



General Contact Process with Rapid Stirring

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Abstract. We study the limiting behavior of an interacting particle system evolving on the lattice \mathbb{Z}^d for $d \geq 3$. The model is known as the contact process with rapid stirring. The process starts with a single particle at the origin. Each particle may die, jump to a neighboring site if it is vacant or split. In the case of splitting, one of the offspring takes the place of the parent while the other, the newborn particle, is sent to another site in \mathbb{Z}^d according to a certain distribution; if the newborn particle lands on an occupied site, its birth is suppressed. We study the asymptotic behavior of the critical branching rate as the jumping rate (also known as the stirring rate) approaches infinity.

1. Introduction

The basic contact process with rapid stirring has state space $S = \{0, 1\}^{\mathbb{Z}^d}$ where 1 at site $x \in \mathbb{Z}^d$ means that site x is occupied by a particle and 0 means that the site is empty. Each particle dies with rate 1 and splits at rate λ . If a split occurs at site x then one of the children replaces its parent at site x while the other child is sent to a site y chosen uniformly within the nearest neighboring sites of x . In addition, each pair of neighboring states exchange values at rate N (stirring).

The asymptotic behavior of this and related processes was studied by [Durrett and Neuhauser \(1994\)](#). They obtained results about the existence of phase transition, for various systems, when the stirring rate goes to infinity. Let λ_c^N denotes the infimum of the rates for which the process survives with positive probability. For the basic contact process with stirring, starting with a single particle at the origin, [Durrett and Neuhauser](#) showed that $\lambda_c^N \rightarrow 1$ as $N \rightarrow \infty$.

Later on, [Konno \(1995\)](#) improved the result of [Durrett and Neuhauser](#) by getting a more detailed description of the asymptotic behavior of λ_c^N . His main result is as

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follows. For $N > 0$ define:

$$\phi_d(N) = \begin{cases} \frac{1}{N^{\frac{1}{3}}} & \text{if } d = 1, \\ \frac{\log N}{N} & \text{if } d = 2, \\ \frac{1}{N} & \text{if } d > 2. \end{cases}$$

Then we have

$$\lambda_c^N \approx 1 + C_* \phi_d(N). \quad (1.1)$$

Here \approx means that if $C_* > 0$ is small (large) then the right-hand side of (1.1) is a lower (upper) bound of the left-hand side for all large enough N . After that, [Katori \(1994\)](#) got bounds on C_* for dimensions $d \geq 3$. He showed:

$$\frac{1}{2d(2d-1)} \leq \liminf_{N \rightarrow \infty} (\lambda_c^N - 1)N \leq \limsup_{N \rightarrow \infty} (\lambda_c^N - 1)N \leq \frac{G(0,0) - 1}{2d},$$

where $G(\cdot, \cdot)$ is the Green function of the simple random walk on \mathbb{Z}^d .

Later on, [Berezin and Mytnik \(2014\)](#) improved the results of Konno and Katori by getting more precise information on the asymptotic behavior of λ_c^N in dimensions $d \geq 3$. They improved the lower bound on λ_c^N and by using Katori's result in [Katori \(1994\)](#) they got the sharp asymptotic behavior of the critical λ_c . Their main result was:

$$\lim_{N \rightarrow \infty} \frac{\lambda_c^N - 1}{\frac{\vartheta}{N}} = 1, \quad \text{where } \vartheta = \frac{G(0,0) - 1}{2d}. \quad (1.2)$$

Finally, [Levit and Valesin \(2017\)](#) studied the same model for $d = 2$ and established a lower bound for λ_c^N . Their main result for $d = 2$, was

$$\liminf_{N \rightarrow \infty} \frac{\lambda_c^N - 1}{\frac{\log N}{N}} \geq \frac{1}{4\pi}.$$

The aim of this paper is to study the limiting behavior of λ_c^N for a more general case. In particular, we generalize the model as follow: the location of one of the offspring is distributed according to some symmetric, probability function P^b which satisfies Assumption 2.2 below, and not according to the uniform distribution on its $2d$ closest neighboring as in the nearest neighbor contact process.

It is worth mentioning that in order to prove the upper bound in the nearest neighboring case, Katori used the binary contact path process with exchange. This is a class of interacting particle systems called the linear system (see [Liggett, 1985](#), Chapter IX). However, this technique is no longer available in the general case; therefore, we believe that there would be exceptional difficulties generalizing the matching upper bound.

The rest of the paper is organized as follows: In Section 2 we present the model and state our main result. Section 3 is dedicated to the proof of our main result. This proof is based on two propositions which we prove in Sections 4 and 5.

2. Model and Main Result

The contact process with rapid stirring is defined on the lattice \mathbb{Z}^d . Let N be a ‘‘large’’ integer parameter, which in the sequel we will send to infinity. The state of the process at time t is given by the function $\xi_t^N : \mathbb{Z}^d \rightarrow \{0, 1\}$, where the value of $\xi_t^N(x)$ represents the number of particles present at x at time t (it equals zero or one). Starting with a single particle at the origin, the basic rules for the evolution of the process are:

- (1) particles die at rate 1 without producing offspring;
- (2) particles split into two at rate λ . If a split occurs at $x \in \mathbb{Z}^d$, then one of the offspring replaces the parent, while the other is sent to a site y with probability $P^b(x, y)$. If a newborn particle lands on an occupied site, its birth is suppressed;
- (3) for each $x, y \in \mathbb{Z}^d$ such that $\|x - y\|_1 = 1$ the values of ξ^N at x and y are exchanged at rate N ;
- (4) the above mechanisms are independent.

Remark 2.1. We say that events occur at some rate if the times between events are independent and exponentially distributed with that rate.

In the rest of this paper we will frequently use the following assumption about the probability function $P^b(x, y)$.

Assumption 2.2. Let $P^b(x, y)$ be a function of $\|x - y\|_1$. It means that there exists a function f such that $P^b(x, y) = f(\|x - y\|_1)$.

Remark 2.3. The process is an interacting particle system with long range interactions (depending on P^b). It is easy to check that Condition (0.3) in Chapter 3 in [Liggett \(1985\)](#) is satisfied and thus by Theorems 3.9 and 1.5 in Chapter 1 in [Liggett \(1985\)](#) the process ξ_t^N exists and, in fact, it is Markov.

Before we present the main result let us introduce the following notation:

- Let $|\xi_t^N| = \sum_{x \in \mathbb{Z}^d} \xi_t^N(x)$ denote the total number of particles at time t .
- Let $\Omega_\infty^N = \{|\xi_t^N| > 0 \quad \forall t \geq 0\}$ denote the event that the process survives.
- Let $\rho_\lambda^N = P(\Omega_\infty^N)$ denote the probability that the process survives.
- Let $\lambda_c^N = \inf\{\lambda \geq 0 : \rho_\lambda^N > 0\}$ denote the infimum of the rates for which the process survives with positive probability (also known as the critical branching rate).

Note that ρ_λ^N describes the probability of survival of the process. For all λ smaller than λ_c^N the process dies out almost surely, and for λ greater than λ_c^N the process survives with probability ρ_λ^N .

Our main result gives a lower bound on the critical branching rate. It is presented in the following theorem.

Theorem 2.4. *Let $d \geq 3$ and assume that $P^b(x, y)$ satisfies Assumption 2.2. Let $G(\cdot, \cdot)$ denote the Green function of the simple random walk on \mathbb{Z}^d . Then,*

$$\liminf_{N \rightarrow \infty} \frac{\lambda_c^N - 1}{\frac{\vartheta}{N}} \geq 1,$$

where

$$\vartheta \equiv (2d)^{-1} \sum_{x, y \in \mathbb{Z}^d} P^b(0, x) P^b(0, y) G(x, y). \quad (2.1)$$

Our main theorem shows that the asymptotic behavior of the critical rate is at least $1 + \frac{\vartheta}{N}$.

Conjecture 2.5. *We conjecture that in our general model the lower bound is tight. That is, we conjecture that $\limsup_{n \rightarrow \infty} \frac{\lambda_c^N - 1}{\vartheta} = 1$ and thus,*

$$\lim_{n \rightarrow \infty} \frac{\lambda_c^N - 1}{\frac{\vartheta}{N}} = 1.$$

Special case considered in Berezin and Mytnik (2014). Let us show that our main result is consistent with the result of Berezin and Mytnik who considered a special case of our model. In Berezin and Mytnik (2014) the same model is considered with a particular distribution of the location of a particle after splitting: when a particle splits its child's location is distributed uniformly over its closest neighbors. That is, the branching probability function is $P^b(y, x) = (2d)^{-1} \mathbb{1}_{\{\|x-y\|_1=1\}}$. Let $\varphi_d = \{x \in \mathbb{Z}^d : \|x\|_1 = 1\}$ denote the set of nearest neighbours of the origin. Thus, according to (2.1),

$$\vartheta = (2d)^{-1} \sum_{x \in \varphi_d} \sum_{y \in \varphi_d} \frac{1}{2d} \frac{1}{2d} G(x, y).$$

For any $x \in \mathbb{Z}^d$, let $G(x, \varphi_d) = \sum_{y \in \varphi_d} G(x, y)$ denote the expected number of times a simple symmetric nearest neighbour random walk visits at the set of sites φ_d , given that it started from site x . Thus,

$$\vartheta = (2d)^{-1} \sum_{x \in \varphi_d} \frac{1}{2d} \frac{1}{2d} G(x, \varphi_d) = (2d)^{-2} \sum_{x \in \varphi_d} \frac{1}{2d} G(x, \varphi_d).$$

Notice that when a simple symmetric nearest neighbour random walk jumps from the origin, then with probability $\frac{1}{2d}$ it jumps to a site x in φ_d . Thus, by using first step analysis we get,

$$\vartheta = (2d)^{-2} G(0, \varphi_d) = \frac{G(0, 0) - 1}{2d}$$

which coincides with the result obtained by Berezin and Mytnik (2014, Theorem 3).

Note that the lower bound in Berezin and Mytnik (2014) is tight. This motivates our Conjecture 2.5.

3. Proof of Main Result

Before we start the proof of the main result let us introduce an additional notation. For any $x, y \in \mathbb{Z}^d$ we say that x, y are ℓ -neighbours if $\|x - y\|_1 = \ell$. In that case, we denote it by $x \underset{\ell}{\sim} y$. Let $h_\ell = \sum_{x \in \mathbb{Z}^d} \mathbb{1}_{\{0 \underset{\ell}{\sim} x\}}$ be the total number of ℓ -neighbours of the origin.

Recall that in this work we focus on the probability function $P^b(x, y)$ that satisfies Assumption 2.2, which means that it depends only on the distance (in l_1 -norm) between x and y . Thus, we may write

$$P^b(x, y) = \sum_{\ell} \frac{p_\ell}{h_\ell} \cdot \mathbb{1}_{\{x \underset{\ell}{\sim} y\}} \quad (3.1)$$

for some discrete distribution p_ℓ on \mathbb{N} . This means that whenever a particle splits one of its offspring jumps to an ℓ -neighbor with probability p_ℓ and its position is distributed uniformly within all the ℓ -neighbors. We will also frequently use the following notation:

- We denote particles by Greek letters α, β, γ .
- Since we start with a single particle and each particle may split into two we use a binary vector representation of particles. For example, particle (1) is the first particle, the one we start the process with; particles (1,0) and (1,1) are the two children of particle (1) and so on.
- For each particle α let $\mathcal{C}_0(\alpha)$ and $\mathcal{C}_1(\alpha)$ denote the two children of particle α . For example, if $\alpha = (1, 1)$ then $\mathcal{C}_0(\alpha) = (1, 1, 0)$ and $\mathcal{C}_1(\alpha) = (1, 1, 1)$. In addition, we say that particle β is the child of particle α if $\beta = \mathcal{C}_0(\alpha)$ or $\beta = \mathcal{C}_1(\alpha)$.
- For each particle α let $\mathcal{P}(\alpha)$ denote the parent of particle α . In addition, we say that particle α is the parent of particle β if $\mathcal{P}(\beta) = \alpha$.

In order to prove Theorem 2.4 we set the branching rate $\lambda = 1 + \frac{\theta}{N}$ and we show that for any $\theta < \vartheta$ the process dies out with probability one, which gives a lower bound for the critical branching rate. It will be more convenient to deal with the speeded-up process

$$\hat{\xi}_t^N = \xi_{Nt}^N. \quad (3.2)$$

This process obeys the same rules as ξ_t^N with the exception that all events occur with rates multiplied by N . Thus, particles die with rate N , split with rate $\theta + N$ and perform stirring with rate N^2 .

The next proposition is crucial for the proof of the main Theorem 2.4. It is proved in Section 4.

Proposition 3.1. *Fix an arbitrary $\theta < \vartheta$. Then there exists $N_\theta > 0$ such that for any $N \geq N_\theta$ and for any $t \geq 0$,*

$$\hat{m}_t^N := \mathbb{E}(|\hat{\xi}_t^N|) \leq \exp \left[\left(\frac{1}{2}(\theta - \vartheta) \right) t + 2 \right]. \quad (3.3)$$

With Proposition 3.1 at hands we are ready to give the proof of Theorem 2.4.

Proof of Theorem 2.4: From Proposition 3.1 it follows that for every $\theta < \vartheta$, uniformly on all N sufficiently large, the expected number of particles approaches zero (exponentially fast) when t approaches infinity. Since the number of particles is a non-negative integer random variable then with probability one the process dies out for every $\theta < \vartheta$. This proves Theorem 2.4. \square

4. Proof of Proposition 3.1

In order to prove Proposition 3.1, first of all, in Section 4.1, we derive the differential equation for the expected number of particles (\hat{m}_t^N): see (4.1) in Lemma 4.2. The last term on the right hand side of (4.1) arises as a result of suppression of births of the particles that are born on the occupied sites. The bounds on this term are crucial for proving Proposition 3.1. The important steps for bounding this term are performed in Section 4.2, and the results from that section allow us to finish the proof of Proposition 3.1 in Section 4.3.

4.1. *Differential Equation for the Expected Number of Particles.* The first ingredient in the proof of Proposition 3.1 is a derivation of the equation for the expected number of particles. To do this we need the following notation:

- Let \mathcal{A}_t denote the set of all particles alive at time t in the process ξ_t^N .

- Let \mathcal{F}_t denote the natural filtration of the process ξ_t^N until time t .
- Let $m_t^N = \mathbb{E}(|\xi_t^N|)$ denote the expected number of particles at time t and let $\hat{m}_t^N = m_{Nt}^N$.
- Let $I_t^{\ell,N} = \mathbb{E}[\sum_{x \sim y} \xi_t^N(x) \xi_t^N(y)]$ denote twice the expected number of ℓ -neighbours alive at time t and let $\hat{I}_t^{\ell,N} = I_{Nt}^{\ell,N}$.

In addition, recall that in the process ξ_t^N , each particle splits at rate $1 + \frac{\theta}{N}$, dies at rate 1 and jumps to each one of its closest vacant neighbouring sites with rate N . Since each particle has $2d$ closest neighbours the total jumping rate of a particle is $2dN$. Thus,

$$\mu := 2 + \frac{\theta}{N} + 2dN$$

denotes the total rate of events (splitting, dying and stirring) that can occur for each alive particle.

Before we present the differential equation for the expected number of particles in the speeded-up process (see (3.2)) let us state a lemma which will be useful for the rest of the paper.

Lemma 4.1. *For any finite time t and any stirring rate N it holds that:*

$$(A1) \quad P(|\xi_t^N| < \infty) = 1 \quad \text{and} \quad m_t^N \leq e^{\mu t} < \infty,$$

and consequently

$$(A2) \quad P(|\hat{\xi}_t^N| < \infty) = 1 \quad \text{and} \quad \hat{m}_t^N \leq e^{\mu N t} < \infty.$$

The proof of this lemma is simple and is based on bounding $|\xi_t^N|$ by the Yule process and the properties of Yule process (see Chapter 2, Example 2.5.1 in [Norris, 1998](#)). In addition, notice that (A2) follows from (A1) when speeding the process by a factor of N . Since the proof is standard it is omitted. The differential equation for the expected number of particles in the speeded-up process, $\{\hat{\xi}_t^N\}_{t \geq 0}$, is given in the next lemma.

Lemma 4.2.

$$\hat{m}_t^N = \hat{m}_0^N + \int_0^t \theta \hat{m}_s^N ds - \left(1 + \frac{\theta}{N}\right) \sum_{\ell} \frac{p_{\ell}}{h_{\ell}} \int_0^t N \hat{I}_s^{\ell,N} ds. \quad (4.1)$$

The analogous results are frequently stated as standard (see page 152 in [Schinazi, 2014](#) and page 842 in [Konno, 1995](#)) so the proof of the lemma is omitted. A complete proof of the lemma can be found in [Shlomov \(2017\)](#). The general idea behind the construction of this equation is to calculate the change of the expectation of the total mass of the process in an infinitesimal period of time. One considers $\frac{d}{dt} m_t^N = \lim_{\epsilon \rightarrow 0} \frac{1}{\epsilon} \mathbb{E}[|\xi_{t+\epsilon}^N| - |\xi_t^N|]$, and shows that this limit exists. One gets an expression for the limit by using the Markov property and conditioning on each particle possible actions on time interval $[t, t + \epsilon]$. Note that on this time interval each particle either jumps or splits or dies, and as $\epsilon \downarrow 0$, the probability that more than one of these events occurs is of order less than ϵ and thus negligible.

Intuitively, the positive term $\int_0^t \theta \hat{m}_s^N ds$ comes from the gap between the splitting rate and the dying rate while the term $\left(1 + \frac{\theta}{N}\right) \sum_{\ell} \frac{p_{\ell}}{h_{\ell}} \int_0^t N \hat{I}_s^{\ell,N} ds$ comes from the loss of mass caused by collisions; collision is an event when one particle splits and the newborn particle lands on the site occupied by another particle.

4.2. *Bounding the expected number of neighbors.* Another ingredient in the proof of Proposition 3.1 is bounding the expected number of ℓ -neighbors. It is clear from Lemma 4.2 that in order to bound from above the expectation of the total mass of the speeded-up process $\hat{\xi}_t^N$, we need to bound from below the expected number of ℓ -neighbors for every ℓ . The main idea that allows us to get a useful lower bound on this quantity at time s is based on counting only particles that are ℓ -neighbors at time s and have a common parent that splits at some time in $[s - \tau_N, s)$, where

$$\tau_N = \frac{\ln(N)}{N^2}. \quad (4.2)$$

The reason for such a definition of τ_N is that two particles need the order of N^{-2} units of time to travel “long” distance so that a collision (recall again that this is an event when one particle splits and the newborn particle lands on the site occupied by another particle) between them, after this time, happens with very small probability. This explains why we choose τ_N a little bit bigger than N^{-2} . A similar technique is used in a number of papers (see [Berezin and Mytnik, 2014](#), [Konno, 1995](#), [Durrett and Perkins, 1999](#)). In Lemmas 4.3 – 4.5 below, we will use this idea in order to bound from below the expected number of ℓ -neighbors.

Before we proceed, we need an additional notation:

- Let $\hat{\mathcal{A}}(t)$ denote the set of all particles alive at time t in the process $\hat{\xi}_t^N$.
- Let $\hat{\mathcal{F}}_t$ denote the natural filtration of the process $\hat{\xi}_t^N$ until time t .
- Let T_α denote the time at which particle α dies or splits in the process $\hat{\xi}_t^N$.
- Let B_t^α denote the position (on the lattice) of particle α at time t in the process $\hat{\xi}_t^N$. If particle α is not alive at time t (dead or not born yet) then we write $B_t^\alpha = \Delta$.
- Let $Z_\alpha^\ell(t) = \mathbb{1}_{\{T_\alpha \in [t, t + \tau_N], B_{t+\tau_N}^{\mathcal{C}_1(\alpha)} \sim B_{t+\tau_N}^{\mathcal{C}_0(\alpha)}, T_{\mathcal{C}_1(\alpha)} \geq t + \tau_N, T_{\mathcal{C}_0(\alpha)} \geq t + \tau_N\}}$ denote the indicator of the event that the two children of particle α created during the time interval $[t, t + \tau_N]$ are ℓ -neighbours at time $t + \tau_N$ and let $Z_1^\ell = Z_1^\ell(0)$.
- Let $\zeta_\alpha(t)$ denote the indicator of the event that one of α 's children created during the time interval $[t, t + \tau_N]$ died at the time of its birth T_α as a result of collision.

Lemma 4.3. *For any $s \geq \tau_N$ and for any $\ell = 1, 2, \dots$*

$$\hat{I}_s^{\ell, N} \geq 2\hat{m}_{s-\tau_N}^N \mathbb{E}[Z_1^\ell] - 2\mathbb{E} \left[\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \zeta_\alpha(s - \tau_N) \right].$$

Proof: Fix arbitrary ℓ and recall the definition of B_s^β . Then, $\hat{I}_s^{\ell, N}$ can be written as follows:

$$\hat{I}_s^{\ell, N} = \mathbb{E} \left[\sum_{\alpha, \beta \in \hat{\mathcal{A}}(s)} \mathbb{1}_{\{B_s^\beta \sim B_s^\alpha\}} \right].$$

We bound the right-hand side from below by summing only pairs of particles that share a common parent who splits during the time interval $[s - \tau_N, s]$. Thus,

$$\hat{I}_s^{\ell, N} \geq 2\mathbb{E} \left[\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} Z_\alpha^\ell(s - \tau_N) \right] \geq 2\mathbb{E} \left[\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \mathbb{E} \left[Z_\alpha^\ell(s - \tau_N) \mid \hat{\mathcal{F}}_{s-\tau_N} \right] \right].$$

Note that given particle α alive at time $s - \tau_N$ and given $\hat{\mathcal{F}}_{s-\tau_N}$ we have for $s \geq \tau_N$

$$\mathbb{E}[Z_\alpha^\ell(s - \tau_N) \mid \hat{\mathcal{F}}_{s-\tau_N}] + \mathbb{E}[\zeta_\alpha(s - \tau_N) \mid \hat{\mathcal{F}}_{s-\tau_N}] \geq \mathbb{E}[Z_1^\ell]. \quad (4.3)$$

Thus,

$$\begin{aligned} \hat{I}_s^{\ell, N} &\geq 2\mathbb{E} \left[\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} (\mathbb{E}[Z_1^\ell] - \mathbb{E}[\zeta_\alpha(s - \tau_N) \mid \hat{\mathcal{F}}_{s-\tau_N}]) \right] \\ &= 2\hat{m}_{s-\tau_N}^N \mathbb{E}[Z_1^\ell] - 2\mathbb{E} \left[\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \zeta_\alpha(s - \tau_N) \right]. \end{aligned}$$

□

Now we need to get an upper bound for $\int_{\tau_N}^t N \mathbb{E} \left[\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \zeta_\alpha(s - \tau_N) \right] ds$. This is done in the next lemma.

Lemma 4.4. *For every $N > N_2 = \theta$ and $t \geq \tau_N$*

$$\int_{\tau_N}^t N \mathbb{E} \left[\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \zeta_\alpha(s - \tau_N) \right] ds \leq \sum_{\ell} \frac{p_\ell}{h_\ell} 4N\tau_N \int_0^t N \hat{I}_r^{\ell, N} dr.$$

Proof: In order to prove this lemma we first need to bound the term inside the integral. That is,

$$\Xi = \mathbb{E} \left[\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \zeta_\alpha(s - \tau_N) \right]. \quad (4.4)$$

To this end, we introduce the following notation: Let T_α^c (resp. T_α^d) denote the potential split (resp. death) time of particle α ; note that $T_\alpha = \min(T_\alpha^c, T_\alpha^d)$. Let J_α be the event that one of the particle α 's children died at the time of its birth as a result of a collision with another particle. In addition, notice that from the memoryless properties of the exponential distribution we can deduce the following property of T_α^c :

(P1) For any time $t \geq 0$, given $\{\hat{\mathcal{F}}_t, \alpha \in \hat{\mathcal{A}}(t)\}$, the random variable $(T_\alpha^c - t)$ has an exponential distribution with rate $\lambda_b := N + \theta$.

We start by bounding Ξ . By the law of total expectation and the linearity of the expectation we have

$$\begin{aligned} \Xi &= \mathbb{E} \left(\mathbb{E} \left[\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \zeta_\alpha(s - \tau_N) \mid \hat{\mathcal{F}}_{s-\tau_N} \right] \right) \\ &= \mathbb{E} \left(\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \mathbb{E} \left[\zeta_\alpha(s - \tau_N) \mid \hat{\mathcal{F}}_{s-\tau_N} \right] \right). \end{aligned}$$

By the definition of $\zeta_\alpha(s - \tau_N)$ we have

$$\Xi = \mathbb{E} \left(\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} P \left[T_\alpha^c \in (s - \tau_N, s), T_\alpha^d > T_\alpha^c, J_\alpha \mid \hat{\mathcal{F}}_{s-\tau_N} \right] \right).$$

Since we sum only over particles which are alive at time $s - \tau_N$ we have

$$\Xi = \mathbb{E} \left(\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} P \left[T_\alpha^c \in (s - \tau_N, s), T_\alpha^d > T_\alpha^c, J_\alpha \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c > s - \tau_N \right] \right).$$

By (P1), conditionally on the splitting time of particle α we get

$$\begin{aligned} \Xi &= \mathbb{E} \left(\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \int_{s-\tau_N}^s \lambda_b e^{-\lambda_b(t-s+\tau_N)} \mathbb{E} \left(\mathbf{1}_{\{J_\alpha, T_\alpha^d > t\}} \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c = t \right) dt \right) \\ &\leq \mathbb{E} \left(\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \int_{s-\tau_N}^s \lambda_b \mathbb{E} \left(\mathbf{1}_{\{J_\alpha, T_\alpha^d > t\}} \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c = t \right) dt \right). \end{aligned}$$

Since the indicator in the expectation above is zero if $T_\alpha^d < t$ we have

$$\begin{aligned} \Xi &\leq \mathbb{E} \left(\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \int_{s-\tau_N}^s \lambda_b \mathbb{E} \left(\mathbf{1}_{\{J_\alpha\}} \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c = t, T_\alpha^d > t \right) \right. \\ &\quad \left. \cdot P \left(T_\alpha^d > t \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c = t \right) dt \right) \end{aligned}$$

Thus,

$$\Xi \leq \mathbb{E} \left(\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \int_{s-\tau_N}^s \lambda_b \mathbb{E} \left(\mathbf{1}_{\{J_\alpha\}} \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c = t, T_\alpha^d > t \right) dt \right).$$

By the tower property, and since T_α^d and T_α^c are continuous random variables we get

$$\begin{aligned} \Xi &\leq \mathbb{E} \left(\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \int_{s-\tau_N}^s \lambda_b \right. \\ &\quad \left. \cdot \mathbb{E} \left[\mathbb{E} \left(\mathbf{1}_{\{J_\alpha\}} \mid F_{t-}, T_\alpha^c = t, T_\alpha^d \geq t \right) \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c = t, T_\alpha^d \geq t \right] dt \right). \end{aligned}$$

Recall that $n_\alpha^\ell(t)$ denotes the number of ℓ -neighbours of particle α alive at time t . In addition, when a particle splits one of its offspring is sent to an ℓ -neighbour with probability p_ℓ and is distributed uniformly within all the ℓ -neighbours. Thus, since the filtration until time $t-$ is given and particle α splits at time t , then the conditional probability that its child lands on an occupied site, given its child is sent to an ℓ -neighbour, is equal to $\frac{n_\alpha^\ell(t)}{h_\ell}$. Therefore, we get

$$\Xi \leq \mathbb{E} \left(\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \int_{s-\tau_N}^s \lambda_b \sum_{\ell} \frac{p_\ell}{h_\ell} \mathbb{E} \left[n_\alpha^\ell(t) \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c \geq t, T_\alpha^d \geq t \right] dt \right).$$

Notice that for each ℓ and a particle α alive at any time t ,

$$\begin{aligned} & \mathbb{E} \left[n_\alpha^\ell(t) \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c \geq t \right] \\ &= \mathbb{E} \left[n_\alpha^\ell(t) \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c \geq t, T_\alpha^d \geq t \right] P \left(T_\alpha^d \geq t \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c \geq t \right) \\ & \quad + \mathbb{E} \left[n_\alpha^\ell(t) \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c \geq t, T_\alpha^d < t \right] P \left(T_\alpha^d < t \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c \geq t \right). \end{aligned}$$

Since $n_\alpha^\ell(t) = 0$ if $T_\alpha^d < t$, we get that

$$\mathbb{E} \left[n_\alpha^\ell(t) \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c \geq t, T_\alpha^d \geq t \right] = \frac{\mathbb{E} \left[n_\alpha^\ell(t) \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c \geq t \right]}{P \left(T_\alpha^d \geq t \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c \geq t \right)}.$$

Thus, for a particle α alive at time $s - \tau_N$ we have

$$\begin{aligned} & P \left(T_\alpha^d \geq t \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c \geq t, T_\alpha^d \geq s - \tau_N \right) \\ & \geq P \left(T_\alpha^d \geq s \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c \geq t, T_\alpha^d \geq s - \tau_N \right) = e^{-N\tau_N} > 0.5, \end{aligned}$$

where the last inequality follows by our choice of $N > 0$. Thus,

$$\Xi \leq \mathbb{E} \left(\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \int_{s-\tau_N}^s 2\lambda_b \sum_\ell \frac{p_\ell}{h_\ell} \mathbb{E} [n_\alpha^\ell(t) \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c \geq t] dt \right).$$

By the linearity of expectation and measurability of $\hat{\mathcal{A}}(s - \tau_N)$ with respect to $\hat{\mathcal{F}}_{s-\tau_N}$ we get

$$\begin{aligned} \Xi & \leq \int_{s-\tau_N}^s 2\lambda_b \sum_\ell \frac{p_\ell}{h_\ell} \mathbb{E} \left(\mathbb{E} \left[\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} n_\alpha^\ell(t) \mid \hat{\mathcal{F}}_{s-\tau_N}, T_\alpha^c \geq t \right] \right) dt \\ & \leq \int_{s-\tau_N}^s 2\lambda_b \sum_\ell \frac{p_\ell}{h_\ell} \mathbb{E} \left(\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} n_\alpha^\ell(t) \right) dt. \end{aligned}$$

Recall that $\hat{I}_t^{\ell, N}$ denotes twice the expected number of ℓ -neighbours alive at time t , and $\lambda_b = N + \theta$. Thus,

$$\Xi \leq \sum_\ell \frac{p_\ell}{h_\ell} 2(N + \theta) \int_{s-\tau_N}^s \hat{I}_t^{\ell, N} dt.$$

Now we can use the bound on Ξ to finish the proof of the lemma. Since $N > \theta$ we have

$$\begin{aligned} \int_{\tau_N}^t N \mathbb{E} \left[\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \zeta_\alpha(s - \tau_N) \right] ds & \leq \int_{\tau_N}^t 2N(N + N) \sum_\ell \frac{p_\ell}{h_\ell} \int_{s-\tau_N}^s \hat{I}_r^{\ell, N} dr ds. \\ & \leq \sum_\ell \frac{p_\ell}{h_\ell} 4N^2 \tau_N \int_0^t \hat{I}_r^{\ell, N} dr, \end{aligned}$$

and we are done. \square

We also need the following two lemmas.

Lemma 4.5. *For any $N > 0$ and $t \geq 0$*

$$\left(1 + \frac{\theta}{N}\right) \sum_{\ell} \frac{p_{\ell}}{h_{\ell}} \int_0^t N \hat{I}_s^{\ell, N} ds \leq 1 + \int_0^t \theta \hat{m}_s^N ds.$$

Proof: From Lemma 4.2 we immediately have

$$\left(1 + \frac{\theta}{N}\right) \sum_{\ell} \frac{p_{\ell}}{h_{\ell}} \int_0^t N \hat{I}_s^{\ell, N} ds = -\hat{m}_t^N + 1 + \int_0^t \theta \hat{m}_s^N ds \leq 1 + \int_0^t \theta \hat{m}_s^N ds$$

□

Lemma 4.6. *For any $N > 0$ and any $s \geq \tau_N$*

$$\hat{m}_{s-\tau_N}^N \geq \hat{m}_s^N e^{-\theta \tau_N}.$$

Proof: From Lemma 4.2 and since $\hat{m}_0^N = 1$ we have

$$\hat{m}_t^N = 1 + \int_0^t \theta \hat{m}_s^N ds - \left(1 + \frac{\theta}{N}\right) \sum_{\ell} \frac{p_{\ell}}{h_{\ell}} \int_0^t N \hat{I}_s^{\ell, N} ds \leq 1 + \int_0^t \theta \hat{m}_s^N ds.$$

Thus, since 1 is a non-decreasing function, and $\hat{m}_r^N < \infty$ for all $r \in [0, t]$ (see Lemma 4.1) Grönwall's lemma immediately implies

$$\hat{m}_t^N \leq \hat{m}_r^N e^{\theta(t-r)} \quad \forall r \in [0, t]. \quad (4.5)$$

□

In order to complete the proof of Proposition 3.1 we need the following proposition, which will be proven in Section 5.

Proposition 4.7. *For every $\epsilon > 0$ there exists $N_{\epsilon} > 0$ such that for every $N > N_{\epsilon}$*

$$\vartheta - \epsilon \leq \sum_{\ell} \frac{p_{\ell}}{h_{\ell}} 2N \mathbb{E}[Z_1^{\ell}].$$

Now we are ready to complete the proof of Proposition 3.1.

4.3. *Proof of Proposition 3.1.* In what follows we assume that $N > N_2$ where N_2 is defined in Lemma 4.4. From Lemma 4.2 we get

$$\hat{m}_t^N = \hat{m}_{\tau_N}^N + \int_{\tau_N}^t \theta \hat{m}_s^N ds - \left(1 + \frac{\theta}{N}\right) \sum_{\ell} \frac{p_{\ell}}{h_{\ell}} \int_{\tau_N}^t N \hat{I}_s^{\ell, N} ds. \quad (4.6)$$

Apply on (4.6) the bound on $\hat{I}_s^{\ell, N}$ that we have from Lemma 4.3 to get

$$\begin{aligned} \hat{m}_t^N &\leq \hat{m}_{\tau_N}^N + \int_{\tau_N}^t \theta \hat{m}_s^N ds \\ &\quad - \left(1 + \frac{\theta}{N}\right) \sum_{\ell} \frac{p_{\ell}}{h_{\ell}} \int_{\tau_N}^t N \left(2\hat{m}_{s-\tau_N}^N \mathbb{E}[Z_1^{\ell}] - 2\mathbb{E} \left[\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \zeta_{\alpha}(s-\tau_N) \right] \right) ds. \end{aligned}$$

Apply the bound on $\int_{\tau_N}^t N \mathbb{E} \left[\sum_{\alpha \in \hat{\mathcal{A}}(s-\tau_N)} \zeta_\alpha(s-\tau_N) \right] ds$ that we have from Lemma 4.4 to get,

$$\begin{aligned} \hat{m}_t^N &\leq \hat{m}_{\tau_N}^N + \int_{\tau_N}^t \theta \hat{m}_s^N ds - \left(1 + \frac{\theta}{N}\right) \sum_\ell \frac{p_\ell}{h_\ell} \int_{\tau_N}^t 2N \hat{m}_{s-\tau_N}^N \mathbb{E}[Z_1^\ell] ds \\ &\quad + 2\left(1 + \frac{\theta}{N}\right) \sum_\ell \frac{p_\ell}{h_\ell} 4N\tau_N \int_0^t N \hat{I}_r^{\ell,N} dr. \end{aligned}$$

Apply the bound on $(1 + \frac{\theta}{N}) \sum_\ell \frac{p_\ell}{h_\ell} \int_0^t N \hat{I}_s^{\ell,N} ds$ from Lemma 4.5, and the bound $\hat{m}_{\tau_N}^N \leq e^{\theta\tau_N}$ from Lemma 4.6 to get

$$\begin{aligned} \hat{m}_t^N &\leq \hat{m}_{\tau_N}^N + \int_{\tau_N}^t \theta \hat{m}_s^N ds - \left(1 + \frac{\theta}{N}\right) \sum_\ell \frac{p_\ell}{h_\ell} \int_{\tau_N}^t 2N \hat{m}_{s-\tau_N}^N \mathbb{E}[Z_1^\ell] ds \\ &\quad + 8N\tau_N \left(1 + \int_0^t \theta \hat{m}_s^N ds\right). \end{aligned}$$

Thus,

$$\begin{aligned} \hat{m}_t^N &\leq \int_{\tau_N}^t \theta \hat{m}_s^N ds - \left(\int_{\tau_N}^t \hat{m}_{s-\tau_N}^N ds\right) 2N \left(1 + \frac{\theta}{N}\right) \sum_\ell \frac{p_\ell}{h_\ell} \mathbb{E}[Z_1^\ell] \\ &\quad + e^{\theta\tau_N} + 8N\tau_N \left(1 + \int_0^t \theta \hat{m}_s^N ds\right). \end{aligned}$$

Apply the bound on $\hat{m}_{s-\tau_N}^N$ that we have from Lemma 4.6 to get

$$\begin{aligned} \hat{m}_t^N &\leq \int_{\tau_N}^t \theta \hat{m}_s^N ds - \left(\int_{\tau_N}^t \hat{m}_s^N e^{-\theta\tau_N} ds\right) 2N \left(1 + \frac{\theta}{N}\right) \sum_\ell \frac{p_\ell}{h_\ell} \mathbb{E}[Z_1^\ell] \\ &\quad + e^{\theta\tau_N} + 8N\tau_N + 8N\tau_N \left(\int_0^{\tau_N} \theta \hat{m}_s^N ds + \int_{\tau_N}^t \theta \hat{m}_s^N ds\right). \end{aligned}$$

Recall that $N\tau_N$ and τ_N converge to zero as N approaches infinity. This and (4.5) imply that there exists N_3 such that for every $N > N_3$ the following holds: $e^{\theta\tau_N} \leq \frac{\epsilon}{3}$, $8N\tau_N \leq \frac{\epsilon}{3}$ and $8N\tau_N \int_0^{\tau_N} \theta \hat{m}_s^N ds \leq \frac{\epsilon}{3}$. Thus, for any $N > \max(N_2, N_3)$

$$\hat{m}_t^N \leq e + \left(\int_{\tau_N}^t \hat{m}_s^N ds\right) \left(\theta(1 + 8N\tau_N) - e^{-\theta\tau_N} \left(1 + \frac{\theta}{N}\right) \sum_\ell \frac{p_\ell}{h_\ell} 2N \mathbb{E}[Z_1^\ell]\right).$$

Use Proposition 4.7, to see that for any $\epsilon > 0$ there exists N_ϵ such that

$$\hat{m}_t^N \leq e + \left(\int_{\tau_N}^t \hat{m}_s^N ds\right) \left(\left(1 + 8N\tau_N\right)\theta - e^{-\theta\tau_N} \left(1 + \frac{\theta}{N}\right)(\vartheta - \epsilon)\right),$$

for every $N > \max(N_2, N_3, N_\epsilon)$. Again, by using Grönwall's lemma we get

$$\hat{m}_t^N \leq e \cdot \exp \left[\left(\left(1 + 8N\tau_N\right)\theta - e^{-\theta\tau_N} \left(1 + \frac{\theta}{N}\right)(\vartheta - \epsilon) \right) (t - \tau_N) \right].$$

Since $\tau_N = \frac{\ln(N)}{N^2}$, there exists N_4 such that for any $N > N_4$ we have $e^{-\theta\tau_N} \left(1 + \frac{\theta}{N}\right) \geq 1$ and $8N\tau_N\theta \leq \epsilon$. Thus, for any $N > \max(N_2, N_3, N_4, N_\epsilon)$ we get

$$\hat{m}_t^N \leq e \cdot \exp [(\theta + \epsilon - (\vartheta - \epsilon))(t - \tau_N)].$$

Recall that $\vartheta - \theta > 0$, and fix $\epsilon := \frac{\vartheta - \theta}{4}$ to get

$$\hat{m}_t^N \leq e \cdot \exp \left[\frac{1}{2} (\theta - \vartheta) (t - \tau_N) \right].$$

Choose N_5 such that for any $N > N_5$ we have $\frac{1}{2}(\vartheta - \theta)\tau_N < 1$. Then,

$$\hat{m}_t^N \leq \exp \left[\frac{1}{2} (\theta - \vartheta) t + 2 \right], \quad \forall N > N_\theta := \max(N_2, N_3, N_4, N_5, N_\epsilon).$$

5. Proof of Proposition 4.7

This section is devoted for the proof of the Proposition 4.7. In particular we show that

$$\vartheta = \liminf_{N \rightarrow \infty} \sum_{\ell} \frac{p_\ell}{h_\ell} 2N \mathbb{E}[Z_1^\ell], \quad (5.1)$$

which is enough to prove the proposition.

Remark 5.1. The intuition behind equation (5.1) follows from ϑ being a quantity that represents rescaled rate with which the mass is lost due to the collisions. Intuitively, this is proportional to an average number of neighbors that can cause the collisions, where the average is based on the probability function P^b (see also (3.1) for a representation of P^b).

Let us first introduce an additional notation:

- Let e_i denote the vector in \mathbb{Z}^d with 1 in the i -th coordinate and 0's elsewhere.
- For $\ell \geq 1$, let $\varphi_d^\ell(x) = \{y \in \mathbb{Z}^d : \|y - x\|_1 = \ell\} \setminus \{0\}$ denote the ℓ -neighbourhood of x excluding the origin.
- For $\ell \geq 1$, let $\varphi_d^\ell = \varphi_d^\ell(0)$ denote the ℓ -neighbourhood of the origin.
- Let $\varphi_d(x) = \varphi_d^1(x)$ denote the nearest neighbours of x (excluding the origin), and recall that $\varphi_d = \varphi_d(0)$.
- Let $A_d(x) = \{y : y \in \varphi_d, y \in \varphi_d(x)\}$ denote the set of points which belong both to the nearest neighbours of the origin and to the nearest neighbours of x .

In addition, recall that for any $x \in \mathbb{Z}^d$, $x[i]$ denotes the i -th coordinate in the vector x .

Before we give the proof of Proposition 4.7 let us state a few auxiliary lemmas.

Lemma 5.2. *For each $x \in \varphi_d$ there is exactly one point (which we denote by z_x) in $\varphi_d(x)$ such that $|A_d(z_x)| = 1$. For any other point, $w_x \in \varphi_d(x)$, the following holds: $|A_d(w_x)| = 2$.*

Proof: Take $x \in \varphi_d$. Without loss of generality assume $x = e_i$, for some $i = 1, 2, \dots, d$.

Take $z_x = e_i + e_i$. Then $z_x \in \varphi_d(x)$ and $A_d(z_x) = \{e_i\} \Rightarrow |A_d(z_x)| = 1$. Now, take $y \in \varphi_d(x)$ such that $y \neq z_x$. Then $y[i] = 1$ and there exists only one $j \in \{1, 2, 3, \dots, d\} \setminus \{i\}$ such that $y[j] = 1$ or $y[j] = -1$. Thus, $A_d(y) = \{e_i, e_j\}$ (or $A_d(y) = \{e_i, -e_j\}$) and the result follows. \square

In order to express $\mathbb{E}(Z_1^\ell)$ in terms of the Green function we are going to couple the stirring process with a simple symmetric random walk. To this end we define the following:

Definition 5.3. Let $\{V_t^N\}_{t \geq 0}$ be a continuous time symmetric nearest neighbour random walk on \mathbb{Z}^d with jump rate $4dN^2$ and let Q^V be its transition rate matrix. Let $\{W_t^N\}_{t \geq 0}$ be a continuous time Markov chain taking values in \mathbb{Z}^d whose transition rate matrix is defined as follows:

$$Q^W(x, y) = \begin{cases} Q^V(x, y) & \text{if } x \notin \varphi_d, \\ N^2 & \text{if } x \in \varphi_d \text{ and } y = -x, \\ 2N^2 & \text{if } x \in \varphi_d \text{ and } y \in \varphi_d(x) \\ 0 & \text{otherwise.} \end{cases}$$

In order to prove Proposition 4.7 we would like to show that the expected times $\{V_t^N\}_{t \geq 0}$ and $\{W_t^N\}_{t \geq 0}$ spend in φ_d are the same. To this end, we partition the points that the processes can exit $\varphi_d \cup \{0\}$ via them. Notice that when $\{V_t^N\}_{t \geq 0}$ or $\{W_t^N\}_{t \geq 0}$ leaves $\varphi_d \cup \{0\}$ it does it via some point in $\varphi_d^2(0)$. Denote

$$J_1 = \bigcup_{x \in \varphi_d^2(0) : |A_d(x)|=1} \{x\}, \quad J_2 = \bigcup_{x \in \varphi_d^2(0) : |A_d(x)|=2} \{x\}.$$

From Lemma 5.2 it follows that for any $x \in \varphi_d^2(0)$ either $|A_d(x)| = 1$ or $|A_d(x)| = 2$. In the following lemma we prove that the probability that the random walk $\{V_t^N\}_{t \geq 0}$ exits the neighbourhood of the origin via points in J_1 is equal to the probability that $\{W_t^N\}_{t \geq 0}$ exits the neighbourhood of the origin via points in J_1 .

Lemma 5.4. Let $V_0^N, W_0^N \in \varphi_d$. Let τ_V and τ_W be the time that $\{V_t^N\}_{t \geq 0}$ and $\{W_t^N\}_{t \geq 0}$ leave $\varphi_d \cup \{0\}$ for the first time respectively. Then,

$$(1) \quad P(V_{\tau_V} \in J_1) = P(W_{\tau_W} \in J_1), \quad (2) \quad P(W_{\tau_V} \in J_2) = P(W_{\tau_W} \in J_2).$$

Proof: As mentioned above, from Lemma 5.2 it follows that

$$J_1 \cap J_2 = \emptyset, \quad J_1 \cup J_2 = \varphi_d^2. \quad (5.2)$$

In addition, notice that τ_V and τ_W are finite stopping times. Thus, (2) follows immediately from (1).

Assume $V_0^N \in \varphi_d$. Thus, using the first step analysis we obtain:

$$P(V_{\tau_V} \in J_1) = \frac{1}{2d} + 0 \cdot \frac{2d-2}{2d} + \frac{1}{2d} P(V_{\tau_V} \in J_1) \Rightarrow P(V_{\tau_V} \in J_1) = \frac{1}{2d-1}.$$

This holds since when $\{V_t^N\}_{t \geq 0}$ jumps from V_0^N , its location after the jump is uniformly distributed within the closest neighbours of V_0^N . Thus, with probability $\frac{1}{2d}$ it jumps to a point in J_1 ; with probability $\frac{2d-2}{2d}$ it jumps to a point in J_2 and with probability $\frac{1}{2d}$ it jumps to the origin (and then jumps back to φ_d).

Now consider $\{W_t^N\}_{t \geq 0}$. Assume $W_0^N = x \in \varphi_d$. Thus, again using the first step and (5.2) we obtain:

$$P(W_{\tau_V} \in J_1) = \frac{2}{4d-1} + 0 \cdot \frac{2(2d-2)}{4d-1} + \frac{1}{4d-1} P(W_{\tau_V} \in J_1).$$

Thus,

$$P(W_{\tau_V} \in J_1) = \frac{1}{2d-1}$$

This holds since when $\{W_t^N\}_{t \geq 0}$ jumps from x , with rate $2N^2$ it jumps to $z_x \in J_1$; with rate N^2 it jumps to $-W_0^N \in \varphi_d$ and with rate $2N^2(2d-2)$ it jumps to a point in J_2 . \square

In the next lemma we show that the expected times V^N and W^N spend in φ_d^ℓ are the same.

Lemma 5.5. *For every $l \in \{1, 2, 3, \dots\}$,*

$$\mathbb{E} \left[\int_0^t 1_{\{V_s^N \in \varphi_d^\ell\}} ds \right] = \mathbb{E} \left[\int_0^t 1_{\{W_s^N \in \varphi_d^\ell\}} ds \right].$$

Proof: It is shown in the proof of Lemma 3.4 in [Berezin and Mytnik \(2014\)](#) that once $\{V_t^N\}_{t \geq 0}$ enters φ_d the time $\{V_t^N\}_{t \geq 0}$ spends in φ_d before leaving the set $\varphi_d \cup \{0\}$ has the same distribution as the time spent by $\{W_t^N\}_{t \geq 0}$ in φ_d at any visit of this set.

By Lemma 5.4, the strong Markov property and the fact that, in distribution, the behavior of these two processes is exactly the same outside of φ_d the result follows. \square

Let us state one more lemma (it is standard so its proof is omitted).

Lemma 5.6. *Let $\{V_s\}_{s > 0}$ be a continuous time symmetric random walk on \mathbb{Z}^d with jump rate λ started at $V_0 = 0$. Let $\{D_n\}_{n > 0}$ be a discrete time symmetric random walk on \mathbb{Z}^d started at $D_0 = 0$. Let $\pi(u)$ be a Poisson process with rate 1 independent of $\{D_n\}_{n > 0}$. Then, for any $l = 1, 2, \dots$ and for any $t > 0$*

$$\int_0^t 1_{\{V_s \in \varphi_d^\ell\}} ds = \int_0^t 1_{\{D_{\pi(\lambda s)} \in \varphi_d^\ell\}} ds.$$

We are now ready to give the proof of Proposition 4.7.

Proof of Proposition 4.7: Recall that we start the process $\{\xi_t^N\}_{t \geq 0}$ with one particle denoted as particle 1. If particle 1 splits, $(1, 0)$ and $(1, 1)$ denote the two children of particle 1. Let $F_1 = \{T_1^c < T_1^d, T_1^c < \tau_N, T_{(1,0)} > \tau_N, T_{(1,1)} > \tau_N\}$ denote the event that particle 1 splits before time τ_N and its two children are alive at time τ_N . In addition recall that $\tau_N = \frac{\ln(N)}{N^2}$. Then, we have

$$P(F_1) = \frac{N + \theta}{2N + \theta} e^{-(2N+\theta)\tau_N} (1 - e^{-(2N+\theta)\tau_N}). \quad (5.3)$$

Now, fix arbitrary ℓ . Notice that $E(Z_1^\ell | F_1^c) = 0$. Thus, we only care about $P(B_{\tau_N}^{(1,0)} \sim B_{\tau_N}^{(1,1)} | F_1)$. Given F_1 , the time at which particle 1 splits is uniformly distributed in $[0, \tau_N]$. Thus,

$$E(Z_1^\ell | F_1) = \frac{1}{\tau_N} \int_0^{\tau_N} P(L + W_{\tau_N - t}^N \in \varphi_d^\ell) dt = \frac{1}{\tau_N} \int_0^{\tau_N} P(L + W_s^N \in \varphi_d^\ell) ds \quad (5.4)$$

where L is the difference in position of the two children of particle 1 right after its split.

Apply Lemma 5.5 on the right-hand side of (5.4) to get

$$E(Z_1^\ell | F_1) = \frac{1}{\tau_N} \int_0^{\tau_N} P(L + V_s^N \in \varphi_d^\ell) ds.$$

Now, let $\{D_n\}_{n \geq 1}$ be a simple (discrete) symmetric random walk on \mathbb{Z}^d independent of ℓ and started at $D_0 = 0$. Let $\{\pi(s)\}_{s > 0}$ be the Poisson process with rate 1 defined on the same probability space and independent of $\{D_n\}_{n \geq 1}$ and L . Thus, since $\{V_s^N\}_{s \geq 0}$ is a continuous time symmetric random walk with jump rate $4dN^2$, by Lemma 5.6 we get

$$\begin{aligned} E(Z_1^\ell | F_1) &= \frac{1}{\tau_N} \int_0^{\tau_N} P(L + D_{\pi(4dN^2s)} \in \varphi_d^\ell) ds \\ &= \frac{1}{\tau_N} \int_0^{4dN^2\tau_N} (4dN^2)^{-1} P(L + D_{\pi(r)} \in \varphi_d^\ell) dr. \end{aligned}$$

Overall we have

$$\begin{aligned} E(Z_1^\ell) &= \frac{N + \theta}{2N + \theta} e^{-(2N+\theta)\tau_N} \left(1 - e^{-(2N+\theta)\tau_N} \right) \\ &\quad \cdot \frac{1}{\tau_N} (4dN^2)^{-1} \int_0^{4dN^2\tau_N} P(L + D_{\pi(r)} \in \varphi_d^\ell) dr. \end{aligned} \quad (5.5)$$

Taking N to infinity, we get that for every $\ell \geq 1$

$$\lim_{N \rightarrow \infty} NE(Z_1^\ell) = (4d)^{-1} \int_0^\infty P(L + D_{\pi(r)} \in \varphi_d^\ell) dr = (4d)^{-1} \sum_{n=0}^\infty P(L + D_n \in \varphi_d^\ell), \quad (5.6)$$

where the second equality follows since the times between jumps of $D_{\pi(r)}$ are exponential with mean 1.

Denote

$$\Theta = \liminf_{N \rightarrow \infty} \sum_\ell \frac{p_\ell}{h_\ell} 2NE[Z_1^\ell]. \quad (5.7)$$

From (5.5), it is easy to check that $NE[Z_1^\ell] \leq (1 + \theta) \sum_{n=0}^\infty P(L + D_n \in \varphi_d^\ell)$ for all $N \geq 1$. Thus, by the Dominated Convergence theorem we get

$$\Theta = 2 \sum_\ell \frac{p_\ell}{h_\ell} \liminf_{N \rightarrow \infty} NE[Z_1^\ell].$$

Applying (5.6) on the right-hand side of the last term we get

$$\Theta = 2 \sum_\ell \frac{p_\ell}{h_\ell} (4d)^{-1} \sum_{n=0}^\infty P(L + D_n \in \varphi_d^\ell) = (2d)^{-1} \sum_\ell \frac{p_\ell}{h_\ell} \sum_{n=0}^\infty \sum_{y \in \varphi_d^\ell} P(L + D_n = y).$$

By Tonelli's theorem for non-negative functions and the total probability law we get

$$\begin{aligned} \Theta &= (2d)^{-1} \sum_\ell \frac{p_\ell}{h_\ell} \sum_{y \in \varphi_d^\ell} \sum_{n=0}^\infty P(L + D_n = y) \\ &= (2d)^{-1} \sum_\ell \frac{p_\ell}{h_\ell} \sum_{y \in \varphi_d^\ell} \sum_{n=0}^\infty \sum_{x \in \mathbb{Z}^d} P(x + D_n = y) P(L = x). \end{aligned} \quad (5.8)$$

Let $P^n(x, y) = P(D_n = y | D_0 = x)$ be the probability that a discrete time symmetric random walk on \mathbb{Z}^d , started at site x , lands on site y after n steps. Recall also

that $P(L = x) = P^b(0, x)$. Thus, we get

$$\Theta = (2d)^{-1} \sum_{\ell} \frac{p_{\ell}}{h_{\ell}} \sum_{y \in \varphi_d^{\ell}} \sum_{n=0}^{\infty} \sum_{x \in \mathbb{Z}^d} P^n(x, y) P^b(0, x).$$

By Tonelli's theorem for non-negative functions, and the definition of the Green function we get

$$\Theta = (2d)^{-1} \sum_{x \in \mathbb{Z}^d} P^b(0, x) \sum_{\ell} \sum_{y \in \varphi_d^{\ell}} \frac{p_{\ell}}{h_{\ell}} G(x, y).$$

By the definition of $P^b(0, y)$ we have

$$\Theta = (2d)^{-1} \sum_{x \in \mathbb{Z}^d} \sum_{y \in \mathbb{Z}^d} P^b(0, x) P^b(0, y) G(x, y),$$

and we are done. □

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