E[D_{k}F \mid \mathcal{F}_{k-1}]R_{k}^{F}(h)\right].$$
 (3.6)

Note that the three expectations in (3.6) are finite. The first one since  $DF \in L^2(\Omega \times \mathbb{N})$  and  $f'_h$  is bounded, indeed by Jensen's inequality

$$\mathbb{E}\left[\left|\sum_{k\geq 0}\mathbb{E}[D_kF \mid \mathcal{F}_{k-1}]D_kFf_h'(F)\right|\right] \leq 4\|h\|_{\infty}\mathbb{E}\left[\sum_{k\geq 0}\mathbb{E}[|D_kF| \mid \mathcal{F}_{k-1}]|D_kF|\right]$$
$$= 4\|h\|_{\infty}$$

$$\times \mathbf{E} \left[ \sum_{k \ge 0} \mathbf{E}[\mathbf{E}[|D_k F| \mid \mathcal{F}_{k-1}] | D_k F| \mid \mathcal{F}_{k-1}] \right]$$

$$= 4 \|h\|_{\infty} \mathbf{E} \left[ \sum_{k \ge 0} \mathbf{E}[\mathbf{E}[|D_k F| \mid \mathcal{F}_{k-1}]^2] \right]$$

$$\le 4 \|h\|_{\infty} \mathbf{E} \left[ \sum_{k \ge 0} \mathbf{E}[\mathbf{E}[|D_k F|^2 \mid \mathcal{F}_{k-1}] \right]$$

$$= 4 \|h\|_{\infty} \sum_{k \ge 0} \mathbf{E}[|D_k F|^2] < \infty;$$

the second relation follows from the boundedness of  $f''_h$  and (3.3), while the third one follows from (2.4) and (3.4). The random variables  $E[D_kF | \mathcal{F}_{k-1}]$ ,  $D_kF$ ,  $F_k^{\pm}$ are independent of  $X_k$  (the first one because it is  $\mathcal{F}_{k-1}$ -measurable and the random variables  $(X_k)_{k\in\mathbb{N}}$  are independent, the others by their definition). Therefore, the equality (3.6) reduces to

$$\begin{split} \mathbf{E}[Ff_{h}(F)] &= \mathbf{E}\left[f_{h}'(F)\sum_{k\geq 0}\mathbf{E}[D_{k}F \mid \mathcal{F}_{k-1}]D_{k}F\right] \\ &+ \sum_{k\geq 0}\frac{1-2p_{k}}{4\sqrt{p_{k}q_{k}}}\mathbf{E}[\mathbf{E}[D_{k}F \mid \mathcal{F}_{k-1}]|D_{k}F|^{2}(f_{h}^{''}(F_{k}^{+}) + f_{h}^{''}(F_{k}^{-}))] \\ &+ \mathbf{E}\left[\sum_{k\geq 0}\mathbf{E}[D_{k}F \mid \mathcal{F}_{k-1}]R_{k}^{F}(h)\right]. \end{split}$$

Inserting this expression into the right-hand side of (3.1) we deduce

$$|\mathbf{E}[h(F)] - \mathbf{E}[h(Z)]| \leq B_1 \min\{4||h||_{\infty}, ||h''||_{\infty}\} + ||h'||_{\infty}B_2$$
(3.7)

$$+ \mathbf{E}\left[\sum_{k\geq 0} |\mathbf{E}[D_k F \mid \mathcal{F}_{k-1}]| |R_k^F(h)|\right], \qquad (3.8)$$

where to get the term (3.7) we used the inequalities  $||f'_h||_{\infty} \leq \min\{4||h||_{\infty}, ||h''||_{\infty}\}$ and  $||f''_h||_{\infty} \leq 2||h'||_{\infty}$  (see Lemma 3.1). Using (2.4) one may easily see that the term in (3.8) is bounded above by  $||h''||_{\infty}B_3$ . The proof is complete.

**Corollary 3.3.** Let  $F \in \text{Dom}(D)$  be a centered random variable and assume that  $B_1: = |1 - ||F||_{L^2(\Omega)}^2 |+ ||\langle D.F, E[D.F | \mathcal{F}_{-1}] \rangle_{\ell^2(\mathbb{N})} - E[\langle D.F, E[D.F | \mathcal{F}_{-1}] \rangle_{\ell^2(\mathbb{N})}]||_{L^2(\Omega)},$ 

$$B_{2}: = \sum_{k \geq 0} \frac{|1 - 2p_{k}|}{\sqrt{p_{k}q_{k}}} \|D_{k}F\|_{L^{2}(\Omega)} \sqrt{\mathbf{E}[|D_{k}F|^{4}]}, \qquad (3.9)$$
$$B_{3}: = \frac{5}{3} \sum_{k \geq 0} \frac{1}{p_{k}q_{k}} \mathbf{E}[|D_{k}F|^{4}]$$

are finite. Then (3.5) holds for all  $h \in \mathbb{C}_b^2$ .

Proof. By the Cauchy-Schwarz and the triangular inequalities we have

$$\mathbb{E}\left[\left|1 - \sum_{k \ge 0} \mathbb{E}[D_k F \mid \mathcal{F}_{k-1}] D_k F\right|\right] \le \left\|1 - \sum_{k \ge 0} \mathbb{E}[D_k F \mid \mathcal{F}_{k-1}] D_k F\right\|_{L^2(\Omega)}$$
  
 
$$\le \|1 - \|F\|_{L^2(\Omega)}^2 + \|\langle D.F, \mathbb{E}[D.F \mid \mathcal{F}_{k-1}] \rangle_{\ell^2(\mathbb{N})} - \|F\|_{L^2(\Omega)}^2 \|_{L^2(\Omega)}.$$

By the Clark-Ocone formula (2.15) we have

$$\mathbf{E}[\langle D.F, \mathbf{E}[D.F \mid \mathcal{F}_{.-1}] \rangle_{\ell^{2}(\mathbb{N})}] = \mathbf{E}\left[\sum_{k=0}^{\infty} D_{k}F \mathbf{E}[D_{k}F \mid \mathcal{F}_{k-1}]\right]$$
$$= \|F\|_{L^{2}(\Omega)}^{2}.$$

Therefore

$$\mathbf{E}\left[\left|1-\sum_{k\geq 0}\mathbf{E}[D_kF\mid \mathcal{F}_{k-1}]D_kF\right|\right]\leq B_1.$$

By the Cauchy-Schwarz and Jensen inequalities we have

$$\begin{split} \mathbf{E}[|\mathbf{E}[D_kF \mid \mathcal{F}_{k-1}]||D_kF|^2] &\leq \sqrt{\mathbf{E}[|\mathbf{E}[D_kF \mid \mathcal{F}_{k-1}]|^2]}\sqrt{\mathbf{E}[|D_kF|^4]} \\ &\leq \sqrt{\mathbf{E}[\mathbf{E}[|D_kF|^2 \mid \mathcal{F}_{k-1}]]}\sqrt{\mathbf{E}[|D_kF|^4]} \\ &= \sqrt{\mathbf{E}[|D_kF|^2]}\sqrt{\mathbf{E}[|D_kF|^4]}, \end{split}$$

and

$$\begin{split} \mathbf{E}[|\mathbf{E}[D_kF \mid \mathcal{F}_{k-1}]||D_kF|^3] &= \sqrt{\mathbf{E}[|\mathbf{E}[D_kF \mid \mathcal{F}_{k-1}]|^2|D_kF|^2]}\sqrt{\mathbf{E}[|D_kF|^4]} \\ &\leq \sqrt{\mathbf{E}[\mathbf{E}[|D_kF|^2 \mid \mathcal{F}_{k-1}]|D_kF|^2]}\sqrt{\mathbf{E}[|D_kF|^4]} \\ &\leq \sqrt{\mathbf{E}[|D_kF|^4]} \times \sqrt{\mathbf{E}[|D_kF|^4]} = \mathbf{E}[|D_kF|^4]. \end{split}$$

The claim follows from Theorem 3.2.

## 3.2. Semigroup bound.

**Theorem 3.4.** Let  $F \in Dom(D)$  be a centered random variable and let

$$C_{1}: = \mathbf{E}\left[\left|1 - \langle DF, -DL^{-1}F \rangle_{\ell^{2}(\mathbb{N})}\right|\right],$$

$$C_{2}: = \sum_{k>0} \frac{|1 - 2p_{k}|}{\sqrt{p_{k}q_{k}}} \mathbf{E}[|D_{k}L^{-1}F||D_{k}F|^{2}],$$
(3.10)

$$C_3: = \frac{5}{3} \sum_{k \ge 0} \frac{1}{p_k q_k} \mathbb{E}[|D_k L^{-1} F| |D_k F|^3]$$
(3.11)

be finite. Then for all  $h \in \mathbb{C}^2_b$  we have

$$|\mathbf{E}[h(F)] - \mathbf{E}[h(Z)]| \le C_1 \min\{4\|h\|_{\infty}, \|h''\|_{\infty}\} + \|h'\|_{\infty}C_2 + \|h''\|_{\infty}C_3.$$
(3.12)

*Proof.* Although the proof is similar to that of Theorem 3.2, we give the details since some points need a different justification. As in the proof of Theorem 3.2 one

has  $f_h(F) \in \text{Dom}(D)$ . Since F is centered, by the covariance representation (2.16) and the chain rule of Proposition 2.1 we have

$$E[Ff_{h}(F)] = -E\left[\sum_{k\geq 0} D_{k}f_{h}(F)D_{k}L^{-1}F\right]$$
  
$$= -E\left[\sum_{k\geq 0} D_{k}Ff_{h}'(F)D_{k}L^{-1}F\right]$$
  
$$+E\left[\sum_{k\geq 0} X_{k}\frac{|D_{k}F|^{2}}{4\sqrt{p_{k}q_{k}}}(f_{h}''(F_{k}^{+}) + f_{h}''(F_{k}^{-}))D_{k}L^{-1}F\right]$$
  
$$-E\left[\sum_{k\geq 0} D_{k}L^{-1}FR_{k}^{F}(h)\right].$$
 (3.13)

Note that the three expectations in (3.13) are finite. The first one since  $DF \in L^2(\Omega \times \mathbb{N})$  and  $f'_h$  is bounded, indeed

$$\begin{split} \mathbf{E}\left[\left|\sum_{k\geq 0} -D_{k}L^{-1}FD_{k}Ff_{h}'(F)\right|\right] &\leq 4\|h\|_{\infty}\mathbf{E}\left[\sum_{k\geq 0} |D_{k}L^{-1}F||D_{k}F|\right] \\ &\leq 4\|h\|_{\infty}\left(\mathbf{E}\sum_{k\geq 0} |D_{k}L^{-1}F|^{2}\right)^{1/2} \\ &\times \left(\mathbf{E}\sum_{k\geq 0} |D_{k}F|^{2}\right)^{1/2} \\ &= 4\|h\|_{\infty}\left(\mathbf{E}\|DL^{-1}F\|_{\ell^{2}(\mathbb{N})}^{2}\right)^{1/2} \left(\mathbf{E}\|DF\|_{\ell^{2}(\mathbb{N})}^{2}\right)^{1/2} \\ &\leq 4\|h\|_{\infty}\mathbf{E}[\|DF\|_{\ell^{2}(\mathbb{N})}^{2}] < \infty, \end{split}$$

where for the latter inequality we used the relation

$$\mathbf{E}[\|DL^{-1}F\|_{\ell^{2}(\mathbb{N})}^{2}] \le \mathbf{E}[\|DF\|_{\ell^{2}(\mathbb{N})}^{2}]$$

(see Lemma 2.13(3) in Nourdin et al. (2010)); the second one due to the boundedness of  $f''_h$  and (3.10); the third one due to (2.4) and (3.11). By Lemma 2.13 in Nourdin et al. (2010) we have that the random variables  $D_k L^{-1}F$ ,  $D_k F$  and  $F_k^{\pm}$ are independent of  $X_k$ . Therefore, the equality (3.13) reduces to

$$\begin{split} \mathbf{E}[Ff_{h}(F)] &= -\mathbf{E}\left[\sum_{k\geq 0} f_{h}'(F) D_{k}F D_{k}L^{-1}F\right] \\ &+ \sum_{k\geq 0} \frac{1-2p_{k}}{4\sqrt{p_{k}q_{k}}} \mathbf{E}[|D_{k}F|^{2}(f''(F_{k}^{+}) + f''(F_{k}^{-}))D_{k}L^{-1}F] \\ &- \mathbf{E}\left[\sum_{k\geq 0} R_{k}^{F}(h)D_{k}L^{-1}F\right]. \end{split}$$

Inserting this expression into the right-hand side of (3.1) we deduce

$$|\mathbf{E}[h(F)] - \mathbf{E}[h(Z)]| \le C_1 \min\{4||h||_{\infty}, ||h''||_{\infty}\} + ||h'||_{\infty}C_2$$
(3.14)

+ E 
$$\left[ \sum_{k \ge 0} |D_k L^{-1} F| |R_k^F(h)| \right]$$
, (3.15)

where to get the term (3.14) we used the inequalities  $||f'_h||_{\infty} \leq \min\{4||h||_{\infty}, ||h''||_{\infty}\}$ and  $||f''_h||_{\infty} \leq 2||h'||_{\infty}$  (see Lemma 3.1). Using (2.4) one may easily see that the term in (3.15) is bounded above by  $||h''||_{\infty}C_3$ . The proof is complete.  $\Box$ 

Note that, formally, the upper bound (3.5) may be obtained by (3.12) substituting the term  $-D_k L^{-1}F$  in the definitions of  $C_1$ ,  $C_2$ ,  $C_3$ , with  $E[D_k F | \mathcal{F}_{k-1}]$ , and vice versa.

**Corollary 3.5.** Let  $F \in Dom(D)$  be a centered random variable and let

$$C_{1}: = |1 - ||F||_{L^{2}(\Omega)}^{2}| + ||\langle D.F, -D.L^{-1}F\rangle_{\ell^{2}(\mathbb{N})} - \mathbb{E}[\langle D.F, -D.L^{-1}F\rangle_{\ell^{2}(\mathbb{N})}]||_{L^{2}(\Omega)},$$

 $C_2$ : =  $B_2$ , where  $B_2$  is defined by (3.9)

and  $C_3$  defined by (3.11) be finite. Then (3.12) holds for all  $h \in C_b^2$ . *Proof.* By the Cauchy-Schwarz and the triangular inequalities we have

$$\mathbb{E}\left[\left|1 - \langle D.F, -D.L^{-1}F \rangle_{\ell^{2}(\mathbb{N})}\right|\right] \leq \left\|1 - \langle D.F, -D.L^{-1}F \rangle_{\ell^{2}(\mathbb{N})}\right\|_{L^{2}(\Omega)} \\ \leq |1 - \|F\|_{L^{2}(\Omega)}^{2}| + \|\langle D.F, -D.L^{-1}F \rangle_{\ell^{2}(\mathbb{N})} - \|F\|_{L^{2}(\Omega)}^{2}\|_{L^{2}(\Omega)}.$$

By the covariance representation formula (2.16) we have

$$\mathbb{E}[\langle D.F, -D.L^{-1}F \rangle_{\ell^{2}(\mathbb{N})}] = ||F||_{L^{2}(\Omega)}^{2}.$$

Therefore

$$\mathbf{E}\left[\left|1-\langle D.F,-D.L^{-1}F\rangle_{\ell^{2}(\mathbb{N})}\right|\right] \leq C_{1}.$$

Let  $F \in \text{Dom}(D)$  be of the form

$$F = \sum_{n \ge 0} J_n(f_n), \quad f_n \in \ell^2_{\mathfrak{s}}(\Delta_n).$$

Then

$$-D_k L^{-1} F = \sum_{n \ge 1} J_{n-1}(f_n(*,k)) \quad \text{and} \quad D_k F = \sum_{n \ge 1} n J_{n-1}(f_n(*,k)).$$

So, by the isometry formula, we have

$$E[|D_k L^{-1} F|^2] = E\left[\left|\sum_{n\geq 1} J_{n-1}(f_n(*,k))\right|^2\right]$$
$$= \sum_{n\geq 1} E[|J_{n-1}(f_n(*,k))|^2]$$
$$= \sum_{n\geq 1} (n-1)! ||f_n(*,k)||^2_{\ell^2(\mathbb{N})^{\otimes (n-1)}}$$

and

$$E[|D_k F|^2] = E\left[\left|\sum_{n\geq 1} nJ_{n-1}(f_n(*,k))\right|^2\right]$$
$$= \sum_{n\geq 1} n^2 E[|J_{n-1}(f_n(*,k))|^2]$$
$$= \sum_{n\geq 1} n^2(n-1)! \|f_n(*,k)\|_{\ell^2(\mathbb{N})^{\otimes (n-1)}}^2.$$

 $\operatorname{So}$ 

$$\mathbb{E}[|D_k L^{-1} F|^2] \le \mathbb{E}[|D_k F|^2]$$
(3.16)

and by the Cauchy-Schwarz inequality, we deduce

$$E[|D_k L^{-1} F||D_k F|^2] \leq \sqrt{E[|D_k L^{-1} F|^2]} \sqrt{E[|D_k F|^4]} \\ \leq \sqrt{E[|D_k F|^2]} \sqrt{E[|D_k F|^4]}.$$

The claim follows from Theorem 3.4.

3.3. Fortet-Mourier distance. In this section we provide bounds in the Fortet-Mourier distance (3.2).

**Theorem 3.6.** Let  $F \in \text{Dom}(D)$  be centered. We have: (i) If (3.5) holds for any  $h \in \mathcal{C}_b^2$  and  $B_1 + B_3 \leq (5 + \mathbb{E}[|F|])/4$ , then

$$d_{\rm FM}(F,Z) \le \sqrt{2(B_1 + B_3)(5 + {\rm E}[|F|])} + B_2.$$
 (3.17)

(ii) If (3.12) holds for any  $h \in \mathbb{C}_b^2$  and  $C_1 + C_3 \leq (5 + \mathbb{E}[|F|])/4$ , then

$$d_{\rm FM}(F,Z) \le \sqrt{2(C_1 + C_3)(5 + {\rm E}[|F|])} + C_2.$$
 (3.18)

*Proof.* We only give the details for the proof of (3.17). The inequality (3.18) can be proved similarly. Take  $h \in \mathcal{FM}$  and define

$$h_t(x) = \int_{\mathbb{R}} h(\sqrt{t}y + \sqrt{1-t}x)\phi(y) \,\mathrm{d}y, \qquad t \in [0,1],$$

where  $\phi$  is the density of the standard  $\mathcal{N}(0,1)$  normal random variable Z. As in the proof of Corollary 3.6 in Nourdin et al. (2010), for  $0 < t \leq 1/2$ , one has  $h_t \in \mathcal{C}_b^2$ and the bounds

$$\|h_t''\|_{\infty} \le 1/\sqrt{t},$$
 (3.19)

and

$$|\mathbf{E}[h(F)] - \mathbf{E}[h_t(F)]| \le \sqrt{t} \left(1 + \frac{\mathbf{E}[|F|]}{2}\right), \qquad |\mathbf{E}[h(Z)] - \mathbf{E}[h_t(Z)]| \le \frac{3}{2}\sqrt{t}.$$

 $\operatorname{So}$ 

$$\begin{split} |\mathbf{E}[h(F)] - \mathbf{E}[h(Z)]| &= |(\mathbf{E}[h(F)] - \mathbf{E}[h_t(F)]) + (\mathbf{E}[h_t(F)] - \mathbf{E}[h_t(Z)]) \\ &+ (\mathbf{E}[h_t(Z)] - \mathbf{E}[h(Z)])| \\ &\leq |\mathbf{E}[h(F)] - \mathbf{E}[h_t(F)]| + |\mathbf{E}[h_t(F)] - \mathbf{E}[h_t(Z)]| \\ &+ |\mathbf{E}[h_t(Z)] - \mathbf{E}[h(Z)]| \\ &\leq \sqrt{t} \left(1 + \frac{\mathbf{E}[|F|]}{2}\right) + B_1 \min\{4\|h_t\|_{\infty}, \|h_t''\|_{\infty}\} + \|h_t^{'}\|_{\infty}B_2 \end{split}$$

$$+ \|h_t''\|_{\infty} B_3 + \frac{3}{2}\sqrt{t} \\ \leq \sqrt{t} \left(\frac{5 + \mathrm{E}[|F|]}{2}\right) + \frac{B_1 + B_3}{\sqrt{t}} + B_2, \qquad (3.20)$$

where in the latter inequality we used (3.19) and that  $||h'_t||_{\infty} \leq 1$ , for all t. Minimizing in  $t \in (0, 1/2]$  the term in (3.20), we have that the optimal is attained at  $t^* = 2(B_1 + B_3)/(5 + \mathbb{E}[|F|]) \in (0, 1/2]$ . The conclusion follows substituting  $t^*$  in (3.20) and then taking the supremum over all the  $h \in \mathcal{FM}$ .

#### 4. First chaos bound for the normal approximation

In this section we specialize the results of Section 3 to first order discrete stochastic integrals. As we shall see, the bounds (3.5) and (3.12) (and the corresponding assumptions) coincide on the first chaos, although they differ on *n*-chaoses,  $n \ge 2$ .

**Corollary 4.1.** Assume that  $\alpha = (\alpha_k)_{k \ge 0}$  is in  $\ell^2(\mathbb{N})$ ,

$$\sum_{k\geq 0} \frac{|1-2p_k|}{\sqrt{p_k q_k}} |\alpha_k|^3 < \infty \quad and \quad \sum_{k\geq 0} \frac{1}{p_k q_k} |\alpha_k|^4 < \infty.$$

$$\tag{4.1}$$

Then for the first chaos

$$F = J_1(\alpha) = \sum_{k \ge 0} \alpha_k Y_k$$

the bound (3.5) (which in this case coincides with the bound (3.12)) holds with

$$B_1 = C_1 = \left| 1 - \sum_{k \ge 0} |\alpha_k|^2 \right|, \qquad B_2 = C_2 = \sum_{k \ge 0} \frac{|1 - 2p_k|}{\sqrt{p_k q_k}} |\alpha_k|^3,$$

and

$$B_3 = C_3 = \frac{5}{3} \sum_{k \ge 0} \frac{1}{p_k q_k} |\alpha_k|^4.$$

*Proof.* Since  $\alpha \in \ell^2(\mathbb{N})$  we have that  $F \in L^2(\Omega)$ . Moreover F is centered, and since

$$D_k F = \alpha_k \sqrt{p_k q_k} \left( \frac{q_k - p_k + 1}{2\sqrt{p_k q_k}} - \frac{q_k - p_k - 1}{2\sqrt{p_k q_k}} \right) = \alpha_k$$

we have  $F \in \text{Dom}(D)$ . The finiteness of the corresponding quantities  $B_1$ ,  $B_2$  and  $B_3$  is guaranteed by  $\alpha \in \ell^2(\mathbb{N})$  and (4.1). The claim follows from e.g. Theorem 3.2.

## Example

Consider the sequence of functionals  $(F_n)_{n\geq 1}$  defined by

$$F_n = \frac{1}{\sqrt{n}} \sum_{k=0}^{n-1} Y_k.$$

Setting

$$\alpha_k = \frac{1}{\sqrt{n}}, \quad k = 0, \dots, n-1, \text{ and } \quad \alpha_k = 0, \quad k \ge n,$$

we have  $B_1 = 0$  and

$$B_2 = \frac{1}{n^{3/2}} \sum_{k=0}^{n-1} \frac{|1-2p_k|}{\sqrt{p_k q_k}} \quad \text{and} \quad B_3 = \frac{5}{3n^2} \sum_{k=0}^{n-1} \frac{1}{p_k q_k}$$

In the symmetric case  $p_k = q_k = 1/2$  we find  $B_2 = 0$  and the bound is of order 1/n, implying a faster rate than in the classical Berry-Esséen estimate (however here we are using  $C_b^2$  test functions; cf. the comment after Corollary 3.3 in Nourdin et al. (2010)).

In the non-symmetric case  $p_k = p$  and  $q_k = q$ ,  $p \neq q$ , the bound is of order  $n^{-1/2}$  as in the classical Berry-Esséen estimate. Indeed we have

$$B_2 = B_2^{(n)} = \frac{1}{\sqrt{n}} \frac{|1 - 2p|}{\sqrt{p(1 - p)}}$$
 and  $B_3 = B_3^{(n)} = \frac{5}{3n} \frac{1}{p(1 - p)}$ 

hence the inequality  $B_1 + B_3 \leq (5 + E[|F_n|])/4$  of Theorem 3.6 reads

$$\frac{5}{3p(1-p)}\frac{1}{\sqrt{n}} \le \frac{5}{4}\sqrt{n} + 4^{-1}\mathrm{E}\left[\left|n\left(\frac{1-2p}{2\sqrt{p(1-p)}}\right) + \sum_{k=0}^{n-1}X_k\right|\right],$$

which holds if e.g.  $n \ge \frac{4}{3p(1-p)}$ . Consequently, by (3.17) it follows that for any  $n \ge \frac{4}{3p(1-p)}$  we have

$$d_{\rm FM}(F_n, Z) \leq \sqrt{2B_3^{(n)}(5 + {\rm E}[|F_n|]) + B_2^{(n)}} \\ = \sqrt{\frac{50}{3p(1-p)}\frac{1}{n} + \frac{10}{3p(1-p)}\frac{1}{n} {\rm E}\left[\left|\frac{1}{\sqrt{n}}\sum_{k=0}^{n-1}\left(\frac{1-2p+X_k}{2\sqrt{p(1-p)}}\right)\right|\right]} \\ + \frac{|1-2p|}{\sqrt{p(1-p)}}\frac{1}{\sqrt{n}} \\ \leq \sqrt{\frac{50}{3p(1-p)}\frac{1}{n} + \frac{10}{3p(1-p)}\frac{1}{n} {\rm E}\left[\left|\frac{1}{\sqrt{n}}\sum_{k=0}^{n-1}\left(\frac{1-2p+X_k}{2\sqrt{p(1-p)}}\right)\right|^2\right]^{1/2}} \\ + \frac{|1-2p|}{\sqrt{p(1-p)}}\frac{1}{\sqrt{n}}.$$

A straightforward computation gives

$$\mathbf{E}\left[\left|\frac{1}{\sqrt{n}}\sum_{k=0}^{n-1}\left(\frac{1-2p+X_k}{2\sqrt{p(1-p)}}\right)\right|^2\right] = 1,$$

hence

$$d_{\rm FM}(F_n, Z) \le \frac{1}{\sqrt{n}} K_1(p), \qquad n \ge \frac{4}{3p(1-p)}$$
 (4.2)

where

$$K_1(p) := \frac{2\sqrt{5} + |1 - 2p|}{\sqrt{p(1 - p)}}$$

In the general case, if

$$a_n := \frac{1}{n^{3/2}} \sum_{k=0}^{n-1} \frac{|1-2p_k|}{\sqrt{p_k q_k}} \to 0 \quad \text{and} \quad b_n := \frac{1}{n^2} \sum_{k=0}^{n-1} \frac{1}{p_k q_k} \to 0, \quad \text{as } n \to \infty, \quad (4.3)$$

then  $F_n \to Z$  in distribution, and the rate depends on the rate of convergence to zero of the sequences  $(a_n)_{n\geq 1}$  and  $(b_n)_{n\geq 1}$ . For instance, if  $p_k = (k+2)^{-\alpha}$ ,  $0 < \alpha < 1, k \geq 0$ , we have  $p_k q_k \geq (n+1)^{-\alpha}(1-(1/2)^{\alpha}), k = 0, \ldots, n-1$ . Consequently we have

$$\frac{1}{n^{3/2}} \sum_{k=0}^{n-1} \frac{|1-2p_k|}{\sqrt{p_k q_k}} \le (1-(1/2)^{\alpha})^{-1/2} \frac{(n+1)^{\alpha/2}}{n^{3/2}} \sum_{k=0}^{n-1} |1-2p_k| \le \frac{1+2^{-(\alpha-1)}}{(1-(1/2)^{\alpha})^{1/2}} \frac{(n+1)^{\alpha/2}}{n^{1/2}},$$

and

$$\frac{1}{n^2} \sum_{k=0}^{n-1} \frac{1}{p_k q_k} \le (1 - (1/2)^{\alpha})^{-1} n^{-1} (n+1)^{\alpha},$$

which yields a bound of order  $n^{-(1-\alpha)/2}$ .

Finally we note that the bound (4.2) in the non-symmetric case  $p_k = p$  and  $q_k = q$ ,  $p \neq q$ , is consistent with the bound on the Wasserstein distance between  $F_n$  and Z provided by Theorem 2.2 in Chatterjee (2008). Indeed, letting  $d_W$  denote the Wasserstein distance and  $Y'_1$  an independent copy of  $Y_1$ , a simple computation shows that

$$\begin{aligned} d_{\rm FM}(F_n,Z) &\leq d_W(F_n,Z) \\ &\leq \frac{1}{2\sqrt{n}} \left( \sqrt{{\rm E}[|Y_1 - Y_1'|^4] - ({\rm E}[|Y_1 - Y_1'|^2])^2} + {\rm E}[|Y_1|^3] \right) \\ &= \frac{1}{2\sqrt{n}} \left( \sqrt{{\rm E}[|Y_1 - Y_1'|^4] - 4} + {\rm E}[|Y_1|^3] \right) \\ &= \frac{1}{\sqrt{n}} K_2(p), \end{aligned}$$

where

$$K_2(p) := \frac{1}{2}\sqrt{\frac{1}{p^2(1-p)} - 4} + \frac{1+2|1-2p|}{4\sqrt{p(1-p)}}$$

since we have

$$\mathbf{E}[|Y_1|^3] \le \frac{1+2|1-2p|}{2\sqrt{p(1-p)}}$$
 and  $\mathbf{E}[|Y_1-Y_1'|^4] = \frac{1}{p^2(1-p)}$ 

We note that when e.g. p is small it holds  $K_2(p) > K_1(p)$ .

Application to determinantal processes. Let E be a locally compact Hausdorff space with countable basis and  $\mathcal{B}(E)$  the Borel  $\sigma$ -field. We fix a Radon measure  $\lambda$  on  $(E, \mathcal{B}(E))$ . The configuration space  $\Gamma_E$  is the family of non-negative N-valued Radon measures on E. We equip  $\Gamma_E$  with the topology which is generated by the functions  $\Gamma_E \ni \xi \longmapsto \xi(A) \in \mathbb{N}, A \in \mathcal{B}(E)$ , where  $\xi(A)$  denotes the number of points of  $\xi$  in A. The existence and uniqueness of a determinantal process with Hermitian kernel K is due to Macchi (1975) and Soshnikov (2000) and can be summarized as follows (we refer the reader to Blank et al. (1994) for notions of functional analysis). **Theorem 4.2.** Let  $\mathcal{K}$  be a self-adjoint integral operator on  $L^2(E, \lambda)$  with kernel K. Suppose that the spectrum of  $\mathcal{K}$  is contained in [0, 1] and that  $\mathcal{K}$  is locally of traceclass, i.e. for any relatively compact  $\Lambda \subset E$ ,  $\mathcal{K}_{\Lambda} = P_{\Lambda} \mathcal{K} P_{\Lambda}$  is of trace-class (here  $P_{\Lambda}f = f \mathbf{1}_{\Lambda}$  is the orthogonal projection.) Then there exists a unique probability measure  $\mu_K$  on  $\Gamma_E$  with n-th correlation measure

$$\lambda_n(\mathrm{d}x_1,\ldots,\mathrm{d}x_n) = \det(K(x_i,x_j))_{1 \le i,j \le n} \lambda(\mathrm{d}x_1)\ldots\lambda(\mathrm{d}x_n),$$

where  $\det(K(x_i, x_j))_{1 \le i,j \le n}$  is the determinant of the  $n \times n$  matrix with ij-entry  $K(x_i, x_j)$ .

The probability measure  $\mu_K$  is called determinantal process with kernel K.

Given a relatively compact set  $\Lambda \subset E$ , we focus on the random variable  $\xi(\Lambda)$  and recall the following basic result (see e.g. Proposition 2.2 in Shirai (2006)).

**Theorem 4.3.** Let  $\mathcal{K}$  be as in the statement of Theorem 4.2 and denote by  $\kappa_k \in [0,1], k \geq 0$ , the eigenvalues of  $\mathcal{K}_{\Lambda}$ . Under  $\mu_K$  the random variable  $\xi(\Lambda)$  has the same distribution of  $\sum_{k\geq 0} Z_k$ , where  $Z_0, Z_1, \ldots$  are independent random variables such that  $Z_k$  obeys the Bernoulli distribution with mean  $\kappa_k$ , i.e.

$$Z_n = \frac{X_n + 1}{2} \in \{0, 1\}, \qquad n \in \mathbb{N}$$

where the X's take values on  $\{-1, 1\}$  and are independent with  $P(X_n = 1) = \kappa_n$ .

The central limit theorem for the number of points on a relatively compact set of a determinantal process may be obtained in different manners, see Costin and Lebowitz (1995), Shirai (2006) and Soshnikov (2002). In the following we provide an alternate derivation which gives the rate of the normal approximation.

**Corollary 4.4.** Let  $\mathcal{K}$  be as in the statement of Theorem 4.2 and  $(\Lambda_n)_{n\geq 0} \subset E$  be an increasing sequence of relatively compact sets such that

$$\operatorname{Var}_{\mu_{K}}(\xi(\Lambda_{n})) = \sum_{k \ge 0} \kappa_{k}^{(n)}(1 - \kappa_{k}^{(n)}) \to \infty, \quad as \ n \to \infty$$

where  $\kappa_k^{(n)} \in [0,1], k \ge 0$ , are the eigenvalues of  $\mathcal{K}_{\Lambda_n}, n \ge 0$ . Setting

$$F_n = \frac{\xi(\Lambda_n) - \mathcal{E}_{\mu_K}[\xi(\Lambda_n)]}{\sqrt{\operatorname{Var}_{\mu_K}(\xi(\Lambda_n))}},$$

for any  $h \in \mathcal{C}_{h}^{2}$ , we have

$$|\mathbf{E}_{\mu_{K}}[h(F_{n})] - \mathbf{E}[h(Z)]| \le ||h'||_{\infty} B_{2}^{(n)} + ||h''||_{\infty} B_{3}^{(n)}, \quad n \ge 0$$

where

$$B_2^{(n)} = \frac{\sum_{k \ge 0} \kappa_k^{(n)} (1 - \kappa_k^{(n)}) |1 - 2\kappa_k^{(n)}|}{\left(\sum_{k \ge 0} \kappa_k^{(n)} (1 - \kappa_k^{(n)})\right)^{3/2}} \le \frac{1}{\left(\sum_{k \ge 0} \kappa_k^{(n)} (1 - \kappa_k^{(n)})\right)^{1/2}}$$

and

$$B_3^{(n)} = \frac{5}{3} \frac{1}{\sum_{k \ge 0} \kappa_k^{(n)} (1 - \kappa_k^{(n)})}.$$

So we have a bound of order  $[\operatorname{Var}_{\mu_K}(\xi(\Lambda_n))]^{-1/2}$ .

*Proof.* For  $n \ge 0$ , let  $(Z_k^{(n)})_{k\ge 0}$  be a sequence of independent  $\{0, 1\}$ -valued random variables with  $Z_k^{(n)} \sim \operatorname{Be}(\kappa_k^{(n)})$  and  $(Y_k^{(n)})_{k\ge 0}$  defined by

$$Y_k^{(n)} = \frac{Z_k^{(n)} - \kappa_k^{(n)}}{\sqrt{\kappa_k^{(n)}(1 - \kappa_k^{(n)})}}$$

By Theorem 4.3 we have

$$\xi(\Lambda_n) \stackrel{d}{=} \sum_{k \ge 0} Z_k^{(n)} = \sum_{k \ge 0} \sqrt{\kappa_k^{(n)} (1 - \kappa_k^{(n)}) Y_k^{(n)}} + \sum_{k \ge 0} \kappa_k^{(n)},$$

where  $\stackrel{d}{=}$  denotes the equality in distribution. Then

$$F_n = \frac{\xi(\Lambda_n) - \mathbf{E}_{\mu_K}[\xi(\Lambda_n)]}{\sqrt{\mathrm{Var}_{\mu_K}(\xi(\Lambda_n))}} \stackrel{d}{=} \sum_{k \ge 0} \alpha_k^{(n)} Y_k^{(n)},$$

where

$$\alpha_k^{(n)} = \frac{\sqrt{\kappa_k^{(n)}(1-\kappa_k^{(n)})}}{\sqrt{\sum_{k\geq 0} \kappa_k^{(n)}(1-\kappa_k^{(n)})}}.$$

We are going to apply Corollary 4.1. Clearly, for any  $n \ge 0$ , the sequence  $(\alpha_k^{(n)})_{k\ge 0}$  is in  $\ell^2(\mathbb{N})$ . Moreover, for any  $n \ge 0$ ,

$$\sum_{k \ge 0} \frac{|\alpha_k^{(n)}|^3}{\sqrt{\kappa_k^{(n)}(1 - \kappa_k^{(n)})}} = \frac{1}{\sqrt{\operatorname{Var}_{\mu_K}(\xi(\Lambda_n))}} < \infty$$

and

$$\sum_{k\geq 0} \frac{|\alpha_k^{(n)}|^4}{\kappa_k^{(n)}(1-\kappa_k^{(n)})} = \frac{1}{\operatorname{Var}_{\mu_K}(\xi(\Lambda_n))} < \infty.$$

So condition (4.1) is satisfied. Moreover, a straightforward computation gives  $B_1 = B_1^{(n)} = 0$ ,  $B_2 = B_2^{(n)}$  and  $B_3 = B_3^{(n)}$ , and the proof is completed.

## Example

Let  $E = \mathbb{C}$  and  $\lambda$  the standard complex Gaussian measure on  $\mathbb{C}$ , i.e.

$$\lambda(\mathrm{d}z) = \frac{1}{\pi} \mathrm{e}^{-|z|^2} \,\mathrm{d}z,$$

where dz is the Lebesgue measure on  $\mathbb{C}$ . The Ginibre process  $\mu_{\exp}$  is the determinantal process with exponential kernel  $K(z, w) = e^{-z\overline{w}}$ , where  $\overline{z}$  is the complex conjugate of  $z \in \mathbb{C}$ . Let b(O, n) be the complex ball centered at the origin with radius n. By Theorem 1.3 in Shirai (2006) we have

$$\operatorname{Var}_{\mu_{\exp}}(\xi(b(O,n))) = \frac{n}{\pi} \int_0^{4n^2} (1 - x/(4n^2))^{1/2} x^{-1/2} \mathrm{e}^{-x} \, \mathrm{d}x$$
$$\sim \frac{n}{\sqrt{\pi}}, \quad \text{as } n \to \infty.$$

So for the Ginibre process Corollary 4.4 provides a bound of order  $n^{-1/2}$ .

#### 5. nth chaos bounds for the normal approximation

In this section we give explicit upper bounds for the constants  $B_i$  and  $C_i$ , i = 1, 2, 3, involved in (3.5) and (3.12), when  $F = J_n(f_n)$ ,  $f_n \in \ell_s^2(\Delta_n)$ . Our approach is based on the multiplication formula (5.3) below, which extends formula (2.11) in Nourdin et al. (2010) (see the discussion after Proposition 5.1).

Given  $f_n \in \ell^2_{\mathfrak{s}}(\Delta_n)$  and  $g_m \in \ell^2_{\mathfrak{s}}(\Delta_m)$ , the contraction  $f_n \otimes_k^l g_m$ ,  $0 \leq l \leq k$ , is defined to be the function of n + m - k - l variables

$$f_n \otimes_k^l g_m(a_{l+1}, \dots, a_n, b_{k+1}, \dots, b_m) := \varphi(a_{l+1}) \cdots \varphi(a_k) f_n \star_k^l g_m(a_{l+1}, \dots, a_n, b_{k+1}, \dots, b_m),$$

where

$$\varphi(n) = \frac{q_n - p_n}{2\sqrt{p_n q_n}}, \quad n \in \mathbb{N}$$
(5.1)

(cf. the structure equation (2.1)) and

а

$$f_n \star'_k g_m(a_{l+1}, \dots, a_n, b_{k+1}, \dots, b_m) := \sum_{1,\dots,a_l \in \mathbb{N}} f_n(a_1, \dots, a_n) g_m(a_1, \dots, a_k, b_{k+1}, \dots, b_m)$$

is the contraction considered in Nourdin et al. (2010) for the symmetric case, see p. 1707 therein. By convention, we define  $\varphi(a_{l+1})\cdots\varphi(a_k) = 1$  if l = k (even when  $\varphi \equiv 0$ ). Denote by  $f_n \otimes_k^l g_m$ ,  $0 \leq l \leq k$ , the symmetrization of  $f_n \otimes_k^l g_m$ . Then, we shall consider the contraction

$$f_{n} \circ_{k}^{l} g_{m}(a_{l+1}, \dots, a_{n}, b_{k+1}, \dots, b_{m}) :=$$

$$= \mathbb{1}_{\Delta_{n+m-k-l}}(a_{l+1}, \dots, a_{n}, b_{k+1}, \dots, b_{m}) f_{n} \otimes_{k}^{l} g_{m}(a_{l+1}, \dots, a_{n}, b_{k+1}, \dots, b_{m}).$$
(5.2)

Note that in the symmetric case  $p_n = q_n = 1/2$  we have  $f_n \otimes_k^k g_m = f_n \star_k^k g_m$ . However,  $f_n \otimes_k^l g_m = 0$  if l < k and so  $f_n \star_k^l g_m \neq f_n \otimes_k^l g_m$  for l < k. The following multiplication formula holds.

Proposition 5.1. We have the chaos expansion

$$J_n(f_n)J_m(g_m) = \sum_{s=0}^{2(n \wedge m)} J_{n+m-s}(h_{n,m,s}),$$
(5.3)

provided the functions

$$h_{n,m,s} := \sum_{s \le 2i \le 2(s \land n \land m)} i! \binom{n}{i} \binom{m}{i} \binom{i}{s-i} f_n \circ_i^{s-i} g_m$$

belong to  $\ell_{\mathfrak{s}}^2(\Delta_{n+m-s})$ ,  $0 \leq s \leq 2(n \wedge m)$ . Here the symbol  $\sum_{s \leq 2i \leq 2(s \wedge n \wedge m)}$  means that the sum is taken over all the integers *i* in the interval  $[s/2, s \wedge n \wedge m]$ .

Since it is not obvious that formula (5.3) extends the product formula (2.11) in Nourdin et al. (2010), it is worthwhile to explain this point in detail. In the symmetric case  $p_n = q_n = 1/2$ , for any  $f_n \in \ell^2_{\mathfrak{s}}(\Delta_n)$ ,  $g_m \in \ell^2_{\mathfrak{s}}(\Delta_m)$ , we have

 $f_n \circ_i^{s-i} g_m = 0$  if  $s < 2i, 0 \le s \le 2(n \land m)$ . Therefore, for any fixed  $0 \le s \le 2(n \land m)$ we have  $h_{n,m,s} = 0$  if s/2 is not an integer and

$$h_{n,m,s} = (s/2)! \binom{n}{s/2} \binom{m}{s/2} f_n \circ_{s/2}^{s/2} g_m,$$

if s/2 is an integer. Note that if s/2 is an integer, we have

$$\|f_n \circ_{s/2}^{s/2} g_m\|_{\ell^2_{\mathfrak{s}}(\Delta_{n+m-s})} = \|f_n \otimes_{s/2}^{s/2} g_m\|_{\ell^2_{\mathfrak{s}}(\Delta_{n+m-s})}$$
  
 
$$\leq \|f_n \otimes_{s/2}^{s/2} g_m\|_{\ell^2(\Delta_{n+m-s})},$$

where we used the straightforward relation

$$\|\tilde{f}\|_{\ell^2(\mathbb{N})^{\otimes n}} \le \|f\|_{\ell^2(\mathbb{N})^{\otimes n}},\tag{5.4}$$

being  $\tilde{f}$  the symmetrization of f. Therefore, by Lemma 2.4(1) in Nourdin et al. (2010) we have  $f_n \circ_{s/2}^{s/2} g_m \in \ell^2_{\mathfrak{s}}(\Delta_{n+m-s})$ , and so  $h_{n,m,s} \in \ell^2_{\mathfrak{s}}(\Delta_{n+m-s})$ , for any  $0 \leq s \leq 2(n \wedge m)$ . By (5.3) we have

$$J_{n}(f_{n})J_{m}(g_{m}) = \sum_{s=0}^{2(n \wedge m)} J_{n+m-s}(h_{n,m,s})$$
  
=  $\sum_{s=0}^{2(n \wedge m)} (s/2)! \binom{n}{s/2} \binom{m}{s/2} J_{n+m-s}(f_{n} \circ_{s/2}^{s/2} g_{m})$   
=  $\sum_{r=0}^{n \wedge m} r! \binom{n}{r} \binom{m}{r} J_{n+m-2r}(f_{n} \circ_{r}^{r} g_{m}),$ 

which is exactly formula (2.11) in Nourdin et al. (2010).

We conclude this part with the following lemma.

**Lemma 5.2.** For any  $f_n \in \ell^2_{\mathfrak{s}}(\Delta_n)$ ,  $g_m \in \ell^2_{\mathfrak{s}}(\Delta_m)$ , we have

$$\sum_{q\geq 0} f_n(*,q) \otimes_k^l g_m(*,q) = f_n \otimes_{k+1}^{l+1} g_m.$$

*Proof.* Note that

$$f_n(*,q) \otimes_k^l g_m(*,q)(a_{l+1},\ldots,a_{n-1},b_{k+1},\ldots,b_{m-1}) = \varphi(a_{l+1})\ldots\varphi(a_k) \sum_{a_1,\ldots,a_l\in\mathbb{N}} f_n(a_1,\ldots,a_{n-1},q)g_m(a_1,\ldots,a_k,b_{k+1},\ldots,b_{m-1},q),$$

and so summing up over  $q \in \mathbb{N}$  we deduce

$$\sum_{q \ge 0} f_n(*,q) \otimes_k^l g_m(*,q)(a_{l+1}, \dots, a_{n-1}, b_{k+1}, \dots, b_{m-1})$$
  
=  $\varphi(a_{l+1}) \dots \varphi(a_k) \sum_{a_1, \dots, a_l, q \in \mathbb{N}} f_n(a_1, \dots, a_{n-1}, q) g_m(a_1, \dots, a_k, b_{k+1}, \dots, b_{m-1}, q)$   
=  $\varphi(a_{l+1}) \dots \varphi(a_k) f_n \star_{k+1}^{l+1} g_m(a_{l+1}, \dots, a_{n-1}, b_{k+1}, \dots, b_{m-1})$   
=  $f_n \otimes_{k+1}^{l+1} g_m(a_{l+1}, \dots, a_{n-1}, b_{k+1}, \dots, b_{m-1}).$ 

5.1. Clark-Ocone bound. By e.g. Lemma 4.6 in Privault (2008), for the *n*th-chaos  $J_n(f_n), n \ge 2, f_n \in \ell^2_{\mathfrak{s}}(\Delta_n)$ , we have

$$\mathbb{E}[J_n(f_n) \mid \mathfrak{F}_k] = J_n(f_n \mathbb{1}_{[0,k]^n}), \qquad k \in \mathbb{N}.$$

Therefore

$$E[D_k J_n(f_n) \mid \mathcal{F}_{k-1}] = n E[J_{n-1}(f_n(*,k)) \mid \mathcal{F}_{k-1}] = n J_{n-1}(f_n)_k, \quad (5.5)$$

where

$$f_{n]k}(*) := f_n(*,k) \mathbb{1}_{[0,k-1]^{n-1}}(*).$$
(5.6)

So by the isometric properties of discrete multiple stochastic integrals we have that the constants  $B_i$  of Corollary 3.3 are equal, respectively, to

$$\dot{B}_{1} := |1 - n! \| f_{n} \|_{\ell_{\mathfrak{s}}^{2}(\Delta_{n})}^{2} | + n^{2} \| \langle J_{n-1}(f_{n}(\ast, \cdot)), J_{n-1}(f_{n}] \cdot (\ast)) \rangle_{\ell^{2}(\mathbb{N})} 
- \mathbb{E}[ \langle J_{n-1}(f_{n}(\ast, \cdot)), J_{n-1}(f_{n}] \cdot (\ast)) \rangle_{\ell^{2}(\mathbb{N})} ] \|_{L^{2}(\Omega)},$$
(5.7)

$$\tilde{B}_{2} := n^{3} \sqrt{(n-1)!} \sum_{k \ge 0} \frac{|1-2p_{k}|}{\sqrt{p_{k}q_{k}}} \|f_{n}(*,k)\|_{\ell^{2}_{\mathfrak{s}}(\Delta_{n-1})} \sqrt{\mathbf{E}[|J_{n-1}(f_{n}(*,k))|^{4}]}, (5.8)$$

$$\tilde{B}_3 := \frac{5n^4}{3} \sum_{k \ge 0} \frac{1}{p_k q_k} \mathbb{E}[|J_{n-1}(f_n(*,k))|^4].$$
(5.9)

In the proof of the next theorem we show that these constants can be bounded above by computable quantities.

**Theorem 5.3.** Let  $n \ge 2$  be fixed and let  $f_n \in \ell^2_{\mathfrak{s}}(\Delta_n)$ . Assume that for any  $k \in \mathbb{N}$  the functions

$$h_{n-1,n-1,s}^{(k)} := \sum_{s \le 2i \le 2(s \land (n-1))} i! \binom{n-1}{i}^2 \binom{i}{s-i} f_n(*,k) \circ_i^{s-i} f_{n]k}(*)$$
(5.10)

and

$$\tilde{h}_{n-1,n-1,s}^{(k)} := \sum_{s \le 2i \le 2(s \land (n-1))} i! \binom{n-1}{i}^2 \binom{i}{s-i} f_n(*,k) \circ_i^{s-i} f_n(*,k)$$
(5.11)

belong to  $\ell_{\mathfrak{s}}^2(\Delta_{2n-2-s}), \ 0 \leq s \leq 2n-2, \ and \ that$ 

$$B_{1} := |1 - n!| |f_{n}||_{\ell_{s}^{2}(\Delta_{n})}^{2} |$$

$$+ n^{2} \left( \sum_{s=0}^{2n-3} (2n-2-s)! \sum_{s \leq \{2i_{1}, 2i_{2}\} \leq 2(s \wedge (n-1))} i_{1}!i_{2}! {\binom{n-1}{i_{1}}}^{2} \right) \times {\binom{n-1}{i_{2}}}^{2} {\binom{i_{1}}{s-i_{1}}} {\binom{i_{2}}{s-i_{2}}} \sum_{k \geq 0} ||f_{n}(*,k) \otimes_{i_{1}}^{s-i_{1}} f_{n}]_{k}(*)||_{\ell^{2}(\Delta_{2n-2-s})} \times \sum_{k \geq 0} ||f_{n}(*,k) \otimes_{i_{2}}^{s-i_{2}} f_{n}]_{k}(*)||_{\ell^{2}(\Delta_{2n-2-s})} \right)^{1/2},$$

$$B_{2} := n^{3} \sqrt{(n-1)!} \sum_{k \geq 0} \frac{|1 - 2p_{k}|}{\sqrt{p_{k}q_{k}}} ||f_{n}(*,k)||_{\ell_{s}^{2}(\Delta_{n-1})} \left( \sum_{s=0}^{2n-2} (2n-2-s)! \right) \times \sum_{s \leq \{2i_{1}, 2i_{2}\} \leq 2(s \wedge (n-1))} i_{1}!i_{2}! {\binom{n-1}{i_{1}}}^{2} {\binom{n-1}{i_{2}}}^{2} {\binom{i_{1}}{s-i_{1}}} {\binom{i_{2}}{s-i_{2}}}$$

$$\times \|f_n(*,k) \otimes_{i_1}^{s-i_1} f_n(*,k)\|_{\ell^2(\Delta_{2n-2-s})} \\ \times \|f_n(*,k) \otimes_{i_2}^{s-i_2} f_n(*,k)\|_{\ell^2(\Delta_{2n-2-s})} \right)^{1/2},$$
(5.12)

$$B_{3}: = \frac{5n^{4}}{3} \sum_{s=0}^{2n-2} (2n-2-s)! \\ \times \sum_{s \leq \{2i_{1}, 2i_{2}\} \leq 2(s \wedge (n-1))} i_{1}!i_{2}! {\binom{n-1}{i_{1}}}^{2} {\binom{n-1}{i_{2}}}^{2} {\binom{i_{1}}{s-i_{1}}} {\binom{i_{2}}{s-i_{2}}} \\ \times \sum_{k \geq 0} \frac{1}{p_{k}q_{k}} \|f_{n}(*,k) \otimes_{i_{1}}^{s-i_{1}} f_{n}(*,k)\|_{\ell^{2}(\Delta_{2n-2-s})} \\ \times \|f_{n}(*,k) \otimes_{i_{2}}^{s-i_{2}} f_{n}(*,k)\|_{\ell^{2}(\Delta_{2n-2-s})}$$
(5.13)

are finite. Then for all  $h \in \mathbb{C}_b^2$  we have

 $|\mathbf{E}[h(J_n(f_n))] - \mathbf{E}[h(Z)]| \le B_1 \min\{4\|h\|_{\infty}, \|h''\|_{\infty}\} + \|h'\|_{\infty}B_2 + \|h''\|_{\infty}B_3.$ 

*Proof.* The claim follows from Corollary 3.3 if we show that the constants  $\tilde{B}_i$ defined by (5.7), (5.8) and (5.9) are bounded above by the constants  $B_i$  defined in the statement, respectively.

Step 1: Proof of  $\tilde{B}_1 \leq B_1$ . By the hypotheses on the functions  $h_{n-1,n-1,s}^{(k)}$  and the multiplication formula (5.3), we deduce

$$J_{n-1}(f_n(*,k))J_{n-1}(f_{n]k}) = \sum_{s=0}^{2n-2} \sum_{s \le 2i \le 2(s \land (n-1))} i! \binom{n-1}{i}^2 \binom{i}{s-i} J_{2n-2-s}(f_n(*,k) \circ_i^{s-i} f_{n]k}(*))$$
  
$$= (n-1)! f_n(*,k) \circ_{n-1}^{n-1} f_{n]k}(*) + \sum_{s=0}^{2n-3} \sum_{s \le 2i \le 2(s \land (n-1))} i! \binom{n-1}{i}^2 \binom{i}{s-i} J_{2n-2-s}(f_n(*,k) \circ_i^{s-i} f_{n]k}(*)).$$
  
(5.14)

Since the constant  $B_1$  in the statement is finite, we have

$$\sum_{k\geq 0} \|f_n(*,k) \otimes_i^{s-i} f_{n]k}(*)\|_{\ell^2(\Delta_{2n-2-s})} < \infty,$$

$$0 \le s \le 2n-3, \ s \le 2i \le 2(s \land (n-1)). \text{ By } (5.4) \text{ this in turn implies}$$
$$\sum_{k \ge 0} \|f_n(*,k) \circ_i^{s-i} f_{n]k}(*)\|_{\ell^2_s(\Delta_{2n-2-s})} < \infty,$$

 $0 \le s \le 2n - 3$ ,  $s \le 2i \le 2(s \land (n - 1))$ , and so

$$\sum_{k\geq 0} f_n(*,k) \circ_i^{s-i} f_{n]k}(*) \in \ell^2_{\mathfrak{s}}(\Delta_{2n-2-s}),$$

 $0 \le s \le 2n-3, s \le 2i \le 2(s \land (n-1))$  (it is worthwhile to note that one can not use Lemma 5.2 to express the infinite sum  $\sum_{k\ge 0} f_n(*,k) \circ_i^{s-i} f_{n]k}(*)$  since the function

of n variables  $f_{n].}(*)$  is not symmetric). Therefore, summing up over  $k \ge 0$  in the equality (5.14), we get

$$\begin{aligned} \langle J_{n-1}(f_n(*,\cdot)), J_{n-1}(f_n](*)) \rangle_{\ell^2(\mathbb{N})} \\ &= (n-1)! \sum_{k \ge 0} f_n(*,k) \circ_{n-1}^{n-1} f_{n]k}(*) \\ &+ \sum_{s=0}^{2n-3} \sum_{s \le 2i \le 2(s \land (n-1))} i! \binom{n-1}{i}^2 \binom{i}{s-i} J_{2n-2-s} \left( \sum_{k \ge 0} f_n(*,k) \circ_i^{s-i} f_{n]k}(*) \right). \end{aligned}$$

Taking the mean and noticing that discrete multiple stochastic integrals are centered, we have

$$\mathbb{E}[\langle J_{n-1}(f_n(*,\cdot)), J_{n-1}(f_{n]}(*)) \rangle_{\ell^2(\mathbb{N})}] = (n-1)! \sum_{k \ge 0} f_n(*,k) \circ_{n-1}^{n-1} f_{n]k}(*),$$

and so

$$\langle J_{n-1}(f_n(*,\cdot)), J_{n-1}(f_n].(*)) \rangle_{\ell^2(\mathbb{N})} - \mathbb{E}[\langle J_{n-1}(f_n(*,\cdot)), J_{n-1}(f_n].(*)) \rangle_{\ell^2(\mathbb{N})}$$

$$= \sum_{s=0}^{2n-3} \sum_{s \le 2i \le 2(s \land (n-1))} i! \binom{n-1}{i}^2 \binom{i}{s-i} J_{2n-2-s} \left( \sum_{k \ge 0} f_n(*,k) \circ_i^{s-i} f_n]_k(*) \right).$$

By means of the orthogonality and isometric properties of discrete multiple stochastic integrals, we have

$$\begin{split} & \mathbf{E}\left[ \left( \sum_{s=0}^{2n-3} \sum_{s \leq 2i \leq 2(s \wedge (n-1))} i! \binom{n-1}{i}^2 \binom{i}{s-i} J_{2n-2-s} \left( \sum_{k \geq 0} f_n(*,k) \circ_i^{s-i} f_{n]k}(*) \right) \right)^2 \right] \\ &= \sum_{s=0}^{2n-3} \mathbf{E}\left[ \left( \sum_{s \leq 2i \leq 2(s \wedge (n-1))} i! \binom{n-1}{i}^2 \binom{i}{s-i} \right) \\ & \times J_{2n-2-s} \left( \sum_{k \geq 0} f_n(*,k) \circ_i^{s-i} f_{n]k}(*) \right) \right)^2 \right] \\ &+ \sum_{s_1 \neq s_2}^{0,2n-3} \mathbf{E}\left[ \left( \sum_{s_1 \leq 2i \leq 2(s_1 \wedge (n-1))} i! \binom{n-1}{i}^2 \binom{i}{s_1-i} \right) \\ & \times J_{2n-2-s_1} \left( \sum_{k \geq 0} f_n(*,k) \circ_i^{s_1-i} f_{n]k}(*) \right) \right) \\ & \times \left( \sum_{s_2 \leq 2i \leq 2(s_2 \wedge (n-1))} i! \binom{n-1}{i}^2 \binom{i}{s_2-i} \\ & \times J_{2n-2-s_2} \left( \sum_{k \geq 0} f_n(*,k) \circ_i^{s_2-i} f_{n]k}(*) \right) \right) \right] \end{split}$$

$$= \sum_{s=0}^{2n-3} \mathbb{E} \left[ \left( \sum_{s \le 2i \le 2(s \land (n-1))} i! \binom{n-1}{i}^2 \binom{i}{s-i} \right) \times J_{2n-2-s} \left( \sum_{k \ge 0} f_n(*,k) \circ_i^{s-i} f_{n]k}(*) \right) \right)^2 \right]$$

$$= \sum_{s=0}^{2n-3} \sum_{s \le \{2i_1, 2i_2\} \le 2(s \land (n-1))} \mathbb{E} \left[ i_1! \binom{n-1}{i_1}^2 \binom{i_1}{s-i_1} i_2! \binom{n-1}{i_2}^2 \binom{i_2}{s-i_2} \times J_{2n-2-s} \left( \sum_{k \ge 0} f_n(*,k) \circ_{i_1}^{s-i_1} f_{n]k}(*) \right) \times J_{2n-2-s} \left( \sum_{k \ge 0} f_n(*,k) \circ_{i_2}^{s-i_2} f_{n]k}(*) \right) \right]$$

$$= \sum_{s=0}^{2n-3} (2n-2-s)! \sum_{s \le \{2i_1, 2i_2\} \le 2(s \land (n-1))} i_1! i_2! \binom{n-1}{i_1}^2 \binom{n-1}{i_2}^2 \binom{i_1}{s-i_1} \times \binom{i_2}{s-i_2} \times \binom{i_2}{s-i_2} \left( \sum_{k \ge 0} f_n(*,k) \circ_{i_2}^{s-i_2} f_{n]k}(*) \right) \right]$$

$$\times \binom{i_2}{s-i_2} \langle \sum_{k \ge 0} f_n(*,k) \circ_{i_1}^{s-i_1} f_{n]k}(*), \sum_{k \ge 0} f_n(*,k) \circ_{i_2}^{s-i_2} f_{n]k}(*) \rangle \ell_s^2 (\Delta_{2n-2-s}).$$
(5.15)

By the above relations and (5.7), we deduce

$$\tilde{B}_{1} = |1 - n! \| f_{n} \|_{\ell_{s}^{2}(\Delta_{n})}^{2} | + n^{2} \left( \sum_{s=0}^{2n-3} (2n-2-s)! \sum_{s \leq \{2i_{1}, 2i_{2}\} \leq 2(s \wedge (n-1))} i_{1}! i_{2}! \right) \\ \times \left( \frac{n-1}{i_{1}} \right)^{2} \binom{n-1}{i_{2}}^{2} \binom{i_{1}}{s-i_{1}} \binom{i_{2}}{s-i_{2}} \\ \times \left( \sum_{k \geq 0} f_{n}(*,k) \circ_{i_{1}}^{s-i_{1}} f_{n}]_{k}(*), \sum_{k \geq 0} f_{n}(*,k) \circ_{i_{2}}^{s-i_{2}} f_{n}]_{k}(*) \right) \ell_{s}^{2}(\Delta_{2n-2-s}) \right)^{1/2}.$$

$$(5.16)$$

Now, note that by the Cauchy-Schwarz inequality

$$|\langle f,g\rangle_{\ell^2(\mathbb{N})^{\otimes n}}| \le ||f||_{\ell^2(\mathbb{N})^{\otimes n}} ||g||_{\ell^2(\mathbb{N})^{\otimes n}}, \quad \text{for any } f,g \in \ell^2(\mathbb{N})^{\otimes n}.$$
(5.17)

By this relation, (5.4) and (5.16) we easily get  $\tilde{B}_1 \leq B_1$ . Step 2: Proof of  $\tilde{B}_i \leq B_i$ , i = 2, 3. By the hypotheses on the functions  $\tilde{h}_{n-1,n-1,s}^{(k)}$  and the multiplication formula (5.3), we deduce

$$J_{n-1}(f_n(*,k))^2 = \sum_{s=0}^{2n-2} \sum_{s \le 2i \le 2(s \land (n-1))} i! \binom{n-1}{i}^2 \binom{i}{s-i} J_{2n-2-s}(f_n(*,k) \circ_i^{s-i} f_n(*,k)).$$

By a similar computation as for (5.15), we have

$$E[|J_{n-1}(f_n(*,k))|^4] = E\left[\left(\sum_{s=0}^{2n-2} \sum_{s \le 2i \le 2(s \land (n-1))} i! \binom{n-1}{i}^2 \binom{i}{s-i} + J_{2n-2-s} \left(f_n(*,k) \circ_i^{s-i} f_n(*,k)\right)\right)^2\right]$$
  
$$= \sum_{s=0}^{2n-2} (2n-2-s)! \sum_{s \le \{2i_1, 2i_2\} \le 2(s \land (n-1))} i_1! i_2 \binom{n-1}{i_1}^2 + \binom{n-1}{i_2}^2 \binom{i_1}{s-i_1} \binom{i_2}{s-i_2} + \langle f_n(*,k) \circ_{i_1}^{s-i_1} f_n(*,k), f_n(*,k) \circ_{i_2}^{s-i_2} f_n(*,k) \rangle_{\ell^2_s(\Delta_{2n-2-s})},$$
  
(5.18)

and so by (5.8) and (5.9) we deduce

$$\tilde{B}_{2} = n^{3} \sqrt{(n-1)!} \sum_{k \geq 0} \frac{|1-2p_{k}|}{\sqrt{p_{k}q_{k}}} \|f_{n}(*,k)\|_{\ell^{2}_{s}(\Delta_{n-1})} \left(\sum_{s=0}^{2n-2} (2n-2-s)!\right) \\
\times \sum_{s \leq \{2i_{1}, 2i_{2}\} \leq 2(s \wedge (n-1))} i_{1}! i_{2}! \binom{n-1}{i_{1}}^{2} \binom{n-1}{i_{2}}^{2} \binom{i_{1}}{s-i_{1}} \binom{i_{2}}{s-i_{2}} \\
\times \langle f_{n}(*,k) \circ_{i_{1}}^{s-i_{1}} f_{n}(*,k), f_{n}(*,k) \circ_{i_{2}}^{s-i_{2}} f_{n}(*,k) \rangle_{\ell^{2}_{s}(\Delta_{2n-2-s})} \right)^{1/2} (5.19)$$

and

$$\tilde{B}_{3} = \frac{5n^{4}}{3} \sum_{s=0}^{2n-2} (2n-2-s)! \\
\times \sum_{s \leq \{2i_{1}, 2i_{2}\} \leq 2(s \wedge (n-1))} i_{1}!i_{2}! {\binom{n-1}{i_{1}}}^{2} {\binom{n-1}{i_{2}}}^{2} {\binom{i_{1}}{s-i_{1}}} {\binom{i_{2}}{s-i_{2}}} \\
\sum_{k \geq 0} \frac{1}{p_{k}q_{k}} \langle f_{n}(*,k) \circ_{i_{1}}^{s-i_{1}} f_{n}(*,k), f_{n}(*,k) \circ_{i_{2}}^{s-i_{2}} f_{n}(*,k) \rangle_{\ell_{s}^{2}(\Delta_{2n-2-s})}.$$
(5.20)

The claim follows from the above equalities and relations (5.17) and (5.4).

5.2. Semigroup bound. For the nth-chaos 
$$J_n(f_n), n \ge 2, f_n \in \ell^2_{\mathfrak{s}}(\Delta_n)$$
, we have

$$-D_k L^{-1} J_n(f_n) = n^{-1} D_k J_n(f_n) = J_{n-1}(f_n(*,k))$$
(5.21)

and the constants  ${\cal C}_i$  of Corollary 3.5 are equal, respectively, to

$$\tilde{C}_1 := |1 - n! \| f_n \|_{\ell^2_{\mathfrak{s}}(\Delta_n)}^2 | + n \| \langle J_{n-1}(f_n(*, \cdot)), J_{n-1}(f_n(*, \cdot)) \rangle_{\ell^2(\mathbb{N})}$$

$$- \operatorname{E}[\langle J_{n-1}(f_n(*,\cdot)), J_{n-1}(f_n(*,\cdot)) \rangle_{\ell^2(\mathbb{N})}] \|_{L^2(\Omega)},$$
(5.22)

$$\tilde{C}_2$$
: =  $\tilde{B}_2$ , where  $\tilde{B}_2$  is defined by (5.8)

$$\tilde{C}_3: = \frac{\tilde{B}_3}{n}$$
, where  $\tilde{B}_3$  is defined by (5.9)

In the next theorem we show that these constants can be bounded above by computable quantities.

**Theorem 5.4.** Let  $n \ge 2$  be fixed and let  $f_n \in \ell^2_{\mathfrak{s}}(\Delta_n)$ . Assume that for any  $k \in \mathbb{N}$  the functions  $\tilde{h}_{n-1,n-1,s}^{(k)}$  defined by (5.11) belong to  $\ell^2_{\mathfrak{s}}(\Delta_{2n-2-s})$ ,  $0 \le s \le 2n-2$ , and that

$$C_{1} := |1 - n!| \|f_{n}\|_{\ell_{s}^{2}(\Delta_{n})}^{2} |$$
  
+  $n \left( \sum_{s=0}^{2n-3} (2n-2-s)! \sum_{s \leq \{2i_{1}, 2i_{2}\} \leq 2(s \wedge (n-1))} i_{1}! i_{2}! {\binom{n-1}{i_{1}}}^{2} {\binom{n-1}{i_{2}}}^{2} {\binom{i_{1}}{s-i_{1}}} \right)$   
 $\times {\binom{i_{2}}{s-i_{2}}} \|f_{n} \otimes_{i_{1}+1}^{s-i_{1}+1} f_{n}\|_{\ell^{2}(\Delta_{2n-2-s})} \|f_{n} \otimes_{i_{2}+1}^{s-i_{2}+1} f_{n}\|_{\ell^{2}(\Delta_{2n-2-s})} \right)^{1/2},$ 

 $C_2 := B_2$ , where  $B_2$  is defined by (5.12), and  $C_3 := B_3/n$  where  $B_3$  is defined by (5.13), are finite. Then for all  $h \in \mathcal{C}_b^2$  we have

$$|\mathbf{E}[h(J_n(f_n))] - \mathbf{E}[h(Z)]| \le C_1 \min\{4||h||_{\infty}, ||h''||_{\infty}\} + ||h'||_{\infty}C_2 + ||h''||_{\infty}C_3$$

*Proof.* The claim follows from Corollary 3.5 if we show that the constant  $\tilde{C}_1$  defined by (5.22) is bounded above by the constant  $C_1$  defined in the statement (for the bounds  $\tilde{C}_i \leq C_i$ , i = 2, 3, see Step 2 of the proof of Theorem 5.3). Along a similar computation as in the Step 1 of the proof of Theorem 5.3, we have

$$\langle J_{n-1}(f_n(*,\cdot)), J_{n-1}(f_n(*,\cdot)) \rangle_{\ell^2(\mathbb{N})}$$

$$= (n-1)! \sum_{k\geq 0} f_n(*,k) \circ_{n-1}^{n-1} f_n(*,k) + \sum_{s=0}^{2n-3} \sum_{s\leq 2i\leq 2(s\wedge(n-1))} i! \binom{n-1}{i}^2 \binom{i}{s-i}$$

$$\times J_{2n-2-s} \left( \sum_{k\geq 0} f_n(*,k) \circ_i^{s-i} f_n(*,k) \right)$$

$$= (n-1)! f_n \circ_n^n f_n$$

$$+ \sum_{s=0}^{2n-3} \sum_{s\leq 2i\leq 2(s\wedge(n-1))} i! \binom{n-1}{i}^2 \binom{i}{s-i} J_{2n-2-s} \left( f_n \circ_{i+1}^{s-i+1} f_n \right), \quad (5.23)$$

where the latter equality follows from Lemma 5.2. By a similar computation as for (5.15), we have

$$\| \|J_{n-1}(f_n(*,\cdot))\|_{\ell^2(\mathbb{N})}^2 - \mathbf{E}[\|J_{n-1}(f_n(*,\cdot))\|_{\ell^2(\mathbb{N})}^2] \|_{L^2(\Omega)}^2$$

$$= \mathbf{E}\left[ \left( \sum_{s=0}^{2n-3} \sum_{s \le 2i \le 2(s \land (n-1))} i! \binom{n-1}{i}^2 \binom{i}{s-i} J_{2n-2-s}(f_n \circ_{i+1}^{s-i+1} f_n) \right)^2 \right]$$

$$= \sum_{s=0}^{2n-3} (2n-2-s)! \sum_{s \le \{2i_1, 2i_2\} \le 2(s \land (n-1))} i_1! i_2! {\binom{n-1}{i_1}}^2 {\binom{n-1}{i_2}}^2 {\binom{n-1}{i_2}}^2 {\binom{n}{s-i_1}}^2 {\binom{n}{s-i_1}} {\binom{n}{s-i_1}} {\binom{n}{s-i_1}}^2 {\binom{n}{s$$

By this relation and (5.22) we deduce

$$\begin{split} \bar{C}_1 &= |1 - n! \| f_n \|_{\ell_s^2(\Delta_n)}^2 | \\ &+ n \left( \sum_{s=0}^{2n-3} (2n-2-s)! \sum_{s \le \{2i_1, \ 2i_2\} \le 2(s \land (n-1))} i_1! i_2! \binom{n-1}{i_1}^2 \binom{n-1}{i_2}^2 \binom{i_1}{s-i_1} \right) \\ &\times \binom{i_2}{s-i_2} \langle f_n \circ_{i_1+1}^{s-i_1+1} f_n, f_n \circ_{i_2+1}^{s-i_2+1} f_n \rangle_{\ell_s^2(\Delta_{2n-2-s})} \right)^{1/2}. \end{split}$$

By this equality and (5.17) and (5.4) we finally have  $\tilde{C}_1 \leq C_1$ .

Connection with Theorem 4.1 in Nourdin et al. (2010). In this subsection we refine a little the bound given by Theorem 5.4 in order to strictly extend the bound provided by Theorem 4.1 in Nourdin et al. (2010). For the *n*th chaos  $J_n(f_n)$ ,  $n \geq 2$ ,  $f_n \in \ell_{\mathfrak{s}}^2(\Delta_n)$ , we have that the constants  $C_i$  of Theorem 3.4 are equal, respectively, to

$$\tilde{C}_1 := \mathbb{E}\left[ \|1 - n\| J_{n-1}(f_n(*, \cdot)) \|_{\ell^2(\mathbb{N})}^2 \| \right],$$
(5.25)

$$\tilde{C}_2 := n^2 \sum_{k \ge 0} \frac{|1 - 2p_k|}{\sqrt{p_k q_k}} \mathbf{E}[|J_{n-1}(f_n(*, k))|^3],$$
(5.26)

and

$$\tilde{C}_3 := \frac{\tilde{B}_3}{n}$$
, where  $\tilde{B}_3$  is defined by (5.9)

In the next theorem we show that these constants can be bounded above by computable quantities.

**Theorem 5.5.** Let  $n \ge 2$  be fixed and let  $f_n \in \ell^2_{\mathfrak{s}}(\Delta_n)$ . Assume that for any  $k \in \mathbb{N}$  the functions  $\tilde{h}_{n-1,n-1,s}^{(k)}$  defined by (5.11) belong to  $\ell^2_{\mathfrak{s}}(\Delta_{2n-2-s})$ ,  $0 \le s \le 2n-2$ , and that

$$C_{1} := \left( |1 - n!| \|f_{n}\|_{\ell_{s}^{2}(\Delta_{n})}^{2} |^{2} + n^{2} \sum_{s=0}^{2n-3} (2n-2-s)! \sum_{s \leq \{2i_{1}, 2i_{2}\} \leq 2(s \wedge (n-1))} i_{1}! i_{2}! {\binom{n-1}{i_{1}}}^{2} {\binom{n-1}{i_{2}}}^{2} {\binom{i_{1}}{s-i_{1}}} \right) \\ \times {\binom{i_{2}}{s-i_{2}}} \|f_{n} \otimes_{i_{1}+1}^{s-i_{1}+1} f_{n}\|_{\ell^{2}(\Delta_{2n-2-s})} \|f_{n} \otimes_{i_{2}+1}^{s-i_{2}+1} f_{n}\|_{\ell^{2}(\Delta_{2n-2-s})} \right)^{1/2},$$

 $C_2 := B_2/n$ , where  $B_2$  is defined by (5.12) and  $C_3 := B_3/n$ , where  $B_3$  is defined by (5.13) are finite. Then for all  $h \in \mathbb{C}^2_b$  we have

$$|\mathbf{E}[h(J_n(f_n))] - \mathbf{E}[h(Z)]| \le C_1 \min\{4\|h\|_{\infty}, \|h''\|_{\infty}\} + \|h'\|_{\infty}C_2 + \|h''\|_{\infty}C_3.$$

*Proof.* The claim follows from Theorem 3.4 if we show that the constants  $\tilde{C}_i$ , i = 1, 2, defined by (5.25) and (5.26) are bounded above by the constants  $C_i$ , i = 1, 2, defined in the statement, respectively (for the bound  $\tilde{C}_3 \leq C_3$  see Step 2 of the proof of Theorem 5.3).

Step 1: Proof of  $\tilde{C}_1 \leq C_1$ . By the Cauchy-Schwarz inequality, (5.23) and (5.24) we have

$$\begin{split} \tilde{C}_{1} &\leq \mathbf{E} \left[ |1 - n| \|J_{n-1}(f_{n}(*, \cdot))\|_{\ell^{2}(\mathbb{N})}^{2} |^{2} \right]^{1/2} \\ &\leq \left( |1 - n!| \|f_{n}\|_{\ell^{2}_{s}(\Delta_{n})}^{2} + n^{2} \mathbf{E} \left[ \left( \sum_{s=0}^{2n-3} \sum_{s \leq 2i \leq 2(s \wedge (n-1))} i! \binom{n-1}{i}^{2} \binom{i}{s-i} \right) \right. \\ &\times J_{2n-2-s}(f_{n} \circ_{i+1}^{s-i+1} f_{n}) \right)^{2} \right] \right)^{1/2} \\ &= \left( |1 - n!| \|f_{n}\|_{\ell^{2}_{s}(\Delta_{n})}^{2} \right|^{2} \\ &+ n^{2} \sum_{s=0}^{2n-3} (2n-2-s)! \sum_{s \leq \{2i_{1}, \ 2i_{2}\} \leq 2(s \wedge (n-1))} i_{1}! i_{2}! \binom{n-1}{i_{1}}^{2} \binom{n-1}{i_{2}}^{2} \\ &\times \binom{i_{1}}{s-i_{1}} \binom{i_{2}}{s-i_{2}} \langle f_{n} \circ_{i_{1}+1}^{s-i_{1}+1} f_{n}, f_{n} \circ_{i_{2}+1}^{s-i_{2}+1} f_{n} \rangle \ell^{2}_{s}(\Delta_{2n-2-s}) \right)^{1/2} \end{split}$$

The claim follows from Relations (5.17) and (5.4).

Step 2: Proof of  $\tilde{C}_2 \leq C_2$ . By the Cauchy-Schwarz inequality we have

$$\mathbf{E}[|J_{n-1}(f_n(*,k))|^3] \le (\mathbf{E}[|J_{n-1}(f_n(*,k))|^2])^{1/2} (\mathbf{E}[|J_{n-1}(f_n(*,k))|^4])^{1/2}.$$

By the isometry for discrete multiple stochastic integrals we have

$$\|J_{n-1}(f_n(*,k))\|_{L^2(\Omega)} = \sqrt{(n-1)!} \|f_n(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_{n-1})}.$$

By the above relations and (5.18) we have  $\tilde{C}_2 \leq \tilde{B}_2/n$ , where  $\tilde{B}_2$  is defined by (5.19). The claim follows from (5.17) and (5.4).

Since it is not obvious that the above theorem extends Theorem 4.1 in Nourdin et al. (2010), it is worthwhile to explain this point in detail. Take  $f_n \in \ell_{\mathfrak{s}}^2(\Delta_n)$ ,  $n \geq 2$ , and let  $\tilde{h}_{n-1,n-1,s}^{(k)}$  be defined by (5.11). In the symmetric case  $p_k = q_k = 1/2$ , by the same arguments as those one after the statement of Proposition 5.1 we have that, for any fixed  $0 \leq s \leq 2(n-1)$  and  $k \in \mathbb{N}$ ,  $\tilde{h}_{n-1,n-1,s}^{(k)} = 0$  if s/2 is not an integer and

$$\tilde{h}_{n-1,n-1,s}^{(k)} = (s/2)! \binom{n-1}{s/2}^2 f_n(*,k) \circ_{s/2}^{s/2} f_n(*,k)$$

otherwise. If s/2 is an integer we also have

$$\begin{split} \|f_n(*,k) \circ_{s/2}^{s/2} f_n(*,k)\|_{\ell^2_s(\Delta_{2n-2-s})} &\leq \|f_n(*,k) \otimes_{s/2}^{s/2} f_n(*,k)\|_{\ell^2(\Delta_{2n-2-s})} \\ &= \|f_n(*,k) \star_{s/2}^{s/2} f_n(*,k)\|_{\ell^2(\Delta_{2n-2-s})} \\ &\leq \|f_n(*,k)\|_{\ell^2_s(\Delta_{n-1})}^2 < \infty, \end{split}$$

where the latter relation follows from Lemma 2.4(1) in Nourdin et al. (2010). So

$$\tilde{h}_{n-1,n-1,s}^{(k)} \in \ell_{\mathfrak{s}}^2(\Delta_{2n-2-s}).$$

In the symmetric case, by the definition of the contraction, for  $0 \le s \le 2n-3$  and  $s \le 2i \le 2(s \land (n-1))$ , we have

$$||f_n \otimes_{i+1}^{s-i+1} f_n||_{\ell^2(\Delta_{2n-2-s})} = 0, \text{ if } s < 2i$$

and

$$\begin{aligned} \|f_n \otimes_{i+1}^{s-i+1} f_n\|_{\ell^2(\Delta_{2n-2-s})} &= \|f_n \star_{i+1}^{i+1} f_n\|_{\ell^2(\Delta_{2n-2-2i})} \\ &\leq \|f_n\|_{\ell^2_s(\Delta_n)}^2, \quad \text{if } s = 2i \end{aligned}$$

where the latter relation follows from Lemma 2.4(1) in Nourdin et al. (2010). Consequently, the constant  $C_1$  in the statement of Theorem 5.5 is finite and reduces to

$$C_{1} = \left( |1 - n!| \|f_{n}\|_{\ell_{s}^{2}(\Delta_{n})}^{2} |^{2} + n^{2} \sum_{s=0}^{2(n-2)} \mathbb{1}\{s/2 \in \mathbb{N}\}(2n-2-s)! \left(\frac{s}{2}!\right)^{2} {\binom{n-1}{s/2}}^{4} \\ \times \|f_{n} \star_{s/2+1}^{s/2+1} f_{n}\|_{\ell^{2}(\Delta_{2n-2-s})}^{2} \right)^{1/2},$$
$$= \left( |1 - n!| \|f_{n}\|_{\ell_{s}^{2}(\Delta_{n})}^{2} |^{2} + n^{2} \sum_{s=1}^{n-1} (2n-2s)! \left[ (s-1)! {\binom{n-1}{s-1}}^{2} \right]^{2} \\ \times \|f_{n} \star_{s}^{s} f_{n}\|_{\ell^{2}(\Delta_{2n-2s})}^{2} \right)^{1/2}.$$

As far as the constant  $C_2$  in the statement of Theorem 5.5 is concerned, in the symmetric case one clearly has

$$C_2 = 0.$$

Finally, consider the constant  $C_3$  in the statement of Theorem 5.5. The following bound holds:

$$\begin{split} C_{3} &\leq \frac{20}{3}n^{3}\sum_{s=0}^{2n-2}(2n-2-s)! \\ &\times \sum_{s \leq \{2i_{1}, 2i_{2}\} \leq 2(s \wedge (n-1))} i_{1}!i_{2}!\binom{n-1}{i_{1}}^{2}\binom{n-1}{i_{2}}^{2}\binom{i_{1}}{s-i_{1}}\binom{i_{2}}{s-i_{2}} \\ &\times \sum_{k \geq 0} \|f_{n}(*,k) \otimes_{i_{1}}^{s-i_{1}} f_{n}(*,k)\|_{\ell^{2}(\mathbb{N})^{\otimes 2n-2-s}} \\ &\times \|f_{n}(*,k) \otimes_{i_{2}}^{s-i_{2}} f_{n}(*,k)\|_{\ell^{2}(\mathbb{N})^{\otimes 2n-2-s}} \\ &= \frac{20n^{3}}{3}\sum_{s=0}^{2n-2} \mathbbm{1}\{s/2 \in \mathbb{N}\}(2n-2-s)! \left(\frac{s}{2}!\right)^{2}\binom{n-1}{s/2}^{4} \\ &\times \sum_{k \geq 0} \|f_{n}(*,k) \star_{s/2}^{s/2} f_{n}(*,k)\|_{\ell^{2}(\mathbb{N})^{\otimes 2n-2-s}} \end{split}$$

$$= \frac{20n^3}{3} \sum_{s=1}^n (2n-2s)! \left[ (s-1)! \binom{n-1}{s-1}^2 \right]^2 \\ \times \sum_{k\geq 0} \|f_n(*,k) \star_{s-1}^{s-1} f_n(*,k)\|_{\ell^2(\mathbb{N})^{\otimes 2n-2s}}^2 \\ = \frac{20n^3}{3} \sum_{s=1}^n (2n-2s)! \left[ (s-1)! \binom{n-1}{s-1}^2 \right]^2 \|f_n \star_s^{s-1} f_n\|_{\ell^2(\mathbb{N})^{\otimes 2n-2s+1}}^2$$

where the latter equality follows from Lemma 2.4(2) (relation (2.4)) in Nourdin et al. (2010) and the constant  $C_3$  is finite again by Lemma 2.4(1) in Nourdin et al. (2010). We recovered the bound provided by Theorem 4.1 in Nourdin et al. (2010).

5.3. Convergence to the normal distribution. The next theorems follow by Theorems 5.3 and 5.4, respectively.

**Theorem 5.6.** Let  $n \geq 2$  be fixed and let  $F_m = J_n(f_m)$ ,  $m \geq 1$ , be a sequence of discrete multiple stochastic integrals such that  $f_m \in \ell^2_{\mathfrak{s}}(\Delta_n)$ , for any  $k \in \mathbb{N}$  the functions  $h_{n-1,n-1,s}^{(k)}$  and  $\tilde{h}_{n-1,n-1,s}^{(k)}$  defined by (5.10) and (5.11) with  $f_m$  in place of  $f_n$  belong to  $\ell^2_{\mathfrak{s}}(\Delta_{2n-2-s})$ ,  $0 \leq s \leq 2n-2$ ,

$$n! \|f_m\|^2_{\ell_s(\Delta_n)} \to 1, \quad as \ m \to \infty \tag{5.27}$$

$$\sum_{k\geq 0} \|f_m(*,k) \otimes_i^{s-i} f_{m]k}(*)\|_{\ell^2(\Delta_{2n-2-s})} \to 0,$$
  
as  $m \to \infty$ , for any  $0 \le s \le 2n-3$  and  $s \le 2i \le 2(s \land (n-1))$  (5.28)

$$\begin{split} \sum_{k\geq 0} &\frac{|1-2p_k|}{\sqrt{p_k q_k}} \|f_m(*,k)\|_{\ell^2_s(\Delta_{n-1})} \\ &\sqrt{\|f_m(*,k)\otimes_{i_1}^{s-i_1} f_m(*,k)\|_{\ell^2(\Delta_{2n-2-s})}} \|f_m(*,k)\otimes_{i_2}^{s-i_2} f_m(*,k)\|_{\ell^2(\Delta_{2n-2-s})}} \to 0, \\ & as \ m \to \infty, \ for \ any \ 0 \le s \le 2n-2 \ and \ s \le \{2i_1,2i_2\} \le 2(s \land (n-1)) \end{split}$$
(5.29)

and  

$$\sum_{k\geq 0} \frac{1}{p_k q_k} \|f_m(*,k) \otimes_{i_1}^{s-i_1} f_m(*,k)\|_{\ell^2(\Delta_{2n-2-s})} \|f_m(*,k) \otimes_{i_2}^{s-i_2} f_m(*,k)\|_{\ell^2(\Delta_{2n-2-s})} \to 0,$$
as  $m \to \infty$ , for any  $0 \le s \le 2n-2$  and  $s \le \{2i_1, 2i_2\} \le 2(s \land (n-1)).$ 
(5.30)

Then

$$F_m \stackrel{Law}{\to} \mathcal{N}(0,1).$$

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**Theorem 5.7.** Let  $n \ge 2$  be fixed and let  $F_m = J_n(f_m)$ ,  $m \ge 1$ , be a sequence of discrete multiple stochastic integrals such that  $f_m \in \ell^2_{\mathfrak{s}}(\Delta_n)$ , for any  $k \in \mathbb{N}$  the functions  $\tilde{h}_{n-1,n-1,s}^{(k)}$  defined by (5.11) with  $f_m$  in place of  $f_n$  belong to  $\ell^2_{\mathfrak{s}}(\Delta_{2n-2-s})$ ,  $0 \le s \le 2n-2$ ,

$$\|f_m \otimes_{i+1}^{s-i+1} f_m\|_{\ell^2(\Delta_{2n-2-s})} \to 0,$$
  
as  $m \to \infty$ , for any  $0 \le s \le 2n-3$  and  $s \le 2i \le 2(s \land (n-1))$  (5.31)

and (5.27), (5.29) and (5.30) hold. Then

$$F_m \stackrel{Law}{\to} \mathcal{N}(0,1).$$

Connection with Proposition 4.3 in Nourdin et al. (2010). In this paragraph we explain the connection between Theorem 5.7, specialized in the symmetric case, and Proposition 4.3 in Nourdin et al. (2010). Take  $f_m \in \ell_s^2(\Delta_n), m \ge 1, n \ge 2$ , and let  $\tilde{h}_{n-1,n-1,s}^{(k)}$  be defined by (5.11) with  $f_m$  in place of  $f_n$ . We already checked (after the proof of Theorem 5.5) that, in the symmetric case, one has  $\tilde{h}_{n-1,n-1,s}^{(k)} \in \ell_s^2(\Delta_{2n-2-s}), k \in \mathbb{N}, 0 \le s \le 2n-2$ . Note that assumption (5.27) is explicitly required in Proposition 4.3 of Nourdin et al. (2010) and, for  $p_k = q_k = 1/2$ , assumption (5.29) is automatically satisfied. In the symmetric case, conditions (5.30) and (5.31) read

$$\sum_{k\geq 0} \|f_m(*,k)\star_i^i f_m(*,k)\|_{\ell^2(\Delta_{2n-2-2i})}^2 \to 0, \quad \text{as } m \to \infty, \text{ for any } 0 \le i \le n-1$$
(5.32)

and

$$\|f_m \star_{i+1}^{i+1} f_m\|_{\ell^2(\Delta_{2n-2-2i})} \to 0, \quad \text{as } m \to \infty, \text{ for any } 0 \le i \le n-2$$
 (5.33)

respectively. We are going to check that assumption (4.44) of Proposition 4.3 in Nourdin et al. (2010), i.e.

$$\|f_m \star_r^r f_m\|_{\ell^2(\mathbb{N})^{\otimes 2n-2r}} \to 0, \quad \text{as } m \to \infty, \text{ for any } 1 \le r \le n-1$$
(5.34)

implies (5.32) and (5.33). Clearly, (5.34) is equivalent to (5.33). Moreover, for any  $0 \le i \le n-1$ , by Lemma 2.4(2) (relation (2.4)) and Lemma 2.4(3) in Nourdin et al. (2010), we have

$$\sum_{k\geq 0} \|f_m(*,k)\star_i^i f_m(*,k)\|_{\ell^2(\Delta_{2n-2-2i})}^2 \leq \sum_{k\geq 0} \|f_m(*,k)\star_i^i f_m(*,k)\|_{\ell^2(\mathbb{N})^{\otimes 2n-2-2i}}^2$$
$$\leq \|f_m\star_n^{n-1} f_m\|_{\ell^2(\mathbb{N})} \|f_m\|_{\ell^2_{\mathfrak{s}}(\Delta_n)}^2$$
$$\leq \|f_m\star_{n-1}^{n-1} f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}} \|f_m\|_{\ell^2_{\mathfrak{s}}(\Delta_n)}^2.$$

So combining (5.27) and condition (5.34) with r = n - 1 we have (5.32).

Quadratic functionals. In the next proposition we apply Theorem 5.7 with n = 2. In comparison with Proposition 4.3 of Nourdin et al. (2010) we require an additional  $\ell^4$  condition in the non-symmetric case.

**Proposition 5.8.** Assume that there exists some  $\varepsilon > 0$  such that

$$0 < \varepsilon < p_k < 1 - \varepsilon, \qquad k \in \mathbb{N},\tag{5.35}$$

and consider a sequence  $f_m \in \ell^2_{\mathfrak{s}}(\Delta_2)$  such that

 $\begin{array}{l} a) \ \lim_{m \to \infty} \|f_m\|_{\ell^2_s(\Delta_2)}^2 = 1/2, \\ b) \ \lim_{m \to \infty} \|f_m \star^1_1 f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}} = 0, \\ c) \ \lim_{m \to \infty} \|f_m\|_{\ell^4_s(\Delta_2)} \to 0. \\ Then \ J_2(f_m) \xrightarrow{Law} \mathcal{N}(0,1) \ as \ m \to \infty. \end{array}$ 

*Proof.* We need to satisfy Conditions (5.27), (5.29), (5.30) and (5.31), in addition to an integrability check which is postponed to the end of this proof. First we note that (5.27) is Condition a) above. Next we note that under (5.35), Conditions (5.29), (5.30) and (5.31) read

$$\begin{split} &\sum_{k\geq 0} \|f_m(*,k)\|_{\ell^2_s(\Delta_1)} \|f_m(*,k) \otimes^0_0 f_m(*,k)\|_{\ell^2_s(\Delta_2)} \to 0, \\ &\sum_{k\geq 0} \|f_m(*,k)\|_{\ell^2_s(\Delta_1)} \|f_m(*,k) \otimes^0_1 f_m(*,k)\|_{\ell^2_s(\Delta_1)} \to 0, \\ &\sum_{k\geq 0} \|f_m(*,k)\|_{\ell^2_s(\Delta_1)} \|f_m(*,k) \otimes^1_1 f_m(*,k)\|_{\ell^2_s(\Delta_0)} \to 0, \\ &\sum_{k\geq 0} \|f_m(*,k) \otimes^0_0 f_m(*,k)\|_{\ell^2_s(\Delta_2)}^2 \to 0, \\ &\sum_{k\geq 0} \|f_m(*,k) \otimes^0_1 f_m(*,k)\|_{\ell^2_s(\Delta_1)}^2 \to 0, \\ &\sum_{k\geq 0} \|f_m(*,k) \otimes^1_1 f_m(*,k)\|_{\ell^2_s(\Delta_0)}^2 \to 0, \end{split}$$

and

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 $||f_m \otimes_1^1 f_m||_{\ell^2(\Delta_2)}, ||f_m \otimes_2^1 f_m||_{\ell^2(\Delta_1)} \to 0,$ 

as  $m \to \infty$ . Using again (5.35), we have that the above conditions are implied by

$$\sum_{k\geq 0} \|f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_1)} \|f_m(*,k) \star^0_0 f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_2)} \to 0,$$
(5.36)

$$\sum_{k\geq 0} \|f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_1)} \|f_m(*,k) \star^0_1 f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_1)} \to 0,$$
(5.37)

$$\sum_{k\geq 0} \|f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_1)} \|f_m(*,k) \star^1_1 f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_0)} \to 0,$$
(5.38)

$$\sum_{k\geq 0} \|f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_1)} \|f_m(*,k)|_{\ell^2_{\mathfrak{s}}(\Delta_2)} \to 0,$$

$$\sum_{k\geq 0} \|f_m(*,k)\star^0_0 f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_2)}^2 \to 0,$$
(5.39)

$$\sum_{k>0} \|f_m(*,k) \star_1^0 f_m(*,k)\|_{\ell_s^2(\Delta_1)}^2 \to 0,$$
(5.40)

$$\sum_{k\geq 0} \|f_m(*,k)\star_1^1 f_m(*,k)\|_{\ell^2_s(\Delta_0)}^2 \to 0,$$
(5.41)

and

$$\|f_m \star_1^1 f_m\|_{\ell^2(\Delta_2)}, \, \|f_m \star_2^1 f_m\|_{\ell^2(\Delta_1)} \to 0, \tag{5.42}$$

as  $m \to \infty$ . Now by Lemma 2.4(3) of Nourdin et al. (2010) we have

$$\|f_m \star_2^1 f_m\|_{\ell^2(\Delta_1)} \le \|f_m \star_1^1 f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}},$$

hence (5.42) is implied by Condition b) above. Next, by Lemma 2.4(2)-(3) in Nourdin et al. (2010) we have

$$\sum_{k\geq 0} \|f_m(*,k)\star_0^0 f_m(*,k)\|_{\ell^2_s(\Delta_2)}^2 \leq \|f_m\star_2^1 f_m\|_{\ell^2(\mathbb{N})} \|f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}}^2$$
$$\leq \|f_m\star_1^1 f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}} \|f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}}^2, \qquad (5.43)$$

hence Condition (5.39) is implied by Conditions a) and b) above. Similarly, by Lemma 2.4(2)-(3) in Nourdin et al. (2010) we have

$$\sum_{k\geq 0} \|f_m(*,k)\star_1^1 f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_0)}^2 \leq \|f_m\star_2^1 f_m\|_{\ell^2(\mathbb{N})} \|f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}}^2$$
$$\leq \|f_m\star_1^1 f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}} \|f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}}^2, \qquad (5.44)$$

hence Condition (5.41) is also implied by Conditions a) and b) above. Now, we have

$$\sum_{k\geq 0} \|f_m(*,k)\|_{\ell^2_s(\Delta_1)} \|f_m(*,k) \star^0_0 f_m(*,k)\|_{\ell^2_s(\Delta_2)}$$

$$\leq \left(\sum_{k\geq 0} \|f_m(*,k)\|_{\ell^2_s(\Delta_1)}^2\right)^{1/2} \left(\sum_{k\geq 0} \|f_m(*,k) \star^0_0 f_m(*,k)\|_{\ell^2_s(\Delta_2)}^2\right)^{1/2} \tag{5.45}$$

$$\leq \left(\sum_{k\geq 0} \|f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_1)}^2\right)^{1/2} \|f_m \star^1_1 f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}}^{1/2} \|f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}}^{1/2}$$
(5.46)

$$= \left(\sum_{k\geq 0} f_m(*,k) \star_1^1 f_m(*,k)\right)^{1/2} \|f_m \star_1^1 f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}}^{1/2} \|f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}}$$
(5.47)

$$= \left(f_m \star_2^2 f_m\right)^{1/2} \|f_m \star_1^1 f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}}^{1/2} \|f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}}$$
(5.48)

$$= \|f_m\|_{\ell_s^2(\Delta_2)}^2 \|f_m \star_1^1 f_m\|_{\ell^2(\mathbb{N})^{\otimes 2}}^{1/2}, \tag{5.49}$$

where in (5.45) we used the Cauchy-Schwartz inequality, in (5.46) we used (5.43), in (5.47) we used the identity

$$f \star_n^n g = \langle f, g \rangle_{\ell^2(\mathbb{N})^{\otimes n}}, \quad f, g \in \ell^2(\mathbb{N})^{\otimes n}$$
(5.50)

in (5.48) we used the equality

$$\sum_{k\geq 0} f_m(*,k) \star_1^1 f_m(*,k) = f_m \star_2^2 f_m,$$

and in (5.49) we used (5.50). Similarly, using (5.44) we have

$$\sum_{k\geq 0} \|f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_1)} \|f_m(*,k)\star^1_1 f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_0)} \leq \|f_m\|^2_{\ell^2_{\mathfrak{s}}(\Delta_2)} \|f_m\star^1_1 f_m\|^{1/2}_{\ell^2(\mathbb{N})^{\otimes 2}}.$$
(5.51)

So Conditions (5.36) and (5.38) are also implied by Conditions a) and b) above. By similar arguments as above, we have

$$\sum_{k\geq 0} \|f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_1)} \|f_m(*,k) \star^0_1 f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_1)}$$
  
$$\leq \|f_m\|_{\ell^2_{\mathfrak{s}}(\Delta_2)} \left( \sum_{k\geq 0} \|f_m(*,k) \star^0_1 f_m(*,k)\|_{\ell^2_{\mathfrak{s}}(\Delta_1)}^2 \right)^{1/2}.$$
(5.52)

We note that

$$\sum_{k\geq 0} \|f_m(*,k)\star_1^0 f_m(*,k)\|_{\ell_s^2(\Delta_1)}^2 \le \sum_{k\geq 0} \sum_{a\geq 0} f_m^4(a,k) = \|f_m^2\|_{\ell_s^2(\Delta_2)}^2,$$
(5.53)

and so Condition c) above implies (5.37) and (5.40). Finally we note that for any  $k \in \mathbb{N}$  the functions  $\tilde{h}_{1,1,s}^{(k)}$ ,  $0 \leq s \leq 2$ , defined by (5.11) with  $f_m$  in place of  $f_n$ , i.e.

$$\tilde{h}_{1,1,0}^{(k)} = f_m(*,k) \circ_0^0 f_m(*,k)$$
$$\tilde{h}_{1,1,1}^{(k)} = f_m(*,k) \circ_1^0 f_m(*,k)$$

and

$$\tilde{h}_{1,1,2}^{(k)} = f_m(*,k) \circ_1^1 f_m(*,k)$$

belong to  $\ell_{\mathfrak{s}}^2(\Delta_2)$ ,  $\ell_{\mathfrak{s}}^2(\Delta_1)$  and  $\ell_{\mathfrak{s}}^2(\Delta_0)$ , respectively. Indeed, this easily follows from (5.4), (5.35) and Lemma 2.4(1) of Nourdin et al. (2010).

#### Example

A straightforward computation shows that examples of function sequences that satisfy the hypotheses of Proposition 5.8 include

$$f_m(k_1, k_2) = \frac{1}{m\sqrt{2}} \mathbb{1}_{[0,m]^2}(k_1, k_2), \qquad m \ge 1.$$

Note that the above example will also satisfy the hypotheses of Theorem 5.6 as well. More generally, any sequence of non-negative kernels satisfying the hypotheses of Proposition 5.8 will satisfy the hypotheses of Theorem 5.6. Indeed, under Condition (5.35), for non-negative kernels, Condition (5.28) is implied by (5.39), (5.40) and (5.41) which, as showed in the proof of Proposition 5.8, are in turn implied by Conditions a), b) and c). However, elementary computations have shown that, in general, it is difficult to compare the second addends of the constants  $B_1$  and  $C_1$ of Theorems 5.3 and 5.4, respectively. Consequently, in general, it is difficult to compare Conditions (5.28) and (5.31) of Theorems 5.6 and 5.7, respectively.

## 6. Poisson approximation of Bernoulli functionals

We recall that the total variation distance between the laws of two N-valued random variables  $Y_i$ , i = 1, 2, is given by

$$d_{TV}(Y_1, Y_2) := \sup_{A \subseteq \mathbb{N}} |P(Y_1 \in A) - P(Y_2 \in A)|.$$

Of course, the topology induced by  $d_{TV}$  on the class of all probability laws on  $\mathbb{N}$  is strictly stronger than the topology induced by the convergence in distribution.

In this section we present two different upper bounds for  $d_{TV}(F, \text{Po}(\lambda))$ , where  $\text{Po}(\lambda)$  is a Poisson random variable with mean  $\lambda > 0$ . The first one is obtained by using the covariance representation formula (2.15), while the second one is obtained by using the covariance representation formula (2.16).

Before proceeding further we recall some necessary background on the Chen-Stein method for the Poisson approximation and refer to Barbour et al. (1992) for more insight into this technique.

Chen-Stein's method for Poisson approximation. Given  $A \subseteq \mathbb{N}$ , it turns out that there exists a unique function  $f_A : \mathbb{N} \to \mathbb{R}$  such that

$$\mathbb{1}_A(k) - P(\operatorname{Po}(\lambda) \in A) = \lambda f_A(k+1) - k f_A(k), \quad k \in \mathbb{N}$$
(6.1)

verifying the boundary condition  $\Delta^2 f(0) = 0$ . The above equation is called Chen-Stein's equation. Combining e.g. Theorem 2.3 in Erhardsson (2005) and Theorem 1.3 in Daly (2008), we deduce that the function  $f_A$  has the following properties:

$$\|f_A\|_{\infty} \le \min\left(1, \sqrt{\frac{2}{\lambda e}}\right), \quad \|\Delta f_A\|_{\infty} \le \frac{1 - e^{-\lambda}}{\lambda}, \quad \|\Delta^2 f_A\|_{\infty} \le \frac{2 - 2e^{-\lambda}}{\lambda^2}.$$
(6.2)

6.1. Clark-Ocone bound.

**Theorem 6.1.** Let  $F \in Dom(D)$  be an N-valued random variable with mean  $\lambda$  and assume that

$$b_1 := \mathbf{E}[|\langle \mathbf{E}[D.F \mid \mathcal{F}_{\cdot-1}], D.F \rangle_{\ell^2(\mathbb{N})} - \lambda|],$$

and

$$\begin{split} b_2 &:= & \mathbf{E}\left[\left\langle |\mathbf{E}[D.F \,|\, \mathcal{F}_{.-1}]|, \left|\sqrt{\frac{q.}{p.}}D.F\left(\frac{D.F}{\sqrt{p.q.}}-1\right)\right|\right\rangle_{\ell^2(\mathbb{N})}\right] \\ &+ \mathbf{E}\left[\left\langle |\mathbf{E}[D.F \,|\, \mathcal{F}_{.-1}]|, \left|\sqrt{\frac{p.}{q.}}D.F\left(\frac{D.F}{\sqrt{p.q.}}+1\right)\right|\right\rangle_{\ell^2(\mathbb{N})}\right] \end{split}$$

are finite. Then we have

$$d_{TV}(F, \operatorname{Po}(\lambda)) \le \frac{1 - e^{-\lambda}}{\lambda} b_1 + \frac{1 - e^{-\lambda}}{\lambda^2} b_2.$$
(6.3)

*Proof.* We start by checking the domain condition  $f_A(F) \in \text{Dom}(D)$ . Assume  $F_k^+ \ge F_k^-$ . We note that

$$D_k f_A(F) = \sqrt{p_k q_k} (f_A(F_k^+) - f_A(F_k^-))$$
  
=  $\sqrt{p_k q_k} \sum_{h=1}^{F_k^+ - F_k^-} (f_A(F_k^+ - h + 1) - f_A(F_k^+ - h)),$ 

and so  $|D_k f_A(F)| \leq ||\Delta f_A||_{\infty} \sqrt{p_k q_k} (F_k^+ - F_k^-)$ . Similarly, if  $F_k^+ < F_k^-$  then  $|D_k f_A(F)| \leq ||\Delta f_A||_{\infty} \sqrt{p_k q_k} (F_k^- - F_k^+)$ . Consequently,

$$E[\|Df_A(F)\|_{\ell^2(\mathbb{N})}^2] = E\left[\sum_{k\geq 0} |D_k f_A(F)|^2\right] \leq \|\Delta f_A\|_{\infty}^2 E\left[\sum_{k\geq 0} |D_k F|^2\right] \\ = E[\|DF\|_{\ell^2(\mathbb{N})}^2],$$

and this latter quantity is finite since  $F \in \text{Dom}(D)$ . The claimed domain condition follows. By the Chen-Stein equation (6.1), the covariance representation (2.15) and Proposition 2.2, we have

$$\begin{aligned} P(\operatorname{Po}(\lambda) \in A) &- P(F \in A) = \operatorname{E}[(F - \lambda)f_A(F) - \lambda(f_A(F + 1) - f_A(F))] \\ &= \operatorname{E}\left[\sum_{k \ge 0} \operatorname{E}[D_k F \mid \mathcal{F}_{k-1}] D_k f_A(F) - \lambda \Delta f_A(F)\right] \\ &= \operatorname{E}\left[\sum_{k \ge 0} \operatorname{E}[D_k F \mid \mathcal{F}_{k-1}] (\Delta f_A(F) D_k F + R_k^F(f_A)) - \lambda \Delta f_A(F)\right] \\ &= \operatorname{E}\left[\Delta f_A(F) (\langle \operatorname{E}[D.F \mid \mathcal{F}_{\cdot-1}], D.F \rangle_{\ell^2(\mathbb{N})} - \lambda)] + \operatorname{E}[\langle \operatorname{E}[D.F \mid \mathcal{F}_{\cdot-1}], R_{\cdot}^F(f_A) \rangle_{\ell^2(\mathbb{N})}] \end{aligned}$$

The desired result follows by taking absolute values on both sides, as well as by applying the estimates (6.2) and (2.13), and noticing that the random variables  $D_k F$  and  $E[D_k F | \mathcal{F}_{k-1}]$  are independent of  $X_k$ .

**Corollary 6.2.** Let  $F \in Dom(D)$  be an N-valued random variable with mean  $\lambda$  and assume that

$$b_1 := |\lambda - \operatorname{Var}(F)| + ||\langle D.F, \operatorname{E}[D.F | \mathcal{F}_{-1}]\rangle_{\ell^2(\mathbb{N})} - \operatorname{E}[\langle D.F, \operatorname{E}[D.F | \mathcal{F}_{-1}]\rangle_{\ell^2(\mathbb{N})}]||_{L^2(\Omega)},$$

and

$$b_2 := \sum_{k \ge 0} \frac{1}{\sqrt{p_k q_k}} \mathbb{E}[|D_k F|^2] + \sum_{k \ge 0} \frac{1}{p_k q_k} ||D_k F||_{L^2(\Omega)} \sqrt{\mathbb{E}[|D_k F|^4]}$$
(6.4)

are finite. Then we have

$$d_{TV}(F, \operatorname{Po}(\lambda)) \le \frac{1 - \mathrm{e}^{-\lambda}}{\lambda} b_1 + \frac{1 - \mathrm{e}^{-\lambda}}{\lambda^2} b_2.$$

*Proof.* We preliminary note that by the Clark-Ocone representation formula (2.15) one has

$$\operatorname{Var}(F) = \operatorname{E}[\langle D.F, \operatorname{E}[D.F | \mathcal{F}_{\cdot-1}] \rangle_{\ell^{2}(\mathbb{N})}].$$
(6.5)

By the Cauchy-Schwarz inequality and (6.5), we have  $E[|\lambda - \langle E[D.F | \mathcal{F}_{-1}], DF \rangle_{\ell^{2}(\mathbb{N})}|]$   $\leq \|\lambda - \langle E[D.F | \mathcal{F}_{-1}], DF \rangle_{\ell^{2}(\mathbb{N})}\|_{L^{2}(\Omega)}$ 

 $\leq |\lambda - \operatorname{Var}(F)| + \|\langle D.F, \operatorname{E}[D.F \mid \mathcal{F}_{\cdot-1}] \rangle_{\ell^{2}(\mathbb{N})} - \operatorname{E}[\langle D.F, \operatorname{E}[D.F \mid \mathcal{F}_{\cdot-1}] \rangle_{\ell^{2}(\mathbb{N})}]\|_{L^{2}(\Omega)}.$ Moreover,

$$\begin{split} & \operatorname{E}\left[\left\langle\left|\operatorname{E}[D.F \mid \mathcal{F}_{\cdot-1}]\right|, \left|\sqrt{\frac{q_{\cdot}}{p_{\cdot}}}D.F\left(\frac{D.F}{\sqrt{p\cdot q_{\cdot}}}-1\right)\right|\right\rangle_{\ell^{2}(\mathbb{N})}\right] \\ & +\operatorname{E}\left[\left\langle\left|\operatorname{E}[D.F \mid \mathcal{F}_{\cdot-1}]\right|, \left|\sqrt{\frac{p_{\cdot}}{q_{\cdot}}}D.F\left(\frac{D.F}{\sqrt{p\cdot q_{\cdot}}}+1\right)\right|\right\rangle_{\ell^{2}(\mathbb{N})}\right] \\ & \leq \operatorname{E}\left[\left\langle\left|\operatorname{E}[D.F \mid \mathcal{F}_{\cdot-1}]\right|, \frac{|D.F|}{\sqrt{p\cdot q_{\cdot}}}\right\rangle_{\ell^{2}(\mathbb{N})}\right] + \operatorname{E}\left[\left\langle\left|\operatorname{E}[D.F \mid \mathcal{F}_{\cdot-1}]\right|, \frac{|D.F|^{2}}{p\cdot q_{\cdot}}\right\rangle_{\ell^{2}(\mathbb{N})}\right] \\ & \leq \operatorname{E}\left[\left\langle\operatorname{E}[|D.F \mid \mathcal{F}_{\cdot-1}], \frac{|D.F|}{\sqrt{p\cdot q_{\cdot}}}\right\rangle_{\ell^{2}(\mathbb{N})}\right] + \operatorname{E}\left[\left\langle\left|\operatorname{E}[D.F \mid \mathcal{F}_{\cdot-1}]\right|, \frac{|D.F|^{2}}{p\cdot q_{\cdot}}\right\rangle_{\ell^{2}(\mathbb{N})}\right] \\ & \leq \operatorname{E}\left[\left\langle\operatorname{E}[|D.F \mid |\mathcal{F}_{\cdot-1}], \frac{|D.F|}{\sqrt{p\cdot q_{\cdot}}}\right\rangle_{\ell^{2}(\mathbb{N})}\right] + \operatorname{E}\left[\left\langle\left|\operatorname{E}[D.F \mid \mathcal{F}_{\cdot-1}]\right|, \frac{|D.F|^{2}}{p\cdot q_{\cdot}}\right\rangle_{\ell^{2}(\mathbb{N})}\right] \\ & \leq \sum_{k\geq0} \frac{1}{\sqrt{p_{k}q_{k}}} \operatorname{E}[|D_{k}F|^{2}] + \sum_{k\geq0} \frac{1}{p_{k}q_{k}} \operatorname{E}[|E[D_{k}F \mid \mathcal{F}_{k-1}]||D_{k}F|^{2}] \\ & \leq \sum_{k\geq0} \frac{1}{\sqrt{p_{k}q_{k}}} \operatorname{E}[|D_{k}F|^{2}] + \sum_{k\geq0} \frac{1}{p_{k}q_{k}} ||D_{k}F ||\mathcal{F}_{k-1}]||L^{2}(\Omega)\sqrt{\operatorname{E}[|D_{k}F|^{4}]}. \end{split}$$

The claim follows from Theorem 6.1.

6.2. Semigroup bound. The next result is formally similar to Theorem 3.1 of Peccati (2011). More precisely, the first addend in the right-hand side of (6.6) coincides with the term in the right-hand side of relation (3.5) in Peccati (2011) when replacing the finite difference operator on the Bernoulli space with the finite difference operator on the Poisson space. As for the second addend in the right-hand side of (6.6), although it has some similarities with the corresponding term in (3.6) of Peccati (2011) (the expectations have the same multiplicative constant), the two terms remain different when replacing the finite difference operator on the Bernoulli space with the finite difference operator on the Bernoulli space with the finite difference operator on the Bernoulli space with the finite difference operator on the Bernoulli space with the finite difference operator on the Bernoulli space with the finite difference operator on the Bernoulli space with the finite difference operator on the Bernoulli space with the finite difference operator on the Bernoulli space with the finite difference operator on the Bernoulli space with the finite difference operator on the Bernoulli space with the finite difference operator on the Bernoulli space with the finite difference operator on the Poisson space.

**Theorem 6.3.** Let  $F \in Dom(D)$  be an N-valued random variable with mean  $\lambda$  and assume that

$$c_1 := \mathbb{E}[|\lambda - \langle DF, -DL^{-1}F \rangle_{\ell^2(\mathbb{N})}|]$$

and

$$c_{2} := E\left[\left\langle \left| \sqrt{\frac{q_{\cdot}}{p_{\cdot}}} D.F\left(\frac{D.F}{\sqrt{p_{\cdot}q_{\cdot}}} - 1\right) \right|, |DL^{-1}F| \right\rangle_{\ell^{2}(\mathbb{N})} \right] \\ + E\left[\left\langle \left| \sqrt{\frac{p_{\cdot}}{q_{\cdot}}} D.F\left(\frac{D.F}{\sqrt{p_{\cdot}q_{\cdot}}} + 1\right) \right|, |DL^{-1}F| \right\rangle_{\ell^{2}(\mathbb{N})} \right]$$

are finite. Then we have

$$d_{TV}(F, \operatorname{Po}(\lambda)) \le \frac{1 - e^{-\lambda}}{\lambda} c_1 + \frac{1 - e^{-\lambda}}{\lambda^2} c_2, \qquad (6.6)$$

*Proof.* Although the proof is similar to that one of Theorem 6.1, we give the details since some points need a different justification. As in the proof of Theorem 6.1 one has  $f_A(F) \in \text{Dom}(D)$ . By the Chen-Stein equation (6.1), the covariance representation (2.16) and Proposition 2.2, we have

$$P(\operatorname{Po}(\lambda) \in A) - P(F \in A) = \operatorname{E}[(F - \lambda)f_A(F) - \lambda(f_A(F + 1) - f_A(F))]$$
  

$$= \operatorname{E}\left[\langle Df_A(F), -DL^{-1}F \rangle_{\ell^2(\mathbb{N})} - \lambda \Delta f_A(F)\right]$$
  

$$= \operatorname{E}\left[\langle \Delta f_A(F)DF + R^F(f_A), -DL^{-1}F \rangle_{\ell^2(\mathbb{N})} - \lambda \Delta f_A(F)\right]$$
  

$$= \operatorname{E}\left[\Delta f_A(F)(\langle -DL^{-1}F, DF \rangle_{\ell^2(\mathbb{N})} - \lambda)\right]$$
  

$$+ \operatorname{E}[\langle -DL^{-1}F, R^F(f_A) \rangle_{\ell^2(\mathbb{N})}].$$

The desired result follows by taking absolute values on both sides, as well as by applying the estimates (6.2) and (2.13), and noticing that the random variables  $D_k F$  and  $D_k L^{-1} F$  are independent of  $X_k$  (see Lemma 2.13 (1) in Nourdin et al. (2010)).

Note that, formally, the upper bound (6.3) may be obtained by (6.6) substituting the term  $-D_kL^{-1}F$  in the definitions of  $c_1$  and  $c_2$  with  $\mathbb{E}[D_kF \mid \mathcal{F}_{k-1}]$ , and vice versa.

**Corollary 6.4.** Let  $F \in Dom(D)$  be an  $\mathbb{N}$ -valued random variable with mean  $\lambda$  and assume that

$$c_1 := |\lambda - \operatorname{Var}(F)| + ||\langle D.F, -D.L^{-1}F \rangle_{\ell^2(\mathbb{N})} - \operatorname{E}[\langle D.F, -D.L^{-1}F \rangle_{\ell^2(\mathbb{N})}]||_{L^2(\Omega)},$$
  
and  $c_2 := b_2$ , where  $b_2$  is defined by (6.4), are finite. Then we have

$$d_{TV}(F, \operatorname{Po}(\lambda)) \leq \frac{1 - \mathrm{e}^{-\lambda}}{\lambda} c_1 + \frac{1 - \mathrm{e}^{-\lambda}}{\lambda^2} c_2.$$

$$\operatorname{Var}(F) = \operatorname{E}[\langle DF, -DL^{-1}F \rangle_{\ell^{2}(\mathbb{N})}].$$
(6.7)

By the Cauchy-Schwarz inequality and (6.7), we have

$$\begin{split} & \mathbf{E}[|\lambda - \langle -DL^{-1}F, DF \rangle_{\ell^{2}(\mathbb{N})}|] \\ & \leq \|\lambda - \langle -DL^{-1}F, DF \rangle_{\ell^{2}(\mathbb{N})}\|_{L^{2}(\Omega)} \\ & \leq |\lambda - \operatorname{Var}(F)| + \|\langle D.F, -DL^{-1}F \rangle_{\ell^{2}(\mathbb{N})} - \mathbf{E}[\langle D.F, -DL^{-1}F \rangle_{\ell^{2}(\mathbb{N})}]\|_{L^{2}(\Omega)}. \end{split}$$

Moreover, using the inequality (3.16) we deduce

$$\begin{split} & \mathbf{E}\left[\langle |DL^{-1}F|, \left|\sqrt{\frac{q.}{p.}}D.F\left(\frac{D.F}{\sqrt{p.q.}}-1\right)\right|\rangle_{\ell^{2}(\mathbb{N})}\right] \\ & +\mathbf{E}\left[\langle |DL^{-1}F|, \left|\sqrt{\frac{p.}{q.}}D.F\left(\frac{D.F}{\sqrt{p.q.}}+1\right)\right|\rangle_{\ell^{2}(\mathbb{N})}\right] \\ & \leq \mathbf{E}\left[\langle |DL^{-1}F|, \frac{|D.F|}{\sqrt{p.q.}}\left(1+\frac{|D.F|}{\sqrt{p.q.}}\right)\rangle_{\ell^{2}(\mathbb{N})}\right] \\ & \leq \mathbf{E}\left[\langle |DL^{-1}F|, \frac{|D.F|}{\sqrt{p.q.}}\rangle_{\ell^{2}(\mathbb{N})}\right] + \mathbf{E}\left[\langle |DL^{-1}F|, \frac{|D.F|^{2}}{p.q.}\rangle_{\ell^{2}(\mathbb{N})}\right] \\ & = \sum_{k\geq 0} \frac{1}{\sqrt{pkq_{k}}}\mathbf{E}[|D_{k}L^{-1}F||D_{k}F|] + \sum_{k\geq 0} \frac{1}{p_{k}q_{k}}\mathbf{E}[|D_{k}L^{-1}F||D_{k}F|^{2}] \\ & \leq \sum_{k\geq 0} \frac{1}{\sqrt{pkq_{k}}}\mathbf{E}[|D_{k}F|^{2}] + \sum_{k\geq 0} \frac{1}{p_{k}q_{k}}||D_{k}F^{-1}F||_{L^{2}(\Omega)}\sqrt{\mathbf{E}[|D_{k}F|^{4}]} \\ & \leq \sum_{k\geq 0} \frac{1}{\sqrt{pkq_{k}}}\mathbf{E}[|D_{k}F|^{2}] + \sum_{k\geq 0} \frac{1}{p_{k}q_{k}}||D_{k}F||_{L^{2}(\Omega)}\sqrt{\mathbf{E}[|D_{k}F|^{4}]}. \end{split}$$

The claim follows from Theorem 6.3.

# 7. First chaos bound for the Poisson approximation

In this section we specialize the results of Section 6 to (shifted) first order discrete stochastic integrals. As we shall see, the bounds (6.3) and (6.6) (and the corresponding assumptions) coincide on functionals of the form  $F = \lambda + J_1(f_1)$ ,  $f_1 \in \ell^2(\mathbb{N})$ , although they differ for  $F = \lambda + J_n(f_n)$ ,  $f_n \in \ell^2_{\mathfrak{s}}(\Delta_n)$ ,  $n \geq 2$ .

**Corollary 7.1.** Assume that  $\alpha = (\alpha_k)_{k\geq 0}$  is in  $\ell^2(\mathbb{N})$  and such that

$$F = \lambda + J_1(\alpha) = \lambda + \sum_{k \ge 0} \alpha_k Y_k$$

is  $\mathbb{N}$ -valued. Assuming

$$\sum_{k\geq 0} \sqrt{\frac{q_k}{p_k}} |\alpha_k|^2 \left| \frac{\alpha_k}{\sqrt{p_k q_k}} - 1 \right| < \infty$$

and

$$\sum_{k\geq 0} \sqrt{\frac{p_k}{q_k}} |\alpha_k|^2 \Big| \frac{\alpha_k}{\sqrt{p_k q_k}} + 1 \Big| < \infty,$$

we have that the bound (6.3) (which in this case coincides with the bound (6.6)) holds for F with

$$b_1 = \Big|\lambda - \sum_{k \ge 0} \alpha_k^2\Big|,$$

and

$$b_{2} = \sum_{k \ge 0} \sqrt{\frac{q_{k}}{p_{k}}} |\alpha_{k}|^{2} \Big| \frac{\alpha_{k}}{\sqrt{p_{k}q_{k}}} - 1 \Big| + \sum_{k \ge 0} \sqrt{\frac{p_{k}}{q_{k}}} |\alpha_{k}|^{2} \Big| \frac{\alpha_{k}}{\sqrt{p_{k}q_{k}}} + 1 \Big|.$$

*Proof.* As in the proof of Corollary 4.1 we have  $F \in \text{Dom}(D)$  with  $D_k F = \alpha_k$ . The claim follows from e.g. Theorem 6.1.

# Example

Let  $(Z_k)_{k\geq 0}$  be a sequence of independent and  $\{0, 1\}$ -valued random variables with  $E[Z_k] = p_k$  and define the random variables

$$Y_k = \frac{Z_k - p_k}{\sqrt{p_k q_k}} = \frac{q_k - p_k + X_k}{2\sqrt{p_k q_k}},$$

where  $(X_k)_{k\geq 0}$  is a sequence of independent and  $\{-1,1\}$ -valued random variables with  $P(X_k = 1) = p_k$ . Let  $(\beta_k)_{k\geq 0} \subset \mathbb{N}$ , assume

$$\begin{split} \lambda &:= \sum_{k \ge 0} p_k \beta_k < \infty, \quad \sum_{k \ge 0} p_k q_k \beta_k^2 < \infty \\ &\sum_{k \ge 0} q_k \sqrt{p_k q_k} \beta_k^2 |\beta_k - 1| < \infty, \quad \sum_{k \ge 0} p_k \sqrt{p_k q_k} \beta_k^2 (\beta_k + 1) < \infty \end{split}$$

and define  $\alpha_k := \sqrt{p_k q_k} \beta_k$ . We clearly have  $\alpha = (\alpha_k)_{k \ge 0} \in \ell^2(\mathbb{N})$  and

$$F = \sum_{k \ge 0} \beta_k Z_k = \lambda + \sum_{k \ge 0} \alpha_k Y_k.$$

Note that, obviously, F takes values in  $\mathbb N.$  Note also that

$$\sum_{k\geq 0} \sqrt{\frac{q_k}{p_k}} |\alpha_k|^2 \left| \frac{\alpha_k}{\sqrt{p_k q_k}} - 1 \right| = \sum_{k\geq 0} q_k \sqrt{p_k q_k} \beta_k^2 |\beta_k - 1|,$$

and

$$\sum_{k\geq 0} \sqrt{\frac{p_k}{q_k}} |\alpha_k|^2 \left| \frac{\alpha_k}{\sqrt{p_k q_k}} + 1 \right| = \sum_{k\geq 0} p_k \sqrt{p_k q_k} \beta_k^2 (\beta_k + 1).$$

So (with  $\lambda$  as above) by Corollary 7.1 we have

$$d_{TV}(F, \operatorname{Po}(\lambda)) \leq \frac{1 - e^{-\lambda}}{\lambda} \Big| \sum_{k \geq 0} p_k \beta_k (1 - \beta_k q_k) \Big|$$
$$+ \frac{1 - e^{-\lambda}}{\lambda^2} \left( \sum_{k \geq 0} \sqrt{p_k q_k} \beta_k^2 (|\beta_k - 1| q_k + (\beta_k + 1) p_k) \right).$$
(7.1)

We note that by the classical bound for independent Bernoulli random variables (see e.g. Barbour et al. (1992)) we have

$$d_{TV}\left(\sum_{k\geq 0} Z_k, \operatorname{Po}(\lambda)\right) \leq \frac{1 - e^{-\lambda}}{\lambda} \sum_{k\geq 0} p_k^2.$$
(7.2)

Although the inequality (7.1) with  $\beta_k = 1$  does not coincide with (7.2) (producing indeed a bigger upper bound), the bound in (7.1) holds, more generally, for sums of "weighted" Bernoulli random variables.

# 8. nth chaos bounds for the Poisson approximation

Throughout this section, for a fixed positive constant  $\lambda > 0$ , we consider an  $\mathbb{N}$ -valued (shifted) *n*th chaos  $F = \lambda + J_n(f_n), f_n \in \ell_{\mathfrak{s}}^2(\Delta_n), n \geq 2$ .

8.1. Clark-Ocone bound. By (5.5) and the isometric properties of discrete multiple stochastic integrals we have that the constants  $b_i$  of Corollary 6.2 are equal to

$$\tilde{b}_1 := |\lambda - n! \| f_n \|_{\ell^2_s(\Delta_n)}^2 |+ n^2 \| \langle J_{n-1}(f_n(*, \cdot)), J_{n-1}(f_n] \cdot (*)) \rangle_{\ell^2(\mathbb{N})} \\ - \mathbf{E}[ \langle J_{n-1}(f_n(*, \cdot)), J_{n-1}(f_n] \cdot (*)) \rangle_{\ell^2(\mathbb{N})}] \|_{L^2(\Omega)},$$

$$\tilde{b}_{2} := n^{2}(n-1)! \sum_{k \ge 0} \frac{1}{\sqrt{p_{k}q_{k}}} \|f_{n}(*,k)\|_{\ell_{\mathfrak{s}}^{2}(\Delta_{n-1})}^{2} + n^{3}\sqrt{(n-1)!} \sum_{k \ge 0} \frac{1}{p_{k}q_{k}} \|f_{n}(*,k)\|_{\ell_{\mathfrak{s}}^{2}(\Delta_{n-1})} \sqrt{\mathbb{E}[|J_{n-1}(f_{n}(*,k))|^{4}]}.$$
 (8.1)

The next theorem follows from the computations in the proof of Theorem 5.3.

**Theorem 8.1.** Let  $n \geq 2$  be fixed and let  $f_n \in \ell^2_{\mathfrak{s}}(\Delta_n)$ . Assume that for any  $k \in \mathbb{N}$  the functions  $h_{n-1,n-1,s}^{(k)}$  and  $\tilde{h}_{n-1,n-1,s}^{(k)}$  defined by (5.10) and (5.11) belong to  $\ell^2_{\mathfrak{s}}(\Delta_{2n-2-s})$ ,  $0 \leq s \leq 2n-2$ , and that

$$b_{1} := |\lambda - n!| |f_{n}||_{\ell_{s}^{2}(\Delta_{n})}^{2} | + n^{2} \left( \sum_{s=0}^{2n-3} (2n-2-s)! \right)$$

$$\times \sum_{s \leq \{2i_{1}, 2i_{2}\} \leq 2(s \wedge (n-1))} i_{1}! i_{2}! {\binom{n-1}{i_{1}}}^{2} {\binom{n-1}{i_{2}}}^{2} {\binom{i_{1}}{s-i_{1}}} {\binom{i_{2}}{s-i_{2}}}$$

$$\times \sum_{k \geq 0} ||f_{n}(*,k) \otimes_{i_{1}}^{s-i_{1}} f_{n}|_{k}(*)||_{\ell^{2}(\Delta_{2n-2-s})}$$

$$\times \sum_{k \geq 0} ||f_{n}(*,k) \otimes_{i_{2}}^{s-i_{2}} f_{n}|_{k}(*)||_{\ell^{2}(\Delta_{2n-2-s})} \int^{1/2}_{1},$$

and

$$b_{2} := n^{2}(n-1)! \sum_{k \geq 0} \frac{1}{\sqrt{p_{k}q_{k}}} \|f_{n}(*,k)\|_{\ell_{s}^{2}(\Delta_{n-1})}^{2} + n^{3}\sqrt{(n-1)!}$$

$$\times \sum_{k \geq 0} \frac{1}{p_{k}q_{k}} \|f_{n}(*,k)\|_{\ell_{s}^{2}(\Delta_{n-1})} \left(\sum_{s=0}^{2n-2} (2n-2-s)! \sum_{s \leq \{2i_{1}, 2i_{2}\} \leq 2(s \wedge (n-1))} i_{1}!i_{2}! \times {\binom{n-1}{i_{1}}}^{2} {\binom{n-1}{i_{2}}}^{2} {\binom{i_{1}}{s-i_{1}}} {\binom{i_{2}}{s-i_{2}}} \right)$$

$$\times \|f_n(*,k) \otimes_{i_1}^{s-i_1} f_n(*,k)\|_{\ell^2(\Delta_{2n-2-s})} \|f_n(*,k) \otimes_{i_2}^{s-i_2} f_n(*,k)\|_{\ell^2(\Delta_{2n-2-s})} \bigg)^{1/2}$$
(8.2)

are finite. Then we have

$$d_{TV}(F, \operatorname{Po}(\lambda)) \leq \frac{1 - e^{-\lambda}}{\lambda} b_1 + \frac{1 - e^{-\lambda}}{\lambda^2} b_2.$$

8.2. Semigroup bound. By (5.21) and the isometric properties of discrete multiple stochastic integrals we have that the constants  $c_i$  of Corollary 6.4 are equal to

$$\tilde{c}_{1} := |\lambda - n!| \|f_{n}\|_{\ell_{\mathfrak{s}}^{2}(\Delta_{n})}^{2} + n \|\langle J_{n-1}(f_{n}(\ast,\cdot)), J_{n-1}(f_{n}(\ast,\cdot))\rangle_{\ell^{2}(\mathbb{N})} - \mathbb{E}[\langle J_{n-1}(f_{n}(\ast,\cdot)), J_{n-1}(f_{n}(\ast,\cdot))\rangle_{\ell^{2}(\mathbb{N})}] \|_{L^{2}(\Omega)},$$

$$\tilde{c}_2 := \tilde{b}_2$$
, where  $\tilde{b}_2$  is defined by (8.1).

Next theorem follows from the computations in the proof of Theorem 5.4.

**Theorem 8.2.** Let  $n \ge 2$  be fixed. Assume that  $f_n \in \ell^2_{\mathfrak{s}}(\Delta_n)$ , that the functions  $\tilde{h}_{n-1,n-1,s}$  defined by (5.11) belong to  $\ell^2_{\mathfrak{s}}(\Delta_{2n-2-s})$ ,  $0 \le s \le 2n-2$ , and that

$$c_{1} := |\lambda - n!| \|f_{n}\|_{\ell_{s}^{2}(\Delta_{n})}^{2} | \\ + n \left( \sum_{s=0}^{2n-3} (2n-2-s)! \sum_{s \leq \{2i_{1}, 2i_{2}\} \leq 2(s \wedge (n-1))} i_{1}! i_{2}! \right) \\ \times \left( \binom{n-1}{i_{1}}^{2} \binom{n-1}{i_{2}}^{2} \binom{i_{1}}{s-i_{1}} \binom{i_{2}}{s-i_{2}} \right) \\ \times \|f_{n} \otimes_{i_{1}+1}^{s-i_{1}+1} f_{n}\|_{\ell^{2}(\Delta_{2n-2-s})} \|f_{n} \otimes_{i_{2}+1}^{s-i_{2}+1} f_{n}\|_{\ell^{2}(\Delta_{2n-2-s})} \right)^{1/2},$$

and  $c_2 := b_2$ , where  $b_2$  is defined by (8.2), are finite. Then we have

$$d_{TV}(F, \operatorname{Po}(\lambda)) \le \frac{1 - \mathrm{e}^{-\lambda}}{\lambda} c_1 + \frac{1 - \mathrm{e}^{-\lambda}}{\lambda^2} c_2.$$

Quadratic functionals. In the next proposition, we provide an explicit bound for  $d_{TV}(\lambda + J_2(f), \text{Po}(\lambda)), \lambda > 0$ . The proof is omitted since it is a simple consequence of Theorem 8.2 with n = 2, Condition (5.35) and (5.49), (5.51), (5.52) and (5.53). We also note that the integrability condition required in Theorem 8.2 can be checked as in the proof of Proposition 5.8 from Lemma 2.4(1) in Nourdin et al. (2010).

**Proposition 8.3.** Assume (5.35),  $f, f^2 \in \ell^2_{\mathfrak{s}}(\Delta_2)$  and suppose that the shifted second chaos  $\lambda + J_2(f)$  is  $\mathbb{N}$ -valued. Then we have

$$d_{TV}(\lambda + J_2(f), \operatorname{Po}(\lambda)) \le \frac{1 - e^{-\lambda}}{\lambda} d_1 + \frac{1 - e^{-\lambda}}{\lambda^2} d_2,$$

where

$$d_1 := |\lambda - 2||f||^2_{\ell^2_s(\Delta_2)}| + 2\sqrt{2}||f \star^1_1 f||_{\ell^2(\mathbb{N})^{\otimes 2}} + \frac{1 - 2\varepsilon}{\varepsilon}||f \star^1_2 f||_{\ell^2(\mathbb{N})^{\otimes 2}}$$

and

$$d_2 := \frac{4}{\varepsilon} \|f\|_{\ell^2_s(\Delta_2)}^2 + \frac{8}{\varepsilon^2} (\sqrt{2} + 1) \|f\|_{\ell^2_s(\Delta_2)}^2 \|f\star_1^1 f\|_{\ell^2(\mathbb{N})^{\otimes 2}}^{1/2}$$

+ 
$$\frac{4(1-2\varepsilon)}{\varepsilon^3} \|f\|_{\ell^2_{\mathfrak{s}}(\Delta_2)} \|f^2\|_{\ell^2_{\mathfrak{s}}(\Delta_2)}.$$

#### Example

Let  $m \geq 2$  be a fixed integer, and suppose that  $p_k = 1/m$ , for any  $k \in \mathbb{N}$ . Define the function

$$f(k_1, k_2) = \frac{m-1}{m} \mathbb{1}_{\{k_1=0\}} \mathbb{1}_{[1,m]}(k_2) + \frac{m-1}{m} \mathbb{1}_{\{k_2=0\}} \mathbb{1}_{[1,m]}(k_1), \qquad k_1, k_2 \in \mathbb{N}$$

and let  $\lambda$  be an integer bigger than or equal to 4m. We are going to check that all the conditions of Proposition 8.3 are satisfied. Since (5.35) and the integrability of f and  $f^2$  are obvious, we only check that  $\lambda + J_2(f)$  is N-valued. Let  $(Z_k)_{k\geq 0}$  be a sequence of independent and  $\{0, 1\}$ -valued random variables with  $E[Z_k] = p_k = 1/m$ . We have

$$\begin{aligned} H_2(f) &= \sum_{(k_1,k_2)\in\Delta_2} f(k_1,k_2)Y_{k_1}Y_{k_2} \\ &= \frac{2(m-1)}{m}Y_0\sum_{k=1}^m Y_k \\ &= \frac{2(m-1)}{m}\left(\frac{mZ_0-1}{\sqrt{m-1}}\right)\sum_{k=1}^m \left(\frac{mZ_k-1}{\sqrt{m-1}}\right) \\ &= \frac{2}{m}(mZ_0-1)\sum_{k=1}^m (mZ_k-1) \\ &= 2(mZ_0-1)\left(\sum_{k=1}^m Z_k-1\right) \ge -4m. \end{aligned}$$

So  $J_2(f)$  is a  $\mathbb{Z}$ -valued random variable bounded below by -4m. Therefore, by the choice of  $\lambda$ , the shifted second chaos is  $\mathbb{N}$ -valued.

Again, the above example will also satisfy the hypotheses of Theorem 8.1 with n = 2, while it is difficult in general to compare the constants in Theorems 8.1 and 8.2.

#### 9. Proof of the multiplication formula

In this section we prove the multiplication formula (5.3) for (possibly nonsymmetric) discrete multiple stochastic integrals. In order to explain and prove such a formula we shall use the notion of continuous-time normal martingale.

Continuous-time normal martingales. Given  $\hat{\varphi}:\mathbb{R}\longrightarrow\mathbb{R}$  a deterministic function, let

$$i_t = \mathbb{1}_{\{\hat{\varphi}(t)=0\}}, \qquad j_t = 1 - i_t = \mathbb{1}_{\{\hat{\varphi}(t)\neq 0\}}, \qquad t \in \mathbb{R}_+,$$
  
and consider the martingale  $(M_t)_{t \in \mathbb{R}_+}$  represented as

$$dM_t = i_t dB_t + \hat{\varphi}(t)(dN_t - \lambda_t dt), \qquad t \in \mathbb{R}_+, \quad M_0 = 0, \tag{9.1}$$

with  $\lambda_t = (1 - i_t)/\hat{\varphi}^2(t)$ ,  $t \in \mathbb{R}_+$ , where  $(B_t)_{t \in \mathbb{R}_+}$  is a standard Brownian motion, and  $(N_t)_{t \in \mathbb{R}_+}$  is a Poisson process independent of  $(B_t)_{t \in \mathbb{R}_+}$ , with intensity  $\lambda_t$ . Then the martingale  $(M_t)_{t \in \mathbb{R}_+}$  has deterministic angle bracket  $\langle M, M \rangle_t = t$  and it solves the structure equation

$$d[M,M]_t = dt + \hat{\varphi}(t)dM_t, \qquad t \in \mathbb{R}_+, \tag{9.2}$$

cf. § 2.10 of Privault (2009). Here  $([M, M]_t)_{t \in \mathbb{R}_+}$  denotes the quadratic variation of  $(M_t)_{t \in \mathbb{R}_+}$ . Note that the continuous part of  $(M_t)_{t \in \mathbb{R}_+}$  is given by  $dM_t^c = i_t dM_t$ and the eventual jump of  $(M_t)_{t \in \mathbb{R}_+}$  at time  $t \in \mathbb{R}_+$  is given by  $\Delta M_t = \hat{\varphi}(t)$  on  $\{\Delta M_t \neq 0\}, t \in \mathbb{R}_+$ , see Émery (1989), p. 70.

In the following, we denote by  $L^2(\mathbb{R}^{\circ n}_+)$  the subspace of  $L^2(\mathbb{R}^n_+)$  made of symmetric functions in n variables. The multiple stochastic integral  $I_n(f_n)$  is defined by

$$I_n(f_n) = n! \int_0^\infty \int_0^{t_n^-} \cdots \int_0^{t_2^-} f_n(t_1, \dots, t_n) dM_{t_1} \cdots dM_{t_n}, \quad f_n \in L^2(\mathbb{R}^{\circ n}_+), \quad n \ge 1$$

and the following isometry formula holds

$$E[I_n(f_n)I_m(g_m)] = n! \mathbb{1}_{\{n=m\}} \langle f_n, g_m \rangle_{L^2(\mathbb{R}^{\circ n}_+)},$$
(9.3)

where the symbol  $\langle \cdot, \cdot \rangle_{L^2(\mathbb{R}^{\circ n}_+)}$  denotes the usual inner product on  $L^2(\mathbb{R}^{\circ n}_+)$ . For any  $\hat{f}_n \in L^2(\mathbb{R}^{\circ n}_+)$  and  $\hat{g}_m \in L^2(\mathbb{R}^{\circ m}_+)$  the contraction  $\hat{f}_n \hat{\circ}_k^l \hat{g}_m$ ,  $0 \leq l \leq k$ , is defined to be the symmetrization of the function

$$(x_{l+1},\ldots,x_n,y_{k+1},\ldots,y_m) \longmapsto$$

$$\hat{\varphi}(x_{l+1})\cdots\hat{\varphi}(x_k) \int_{\mathbb{R}^l_+} \hat{f}_n(x_1,\ldots,x_n)\hat{g}_m(x_1,\ldots,x_k,y_{k+1},\ldots,y_m) \,\mathrm{d}x_1\cdots\mathrm{d}x_l$$

$$(9.4)$$

in n + m - k - l real variables. We recall the multiplication formula in the general context of normal martingales

$$I_n(\hat{f}_n)I_m(\hat{g}_m) = \sum_{s=0}^{2(n \wedge m)} I_{n+m-s}(\hat{h}_{n,m,s}),$$
(9.5)

cf. Proposition 2 of Privault (1996) or Proposition 4.5.6 of Privault (2009), provided the functions

$$\hat{h}_{n,m,s} := \sum_{s \le 2i \le 2(s \land n \land m)} i! \binom{n}{i} \binom{m}{i} \binom{i}{s-i} \hat{f}_n \hat{\circ}_i^{s-i} \hat{g}_m$$

belong to  $L^2(\mathbb{R}^{\circ n+m-s}_+)$ ,  $0 \le s \le 2(n \land m)$ , and we remark that (9.5) is of the same form as (5.3).

For later purposes, we provide the relation between the contraction  $\hat{f}_n \hat{\circ}_k^l \hat{g}_m$  and the one defined by (5.2). Given  $f_n \in \ell_{\mathfrak{s}}^2(\Delta_n)$  and  $g_m \in \ell_{\mathfrak{s}}^2(\Delta_m)$ , we let

$$\hat{f}_n(x_1,\dots,x_n) := \sum_{a_1,\dots,a_n \in \mathbb{N}} f_n(a_1,\dots,a_n) \mathbb{1}_{[a_1,a_1+1)}(x_1) \cdots \mathbb{1}_{[a_n,a_n+1)}(x_n), \quad (9.6)$$

 $x_1, \ldots, x_n \in \mathbb{R}_+$ , and

$$\hat{g}_m(y_1,\ldots,y_m) := \sum_{b_1,\ldots,b_m \in \mathbb{N}} g_m(b_1,\ldots,b_m) \mathbb{1}_{[b_1,b_1+1)}(y_1) \cdots \mathbb{1}_{[b_m,b_m+1)}(y_m), \quad (9.7)$$

 $y_1, \ldots, y_m \in \mathbb{R}_+$ , and

$$\hat{\varphi}(y) := \sum_{k \in \mathbb{N}} \varphi(k) \mathbb{1}_{[k,k+1)}(y), \qquad y \in \mathbb{R}_+$$
(9.8)

where  $\varphi$  is defined in (5.1). Then

$$f_n \circ_k^l g_m(a_1, \dots, a_{n+m-k-l}) =$$

$$\mathbf{1}_{\Delta_{n+m-k-l}}(a_1, \dots, a_{n+m-k-l}) \hat{f}_n \circ_k^l \hat{g}_m(a_1, \dots, a_{n+m-k-l}).$$
(9.9)

Proof of Proposition 5.1. By the definition of the multiple stochastic integral, the contraction and the structure equation (2.1), for any  $f_n \in \ell^2_{\mathfrak{s}}(\Delta_n)$  and  $g \in \ell^2(\mathbb{N})$ , we deduce

$$J_{n}(f_{n})J_{1}(g) = \sum_{i_{n+1}=0}^{\infty} \sum_{i_{1}\neq i_{2}\neq\cdots\neq i_{n}} f_{n}(i_{1},\ldots,i_{n})g(i_{n+1})Y_{i_{1}}\ldots Y_{i_{n}}Y_{i_{n+1}}$$

$$= \sum_{i_{1}\neq i_{2}\neq\cdots\neq i_{n}\neq i_{n+1}} f_{n}(i_{1},\ldots,i_{n})g(i_{n+1})Y_{i_{1}}\ldots Y_{i_{n}}Y_{i_{n+1}}$$

$$+n\sum_{i_{1}\neq i_{2}\neq\cdots\neq i_{n}} f_{n}(i_{1},\ldots,i_{n})g(i_{n})Y_{i_{1}}\ldots Y_{i_{n-1}}Y_{i_{n}}^{2}$$

$$= \sum_{i_{1}\neq i_{2}\neq\cdots\neq i_{n}\neq i_{n+1}} f_{n}(i_{1},\ldots,i_{n})g(i_{n+1})Y_{i_{1}}\ldots Y_{i_{n}}Y_{i_{n+1}}$$

$$+n\sum_{i_{1}\neq i_{2}\neq\cdots\neq i_{n}} \varphi(i_{n})f_{n}(i_{1},\ldots,i_{n})g(i_{n})Y_{i_{1}}\ldots Y_{i_{n}}$$

$$= \sum_{i_{1}\neq i_{2}\neq\cdots\neq i_{n}\neq i_{n+1}} f_{n}(i_{1},\ldots,i_{n})g(i_{n+1})Y_{i_{1}}\ldots Y_{i_{n}}Y_{i_{n+1}}$$

$$+n\sum_{i_{1}\neq i_{2}\neq\cdots\neq i_{n}} f_{n}(i_{1},\ldots,i_{n})g(i_{n})Y_{i_{1}}\ldots Y_{i_{n}}Y_{i_{n+1}}$$

$$+n\sum_{i_{1}\neq i_{2}\neq\cdots\neq i_{n}} \varphi(i_{n})f_{n}(i_{1},\ldots,i_{n})g(i_{n})Y_{i_{1}}\ldots Y_{i_{n}}$$

$$+n\sum_{i_{1}\neq i_{2}\neq\cdots\neq i_{n}} \sum_{i_{n}=0}^{\infty} f_{n}(i_{1},\ldots,i_{n})g(i_{n})Y_{i_{1}}\ldots Y_{i_{n-1}}$$

$$= J_{n+1}(f_{n} \circ_{0}^{0}g) + nJ_{n}(f_{n} \circ_{1}^{0}g) + nJ_{n-1}(f_{n} \circ_{1}^{1}g), \quad (9.10)$$

which is exactly (5.3) with  $g_m = g_1 = g$ . Next, consider  $h_i(k) = \mathbb{1}_{\{d_i\}}(k)$ , i = 1, ..., m,  $d_i \neq d_j$ ,  $1 \leq i \neq j \leq m$ , and let  $g_m = h_1 \circ_0^0 \cdots \circ_0^0 h_m$ , i.e.  $J_m(g_m) = J_1(h_1) \cdots J_1(h_m)$ . We shall show (5.3) by induction on m = 1, ..., n. We already proved that (5.3) holds for m = 1. Next, assuming that (5.3) holds at the rank  $m \in \{2, \ldots, n-1\}$  we have

$$J_{n}(f_{n})J_{m+1}(g_{m+1}) = J_{n}(f_{n})J_{m}(g_{m})J_{1}(h_{m+1})$$

$$= \sum_{s=0}^{2m} J_{n+m-s}(h_{n,m,s})J_{1}(h_{m+1})$$

$$= \sum_{s=0}^{2m} J_{n+m-s+1}(h_{n,m,s}\circ_{0}^{0}h_{m+1}) + \sum_{s=0}^{2m} (n+m-s)J_{n+m-s}(h_{n,m,s}\circ_{1}^{0}h_{m+1})$$

$$+ \sum_{s=0}^{2m} (n+m-s)J_{n+m-s-1}(h_{n,m,s}\circ_{1}^{1}h_{m+1})$$

$$= \sum_{s=0}^{2m} J_{n+m-s+1}(h_{n,m,s}\circ_{0}^{0}h_{m+1})$$

$$+ \sum_{s=1}^{2m} (n+m+1-s)J_{n+m+1-s}(h_{n,m,s-1}\circ_{1}^{0}h_{m+1})$$

$$+\sum_{s=2}^{2+2m} (n+m+2-s)J_{n+m+1-s}(h_{n,m,s-2}\circ_1^1 h_{m+1})$$
$$=\sum_{s=0}^{2m+2} J_{n+m+1-s}(h_{n,m+1,s}),$$

since

 $h_{n,m+1,s} = \mathbb{1}_{\{0 \le s \le 2m\}} h_{n,m,s} \circ_0^0 h_{m+1}$  $+ \mathbb{1}_{\{1 \le s \le 1+2m\}} (n+m+1-s) h_{n,m,s-1} \circ_1^0 h_{m+1}$  $+ \mathbb{1}_{\{2 \le s \le 2+2m\}} (n+m+2-s) h_{n,m,s-2} \circ_1^1 h_{m+1},$ (9.11)

as follows from Lemma 9.1 below. We have shown that (9.5) holds for any  $g_m$  of the form

$$g_m = 1\!\!1_{\{d_1\}} \circ^0_0 \cdots \circ^0_0 1\!\!1_{\{d_m\}}$$

and the formula extends to all  $g_m \in \ell^2_{\mathfrak{s}}(\Delta_m)$  by summation and linearity. The proof is completed.

**Lemma 9.1.** The identity (9.11) holds for  $g_m = h_1 \circ_0^0 \cdots \circ_0^0 h_m$  and  $h_i(k) = \mathbb{1}_{\{d_i\}}(k)$ , i = 1, ..., m.

*Proof.* Letting  $\hat{h}_i(x) := \mathbb{1}_{[d_i, d_i+1)}(x), i = 1, ..., m$ , and  $\hat{g}_m = \hat{h}_1 \circ_0^0 \cdots \circ_0^0 \hat{h}_m$ , by (9.5) we have

$$I_n(\hat{f}_n)I_1(\hat{g}_1) = I_{n+1}(\hat{f}_n \hat{\circ}_0^0 \hat{g}_1) + nI_n(\hat{f}_n \hat{\circ}_1^0 \hat{g}_1) + nI_{n-1}(\hat{f}_n \hat{\circ}_1^1 \hat{g}_1).$$
(9.12)

By (9.5) and (9.12) it follows that

$$\begin{split} &\sum_{s=0}^{2m+2} I_{n+m+1-s}(\hat{h}_{n,m+1,s}) = I_n(\hat{f}_n)I_{m+1}(\hat{g}_{m+1}) \\ &= I_n(\hat{f}_n)I_m(\hat{g}_m)I_1(\hat{h}_{m+1}) \\ &= \sum_{s=0}^{2m} I_{n+m-s}(\hat{h}_{n,m,s})I_1(\hat{h}_{m+1}) \\ &= \sum_{s=0}^{2m} I_{n+m-s+1}(\hat{h}_{n,m,s}\hat{\circ}_0^0\hat{h}_{m+1}) + \sum_{s=0}^{2m} (n+m-s)I_{n+m-s}(\hat{h}_{n,m,s}\hat{\circ}_1^0\hat{h}_{m+1}) \\ &+ \sum_{s=0}^{2m} (n+m-s)I_{n+m-s-1}(\hat{h}_{n,m,s}\hat{\circ}_1^1\hat{h}_{m+1}) \\ &= \sum_{s=0}^{2m} I_{n+m-s+1}(\hat{h}_{n,m,s}\hat{\circ}_0^0\hat{h}_{m+1}) \\ &+ \sum_{s=1}^{2m} (n+m+1-s)I_{n+m+1-s}(\hat{h}_{n,m,s-1}\hat{\circ}_1^0\hat{h}_{m+1}) \\ &+ \sum_{s=2}^{2+2m} (n+m+2-s)I_{n+m+1-s}(\hat{h}_{n,m,s-2}\hat{\circ}_1^1\hat{h}_{m+1}), \end{split}$$

which, due to the isometry property of the multiple stochastic integrals  $I_k$ , shows that the identity (9.11) holds with  $\hat{h}_{n,m,s}$  and  $\hat{h}_{m+1}$  in place of  $h_{n,m,s}$  and  $h_{m+1}$ , and with  $\hat{\varphi}$  defined from (9.8). Using (9.9) we conclude that (9.11) holds for  $h_{n,m,s}$  and  $h_{m+1}$  as well, and in this case the identity holds for all non-diagonal terms while all functions in the relation vanish on the diagonals.

## Acknowledgment

This work was initiated during a visit of GLT to the School of Physics and Mathematical Sciences of the Nanyang Technological University with the support of NTU MOE Tier 2 Grant No. M4020140.

# References

- A.D. Barbour. Stein's method for diffusion approximations. Probab. Theory Related Fields 84 (3), 297–322 (1990). MR1035659.
- A.D. Barbour, L. Holst and S. Janson. Poisson approximation, volume 2 of Oxford Studies in Probability. The Clarendon Press, Oxford University Press, New York (1992). ISBN 0-19-852235-5. Oxford Science Publications. MR1163825.
- J. Blank, P. Exner and M. Havlíček. *Hilbert space operators in quantum physics*. AIP Series in Computational and Applied Mathematical Physics. American Institute of Physics, New York (1994). ISBN 1-56396-142-3. MR1275370.
- S. Chatterjee. A new method of normal approximation. Ann. Probab. 36 (4), 1584–1610 (2008). MR2435859.
- O. Costin and J.L. Lebowitz. Gaussian fluctuation in random matrices. *Phys. Rev. Lett.* **75** (1), 69–72 (1995). MR3155254.
- F. Daly. Upper bounds for Stein-type operators. *Electron. J. Probab.* 13, no. 20, 566–587 (2008). MR2399291.
- R.M. Dudley. Real analysis and probability, volume 74 of Cambridge Studies in Advanced Mathematics. Cambridge University Press, Cambridge (2002). ISBN 0-521-00754-2. Revised reprint of the 1989 original. MR1932358.
- M. Émery. On the Azéma martingales. In Séminaire de Probabilités, XXIII, volume 1372 of Lecture Notes in Math., pages 66–87. Springer, Berlin (1989). MR1022899.
- T. Erhardsson. Stein's method for Poisson and compound Poisson approximation. In An introduction to Stein's method, volume 4 of Lect. Notes Ser. Inst. Math. Sci. Natl. Univ. Singap., pages 61–113. Singapore Univ. Press, Singapore (2005). MR2235449.
- F. Götze. On the rate of convergence in the multivariate CLT. Ann. Probab. 19 (2), 724–739 (1991). MR1106283.
- K. Krokowski, A. Reichenbachs and C. Thaele. Discrete malliavin-stein method: Berry-esseen bounds for random graphs, point processes and percolation. ArXiv Mathematics e-prints (2015). arXiv: 1208.4264.
- O. Macchi. The coincidence approach to stochastic point processes. Advances in Appl. Probability 7, 83–122 (1975). MR0380979.
- I. Nourdin and G. Peccati. Stein's method on Wiener chaos. Probab. Theory Related Fields 145 (1-2), 75–118 (2009). MR2520122.
- I. Nourdin, G. Peccati and G. Reinert. Stein's method and stochastic analysis of Rademacher functionals. *Electron. J. Probab.* 15, no. 55, 1703–1742 (2010). MR2735379.
- G. Peccati. The Chen-Stein method for Poisson functionals. ArXiv Mathematics e-prints (2011). arXiv: 1112.5051v3.

- G. Peccati, J.L. Solé, M.S. Taqqu and F. Utzet. Stein's method and normal approximation of Poisson functionals. Ann. Probab. 38 (2), 443–478 (2010). MR2642882.
- N. Privault. On the independence of multiple stochastic integrals with respect to a class of martingales. C. R. Acad. Sci. Paris Sér. I Math. 323 (5), 515–520 (1996). MR1408987.
- N. Privault. Stochastic analysis of Bernoulli processes. Probab. Surv. 5, 435–483 (2008). MR2476738.
- N. Privault. Stochastic analysis in discrete and continuous settings with normal martingales, volume 1982 of Lecture Notes in Mathematics. Springer-Verlag, Berlin (2009). ISBN 978-3-642-02379-8. MR2531026.
- N. Privault and G.L. Torrisi. Probability approximation by Clark-Ocone covariance representation. *Electron. J. Probab.* 18, no. 91, 25 (2013). MR3126574.
- T. Shirai. Large deviations for the fermion point process associated with the exponential kernel. J. Stat. Phys. 123 (3), 615–629 (2006). MR2252160.
- A. Soshnikov. Determinantal random point fields. Uspekhi Mat. Nauk 55 (5(335)), 107–160 (2000). MR1799012.
- A. Soshnikov. Gaussian limit for determinantal random point fields. Ann. Probab. 30 (1), 171–187 (2002). MR1894104.
- C. Stein. A bound for the error in the normal approximation to the distribution of a sum of dependent random variables. In Proceedings of the Sixth Berkeley Symposium on Mathematical Statistics and Probability (Univ. California, Berkeley, Calif., 1970/1971), Vol. II: Probability theory, pages 583–602. Univ. California Press, Berkeley, Calif. (1972). MR0402873.
- C. Stein. Approximate computation of expectations. Institute of Mathematical Statistics Lecture Notes—Monograph Series, 7. Institute of Mathematical Statistics, Hayward, CA (1986). ISBN 0-940600-08-0. MR882007.